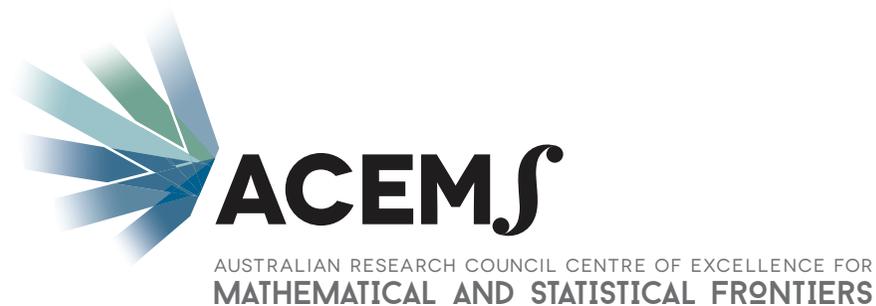


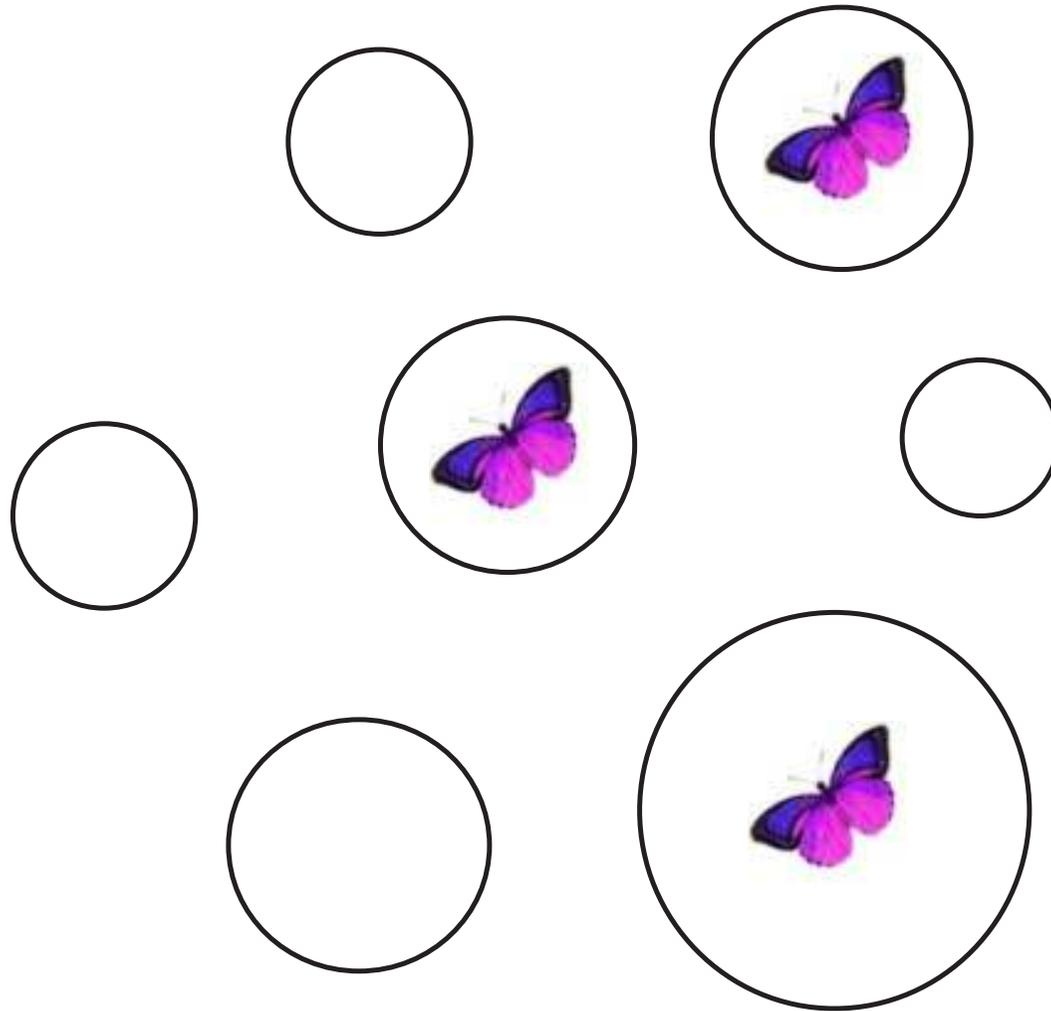
Metapopulations in dynamic landscapes

Phil Pollett and Ross McVinish

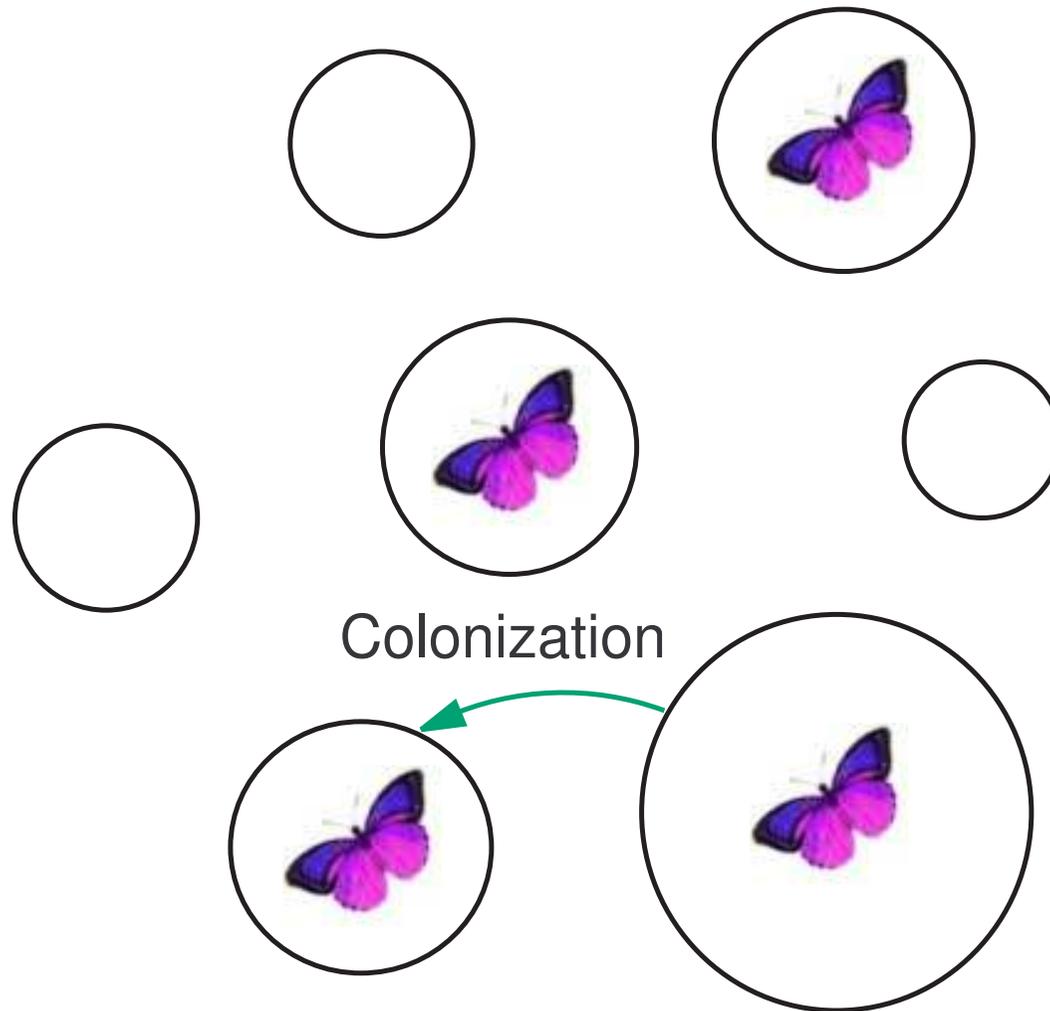
The University of Queensland
<http://www.maths.uq.edu.au/~pkp>



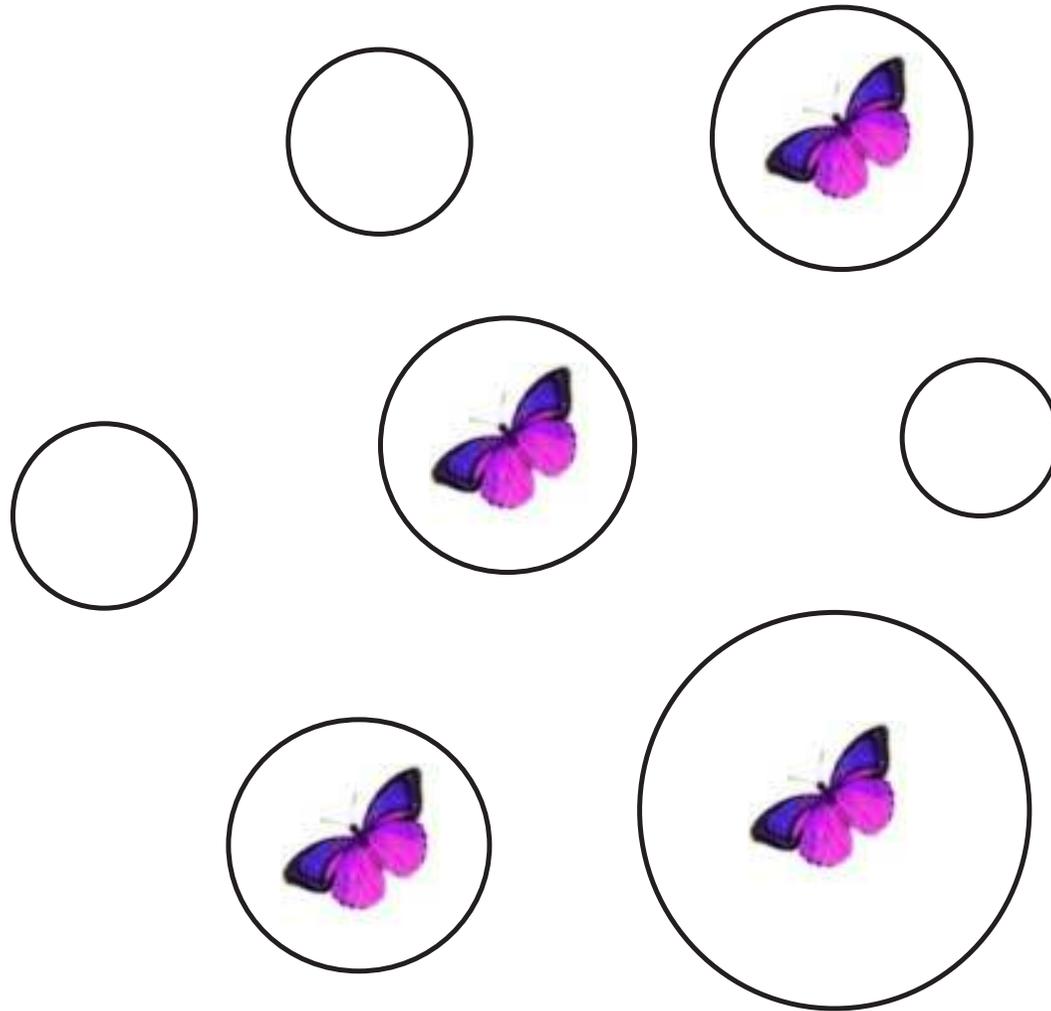
Metapopulations



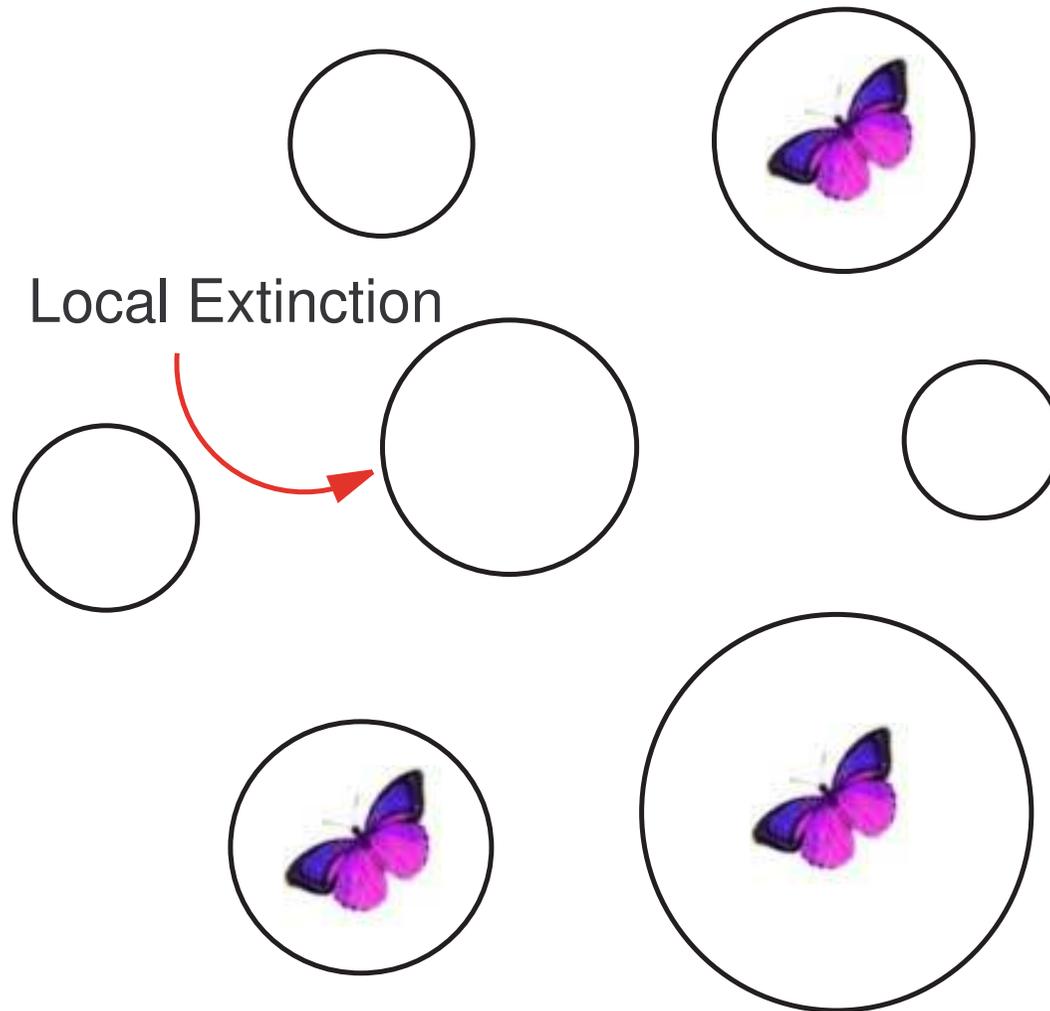
Metapopulations



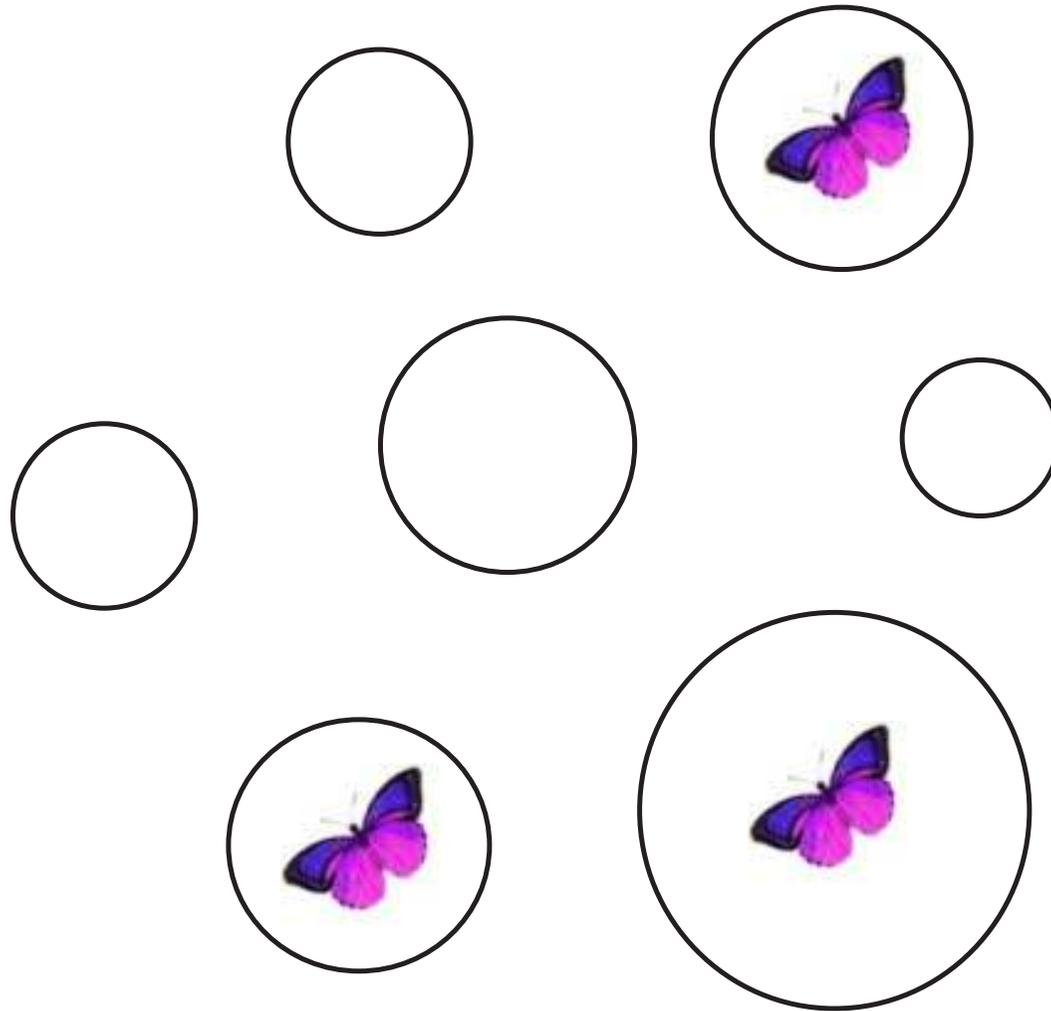
Metapopulations



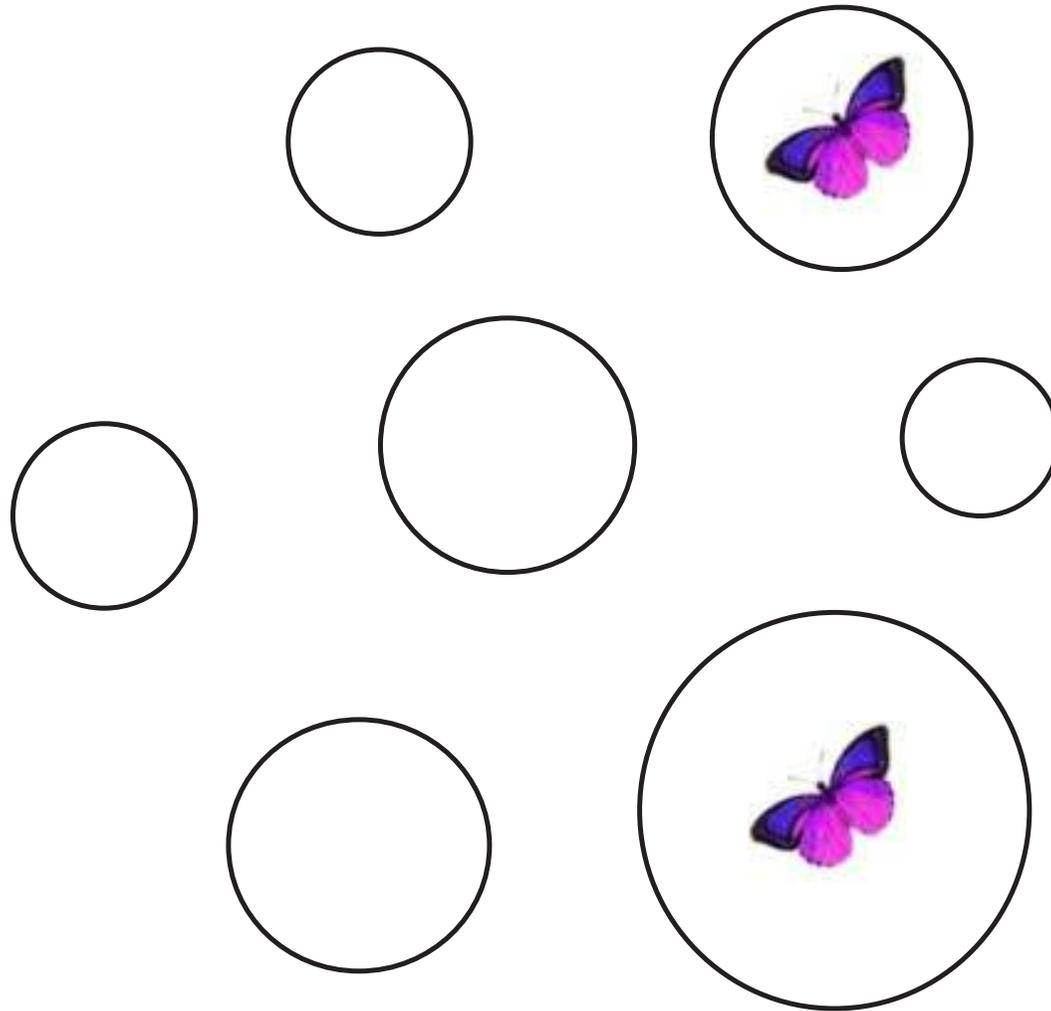
Metapopulations



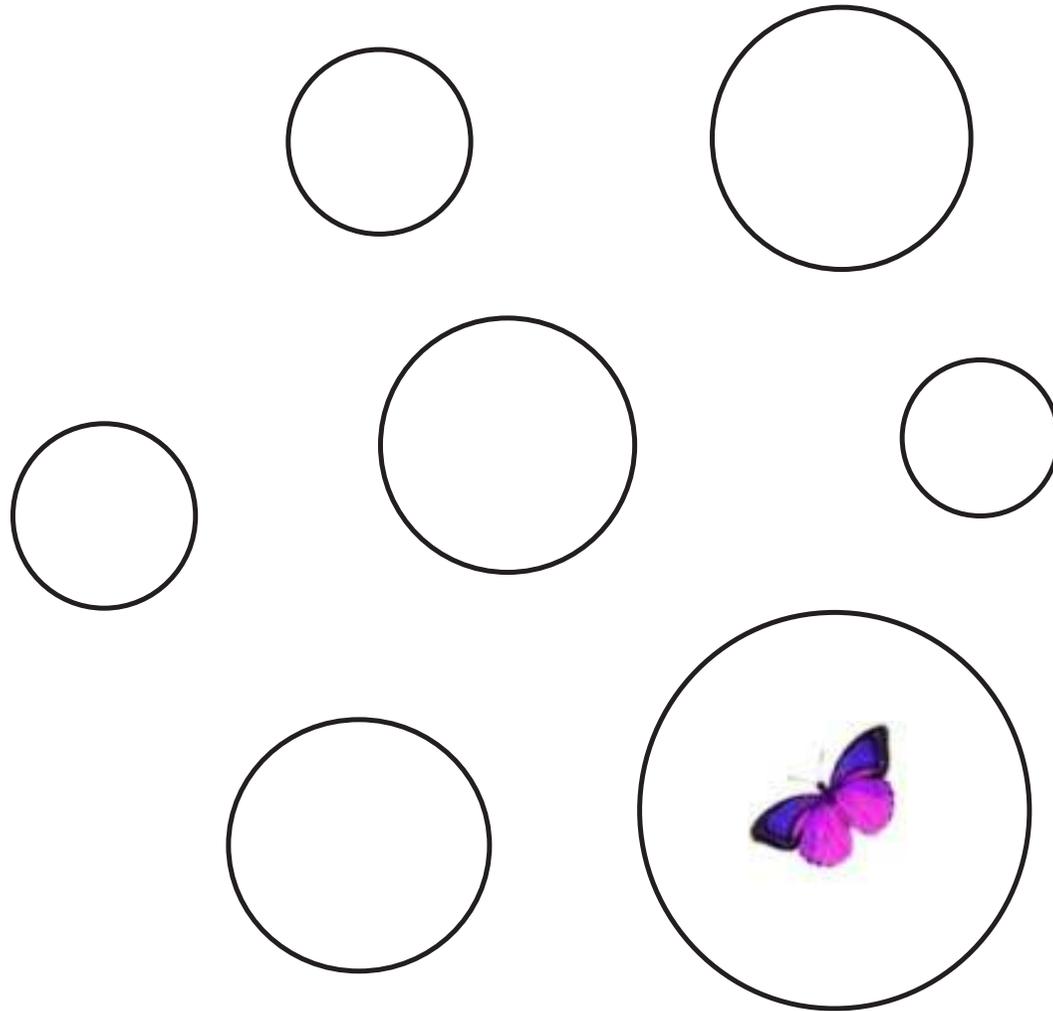
Metapopulations



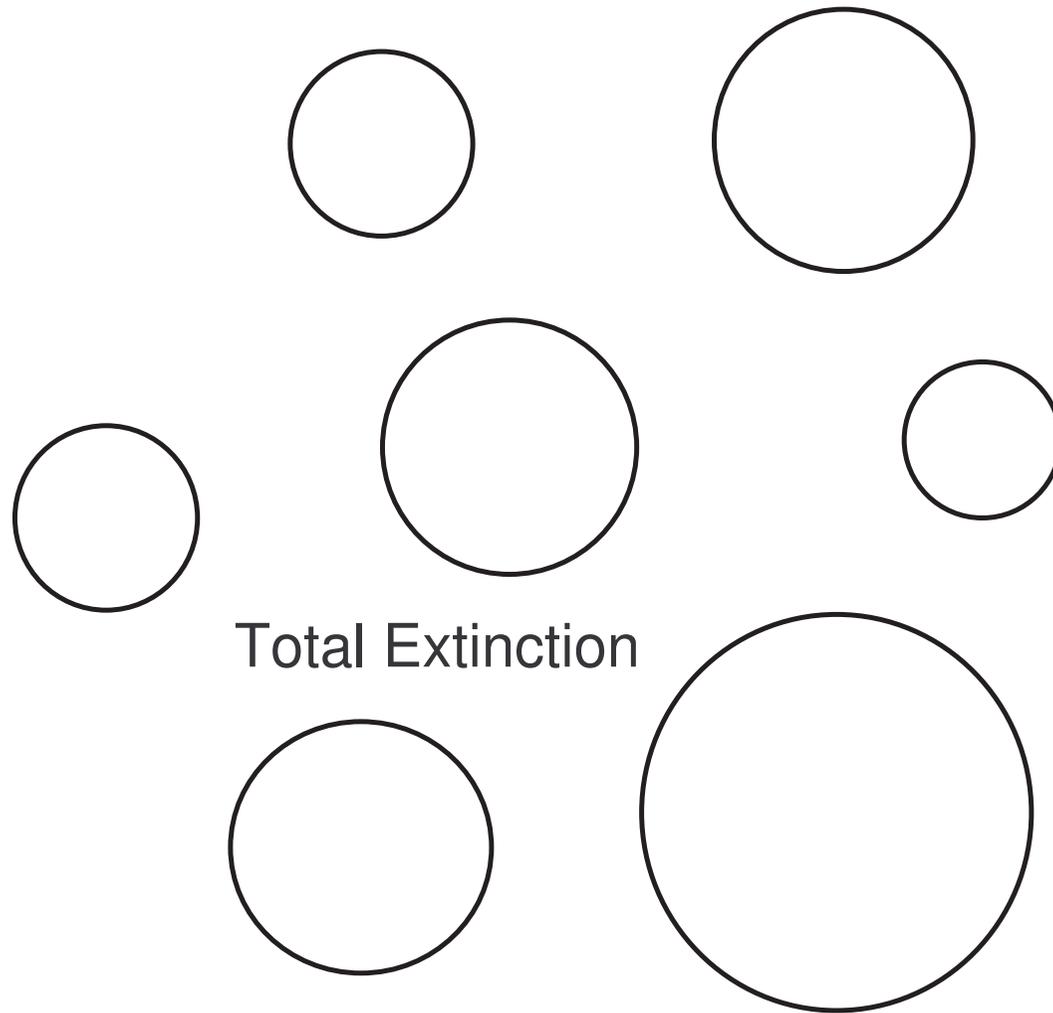
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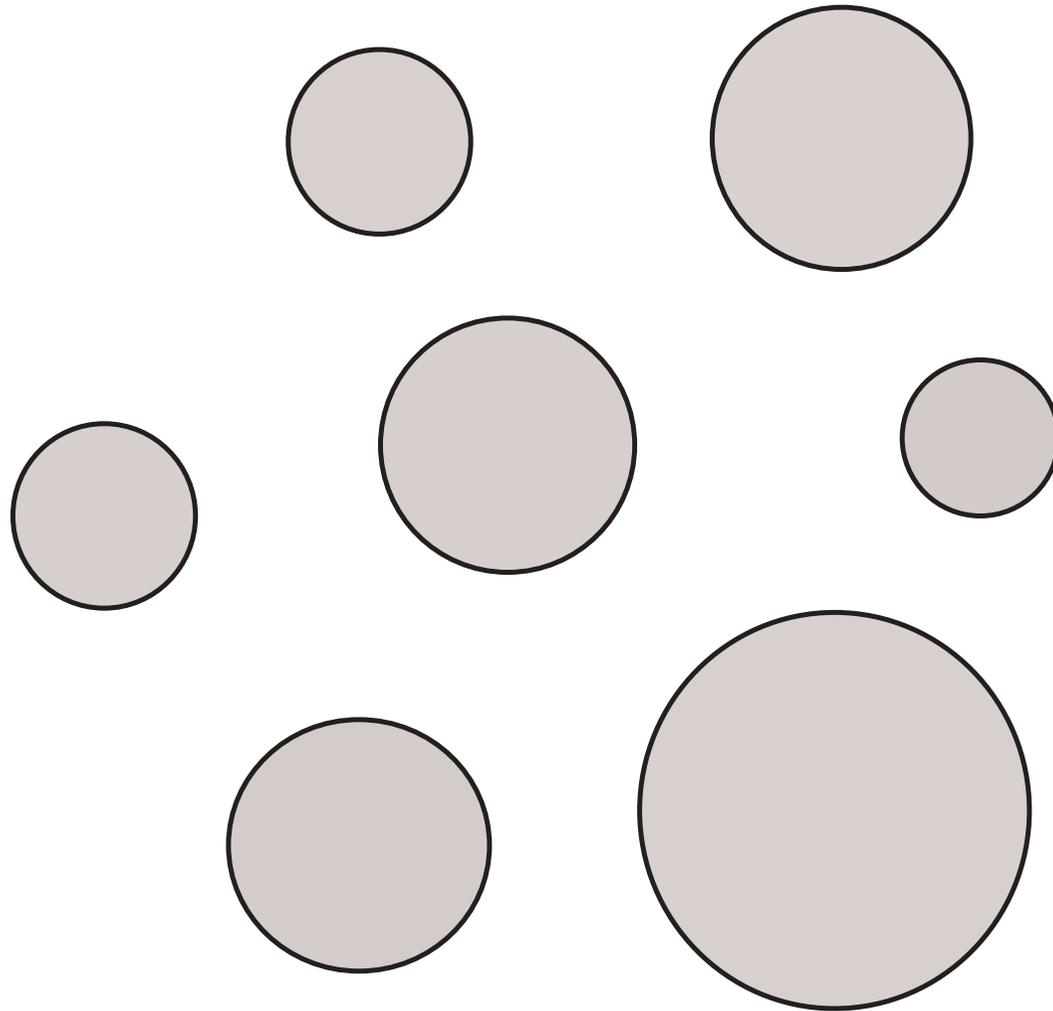
Metapopulations



Metapopulations



Metapopulations



A stochastic patch occupancy model (SPOM)

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Suppose that there are n patches.

Let $X_t^{(n)} = (X_{1,t}^{(n)}, \dots, X_{n,t}^{(n)})$, where $X_{i,t}^{(n)}$ is a binary variable indicating whether or not patch i is occupied at time t .

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Colonization and extinction happen in distinct, successive phases.

SPOM - Phase structure

For many species the propensity for colonization and local extinction is markedly different in different phases of their life cycle. Examples:

The Vernal pool fairy shrimp (*Branchinecta lynchi*) and the California linderiella (*Linderiella occidentalis*), both listed under the Endangered Species Act (USA)

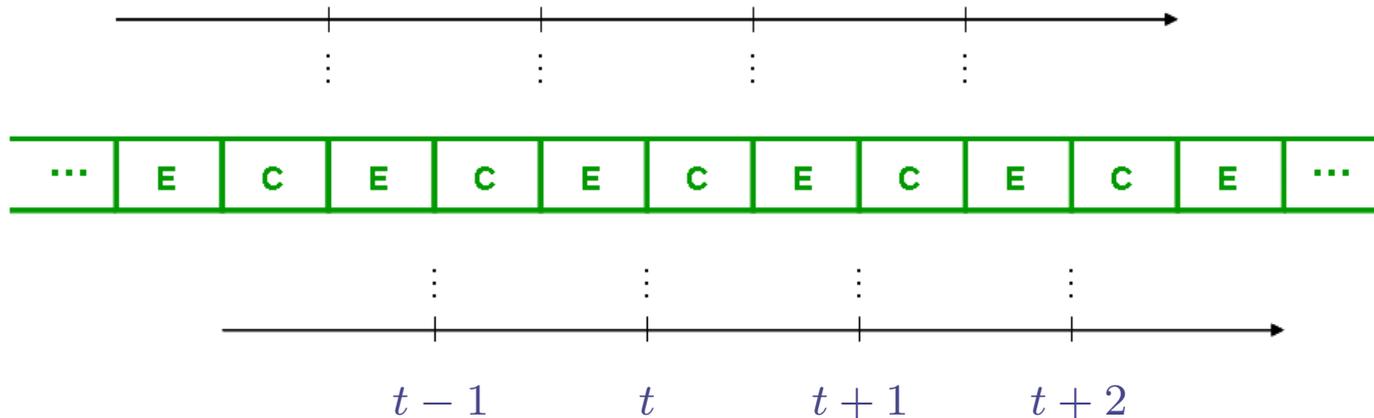


The Jasper Ridge population of Bay checkerspot butterfly (*Euphydryas editha bayensis*), now extinct



SPOM - Phase structure

Colonization and extinction happen in distinct, successive phases.



We will assume that the population is *observed after successive extinction phases* (CE Model).

SPOM - Phase structure

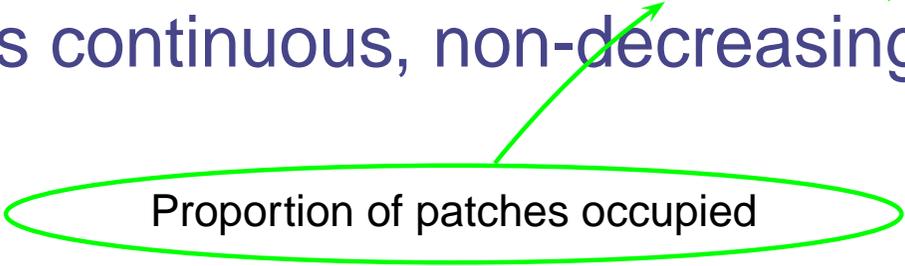
Colonization and extinction happen in distinct, successive phases, as independent trials.

Colonization: unoccupied patches become occupied independently with probability $c(n^{-1} \sum_{i=1}^n X_{i,t}^{(n)})$, where $c : [0, 1] \rightarrow [0, 1]$ is continuous, non-decreasing and concave.

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Proportion of patches occupied

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[More generally, we can allow $c(\cdot)$ to depend on the *relative positions* of all patches and their *areas*, and allow the survival probabilities to *evolve in time*.]

SPOM - Phase structure

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Extinction: occupied patch i remains occupied independently with probability s_i (fixed or random).

SPOM - example

$n = 30$ patches

0 0 0 0 1 0 1 1 0 1 0 1 0 0 0 0 1 1 1 0 1 0 1 0 0 0 1 0 0 0

(11 patches occupied)

SPOM - example

$$n = 30, c(x) = 0.7x$$

0 0 0 0 1 0 1 1 0 1 0 1 0 0 0 0 1 1 1 0 1 0 1 0 0 0 1 0 0 0

$$c(x) = c\left(\frac{11}{30}\right) = 0.7 \times 0.3\dot{6} = 0.25\dot{6}$$

SPOM - example

$$n = 30, c(x) = 0.7x$$

0 0 0 0 1 0 1 1 0 1 0 1 0 0 0 0 1 1 1 0 1 0 1 0 0 0 1 0 0
C 1 0 0 0 1 1 1 1 0 1 0 1 0 0 0 0 1 1 1 1 1 1 0 0 0 1 0 1 0

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0 0 0 0 1 0 1 1 0 0 0 1 0 0 0 0 1 1 1 0 1 0 1 0 0 0 1 0 0 0
C 1 0 0 0 1 1 1 1 0 1 0 1 0 0 0 0 1 1 1 1 1 1 1 0 0 0 1 0 1 0

SPOM - example

$n = 30$, $c(x) = 0.7x$ and $s_i \sim \text{Beta}(25.2, 19.8)$ ($\mathbb{E}s_i = 0.56$)

	0	0	0	0	1	0	1	1	0	0	0	1	0	0	0	0	0	1	1	1	0	1	0	1	0	0	0	0	1	0	0	0
C	1	0	0	0	1	1	1	1	0	1	0	1	0	0	0	0	0	1	1	1	1	1	1	1	1	0	0	0	1	0	1	0
	0.60				0.56	0.63		0.62	0.52								0.61	0.68	0.49	0.49								0.49	0.50			
					0.41	0.59											0.63	0.60	0.61													

[Survival probabilities listed for occupied patches only]

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C	1	0	0	0	1	1	1	1	0	1	0	1	0	0	0	0	1	1	1	1	1	1	1	1	0	0	0	1	0	1	0
E	0	0	0	0	1	0	0	1	0	1	0	1	0	0	0	0	1	0	1	1	1	1	1	0	0	0	0	0	0	1	0

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	0	0	0	0	1	0	1	1	0	1	0	1	0	0	0	0	1	1	1	0	1	0	1	0	0	0	1	0	0	0		
C	1	0	0	0	1	1	1	1	0	1	0	1	0	0	0	0	1	1	1	1	1	1	1	1	0	0	0	1	0	1	0	
E	0	0	0	0	1	0	0	1	0	1	0	1	0	0	0	0	1	0	1	1	1	1	0	0	0	0	0	0	0	0	1	0

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C	1	0	0	0	1	1	1	1	0	1	0	1	0	0	0	0	1	1	1	1	1	1	1	1	0	0	0	1	0	1	0	
E	0	0	0	0	1	0	0	1	0	1	0	1	0	0	0	0	1	0	1	1	1	1	0	0	0	0	0	0	0	0	1	0

$$c(x) = c\left(\frac{10}{30}\right) = 0.7 \times 0.\dot{3} = 0.2\dot{3}$$

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	0	0	0	0	1	0	1	1	0	1	0	1	0	0	0	0	1	1	1	0	1	0	1	0	0	0	1	0	0	0		
C	1	0	0	0	1	1	1	1	0	1	0	1	0	0	0	0	1	1	1	1	1	1	1	1	0	0	0	1	0	1	0	
E	0	0	0	0	1	0	0	1	0	1	0	0	0	0	1	0	1	1	1	1	1	0	0	0	0	0	0	0	0	1	0	
C	0	0	1	0	1	0	0	1	1	1	0	1	0	0	1	0	1	1	1	1	1	1	0	0	0	0	0	0	0	0	1	0

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C	1	0	0	0	1	1	1	1	0	1	0	1	0	0	0	0	1	1	1	1	1	1	1	1	0	0	0	1	0	1	0	
E	0	0	0	0	1	0	0	1	0	1	0	1	0	0	0	0	1	0	1	1	1	1	0	0	0	0	0	0	0	1	0	
C	0	0	1	0	1	0	0	1	1	1	0	1	0	0	1	0	1	1	1	1	1	1	1	0	0	0	0	0	0	0	1	0

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C	1	0	0	0	1	1	1	1	0	1	0	1	0	0	0	0	1	1	1	1	1	1	1	0	0	0	1	0	1	0		
E	0	0	0	0	1	0	0	1	0	1	0	1	0	0	0	0	1	0	1	1	1	1	0	0	0	0	0	0	0	1	0	
C	0	0	1	0	1	0	0	1	1	1	0	1	0	0	1	0	1	1	1	1	1	1	0	0	0	0	0	0	0	1	0	
E	0	0	0	0	1	0	0	1	0	1	0	1	0	0	0	0	0	1	0	0	0	0	1	0	0	0	0	0	0	0	1	0

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```
      0 0 0 0 1 0 1 1 0 1 0 1 0 0 0 0 1 1 1 0 1 0 1 0 0 0 1 0 0 0
C 1 0 0 0 1 1 1 1 0 1 0 1 0 0 0 0 1 1 1 1 1 1 1 0 0 0 1 0 1 0
E 0 0 0 0 1 0 0 1 0 1 0 1 0 0 0 0 1 0 1 1 1 1 0 0 0 0 0 0 1 0
C 0 0 1 0 1 0 0 1 1 1 0 1 0 0 1 0 1 1 1 1 1 1 0 0 0 0 0 0 1 0
E 0 0 0 0 1 0 0 1 0 1 0 1 0 0 0 0 0 1 0 0 0 1 0 0 0 0 0 0 1 0
```

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```
      0 0 0 0 1 0 1 1 0 1 0 1 0 0 0 0 1 1 1 0 1 0 1 0 0 0 1 0 0 0
C 1 0 0 0 1 1 1 1 0 1 0 1 0 0 0 0 1 1 1 1 1 1 1 0 0 0 1 0 1 0
E 0 0 0 0 1 0 0 1 0 1 0 1 0 0 0 0 1 0 1 1 1 1 0 0 0 0 0 0 1 0
C 0 0 1 0 1 0 0 1 1 1 0 1 0 0 1 0 1 1 1 1 1 1 0 0 0 0 0 0 1 0
E 0 0 0 0 1 0 0 1 0 1 0 1 0 0 0 0 0 1 0 0 0 1 0 0 0 0 0 0 1 0
.
.
.
C 0 0 0 0 1 0 0 0 0 0 0 0 0 0 0 0 1 0 0 0 0 0 0 0 0 0 0 0 0 0
E 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0
```

The evolution of the process can be summarized by

$$X_{i,t+1}^{(n)} \stackrel{d}{=} \mathbf{Bin}\left(X_{i,t}^{(n)} + \mathbf{Bin}\left(1 - X_{i,t}^{(n)}, c\left(\frac{1}{n} \sum_{j=1}^n X_{j,t}^{(n)}\right)\right), s_i\right),$$

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a “*Chain Bernoulli*” structure.

In the *homogeneous case*, where $s_i = s$ is the same for each i , the *number* $N_t^{(n)}$ of occupied patches at time t is Markovian. It has the following *Chain Binomial* structure:

$$N_{t+1}^{(n)} \stackrel{d}{=} \mathbf{Bin}\left(N_t^{(n)} + \mathbf{Bin}\left(n - N_t^{(n)}, c\left(\frac{1}{n} N_t^{(n)}\right)\right), s\right).$$

A deterministic limit

Letting the initial number $N_0^{(n)}$ of occupied patches grow at the same rate as $n \dots$

Theorem If $N_0^{(n)} / n \xrightarrow{p} x_0$ (a constant), then

$$N_t^{(n)} / n \xrightarrow{p} x_t, \quad \text{for all } t \geq 1,$$

with (x_t) determined by $x_{t+1} = f(x_t)$, where

$$f(x) = s(x + (1 - x)c(x)).$$

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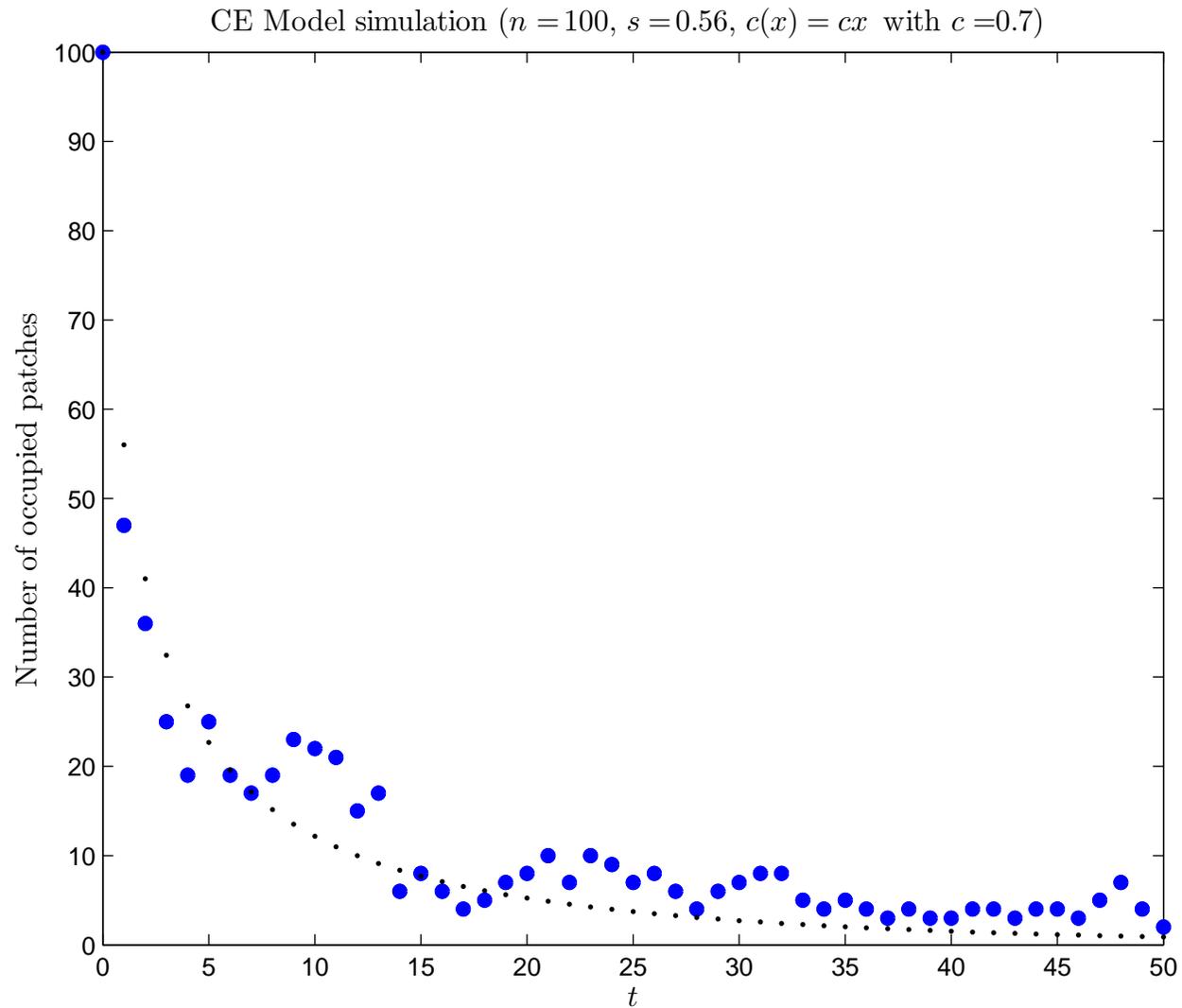
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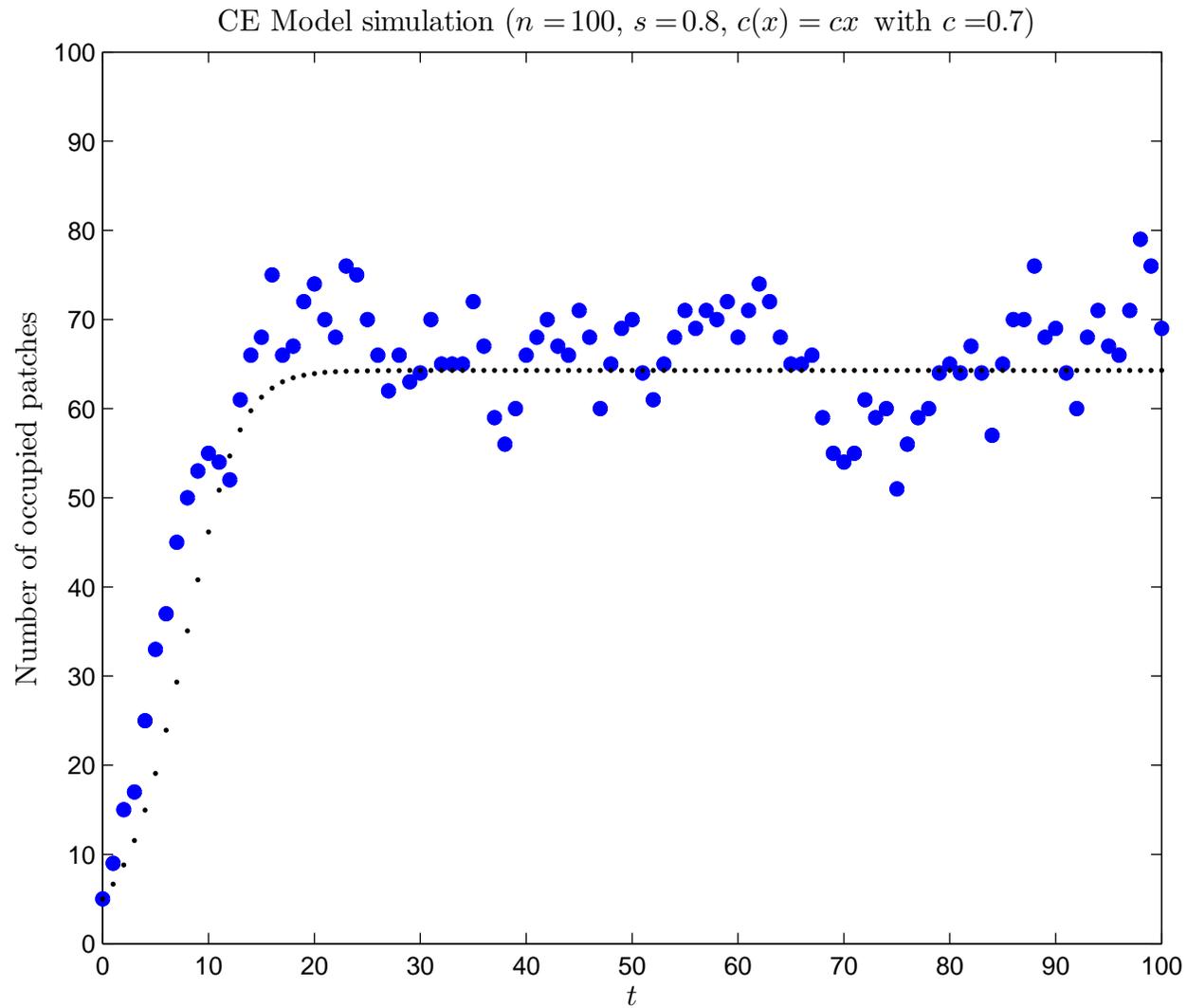
Survival probability

Colonization probability

CE Model - Evanescence



CE Model - Quasi stationarity



Stability

$x_{t+1} = f(x_t)$, where $f(x) = s(x + (1 - x)c(x))$.

Stationarity: $c(0) > 0$. There is a unique fixed point $x^* \in [0, 1]$. It satisfies $x^* \in (0, 1)$ and is stable.

Evanescence: $c(0) = 0$ and $1 + c'(0) \leq 1/s$. Now 0 is the unique fixed point in $[0, 1]$. It is stable.

Quasi stationarity: $c(0) = 0$ and $1 + c'(0) > 1/s$. There are two fixed points in $[0, 1]$: 0 (unstable) and $x^* \in (0, 1)$ (stable).

Stability

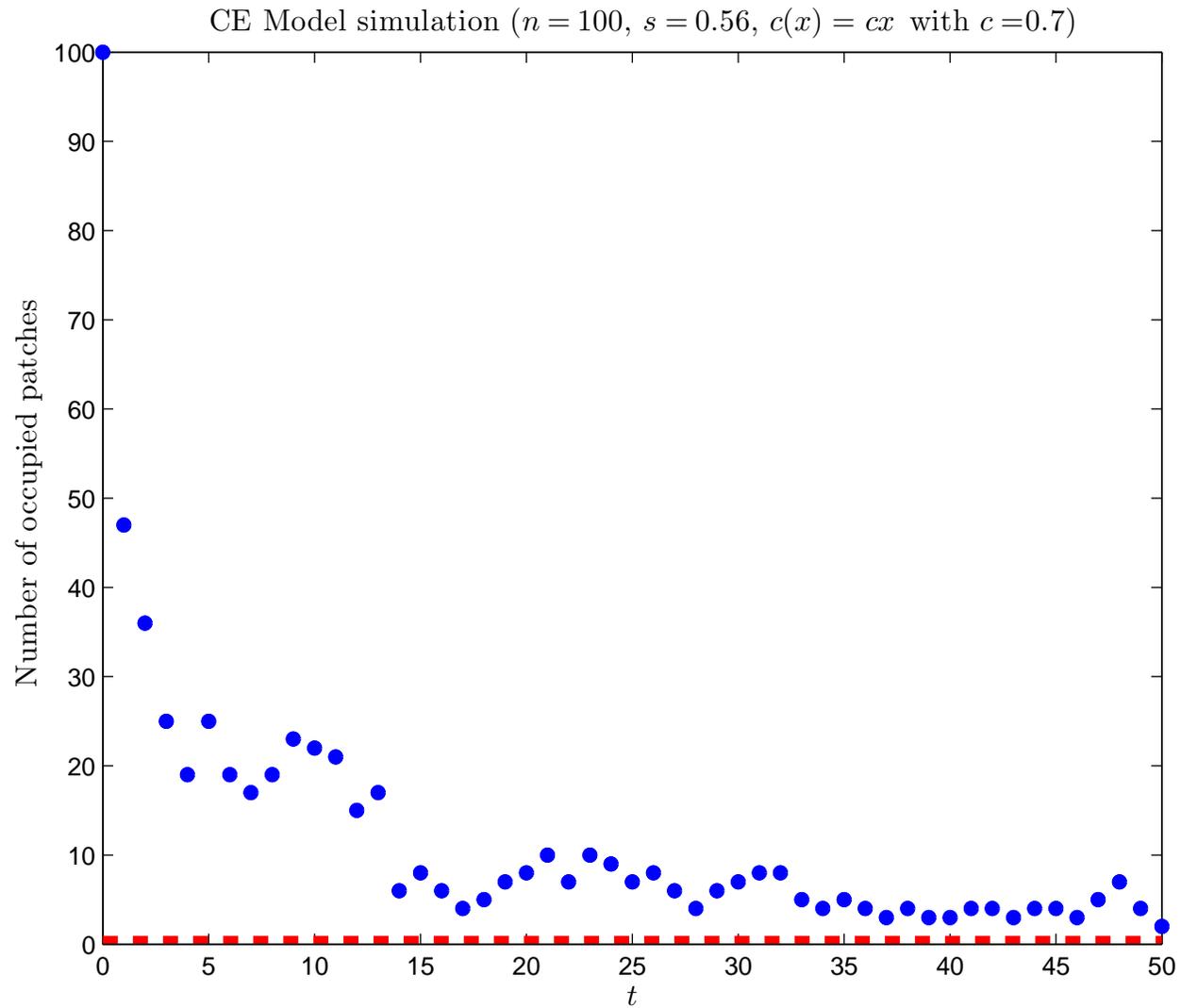
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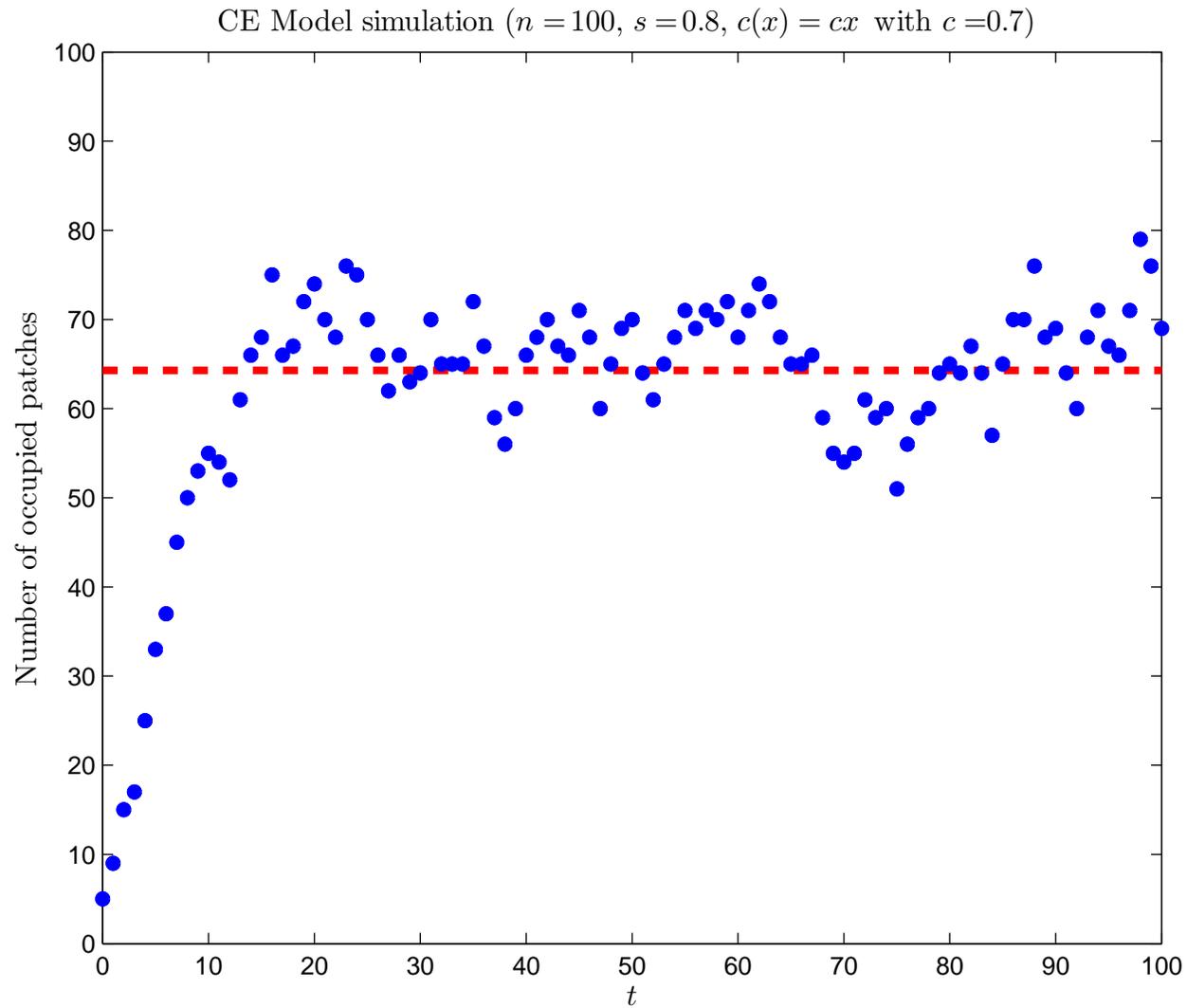
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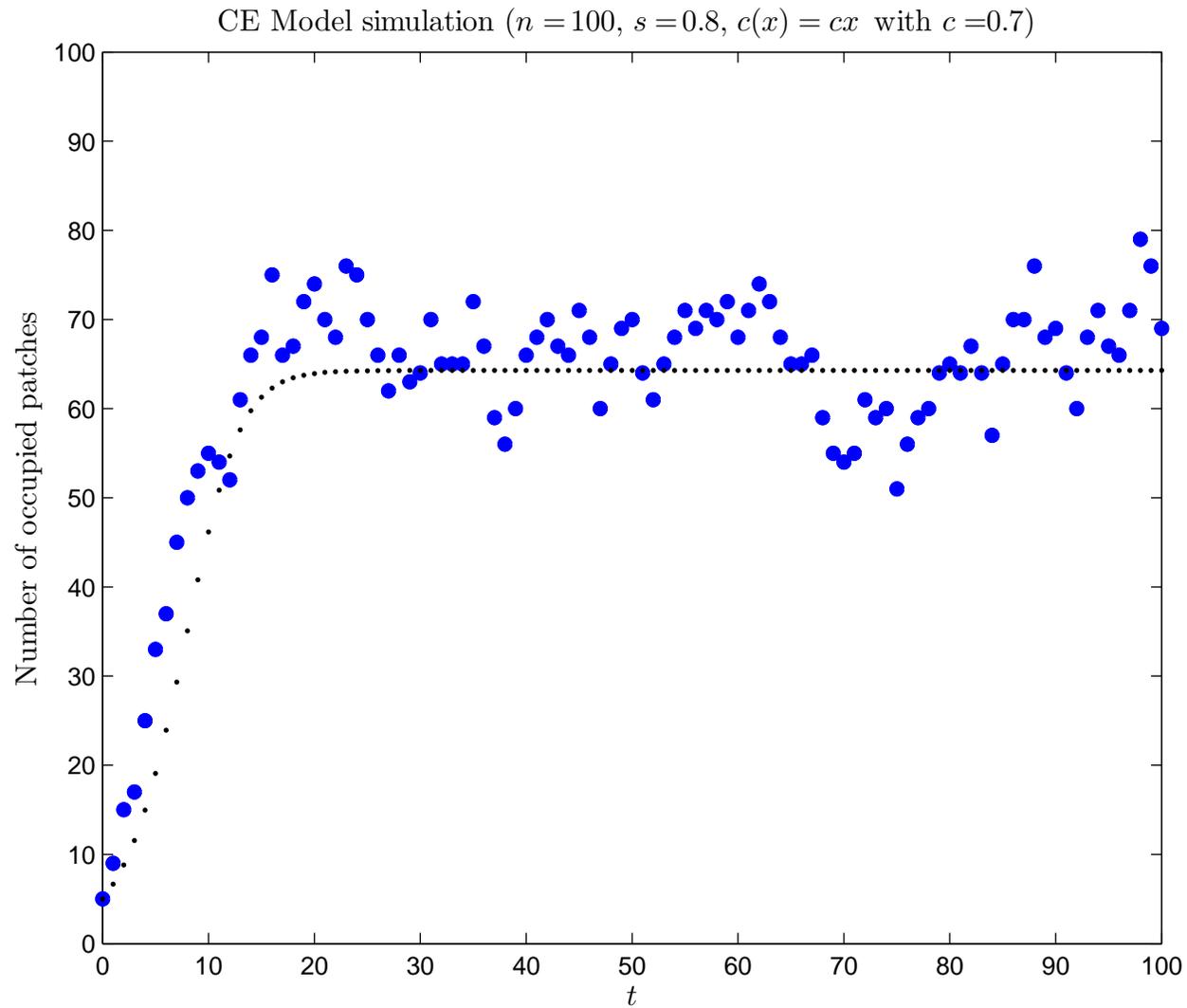
CE Model - Evanescence



CE Model - Quasi stationarity



CE Model - Quasi stationarity



A Gaussian limit

Theorem Further suppose that $c(x)$ is twice continuously differentiable, and let

$$Z_t^{(n)} = \sqrt{n}(N_t^{(n)} / n - x_t).$$

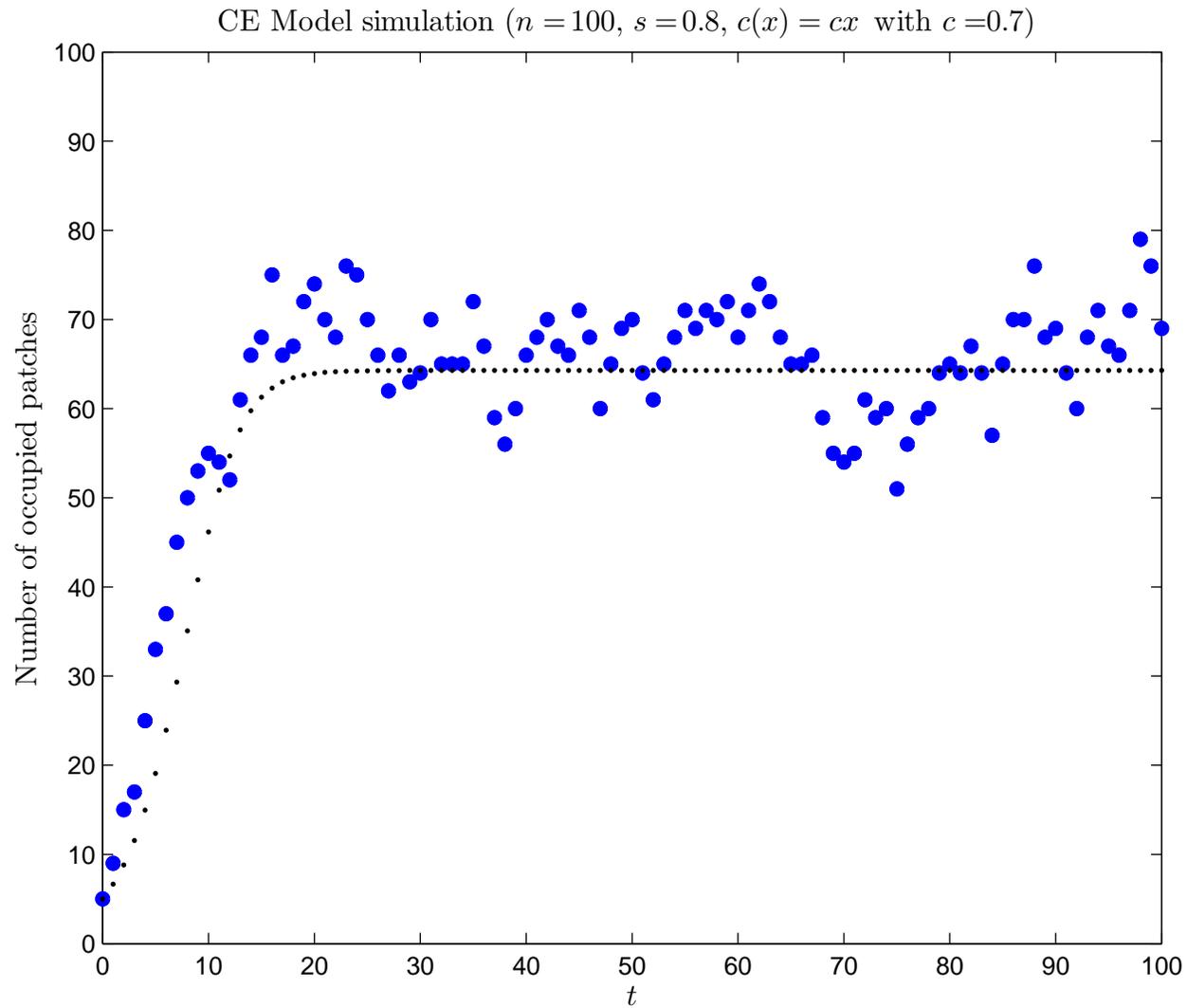
If $Z_0^{(n)} \xrightarrow{d} z_0$, then $Z_{\bullet}^{(n)}$ converges weakly to the Gaussian Markov chain Z_{\bullet} defined by

$$Z_{t+1} = f'(x_t)Z_t + E_t \quad (Z_0 = z_0),$$

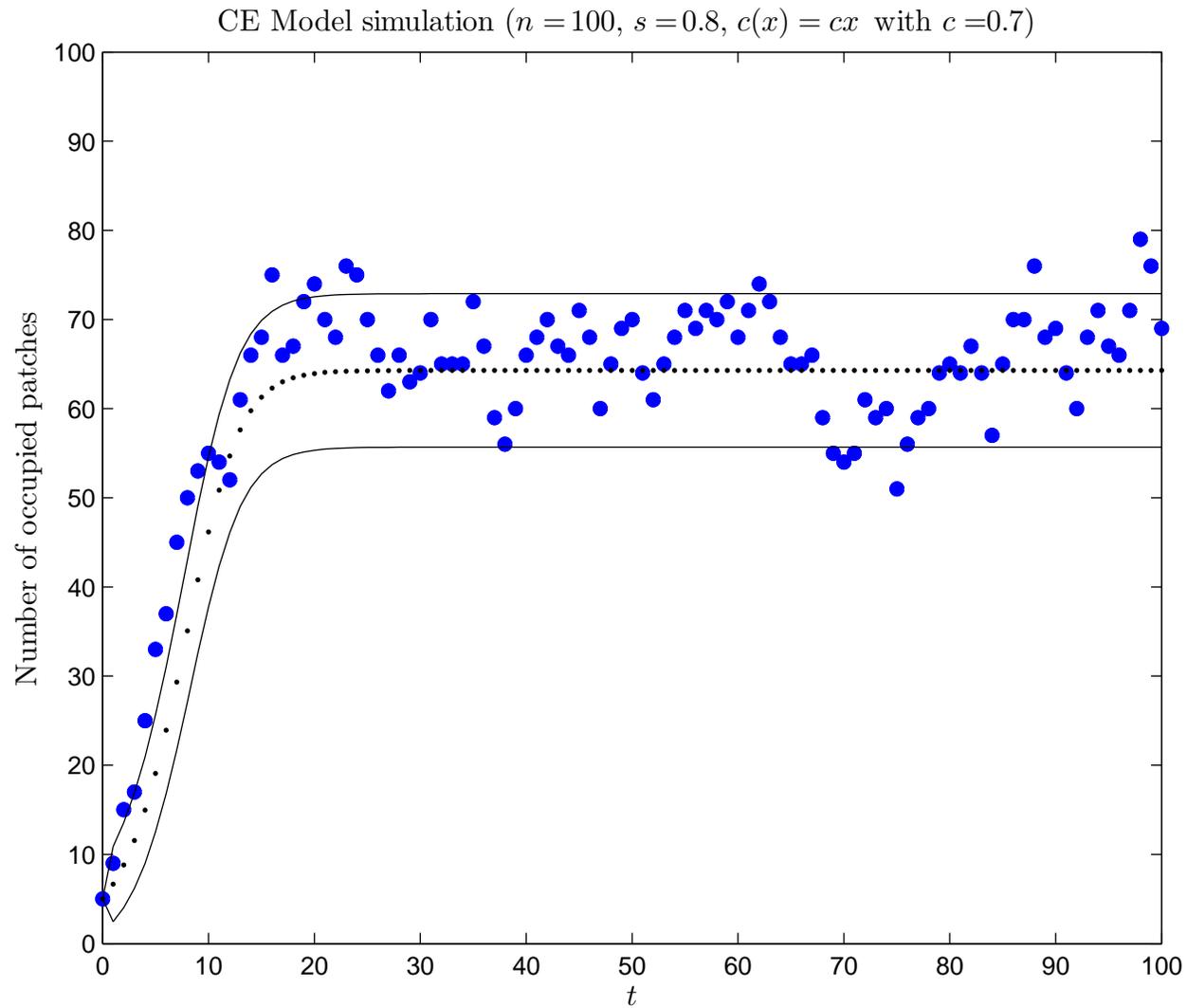
with (E_t) independent and $E_t \sim N(0, v(x_t))$, where

$$v(x) = s \left[(1 - s)x + (1 - x)c(x)(1 - sc(x)) \right].$$

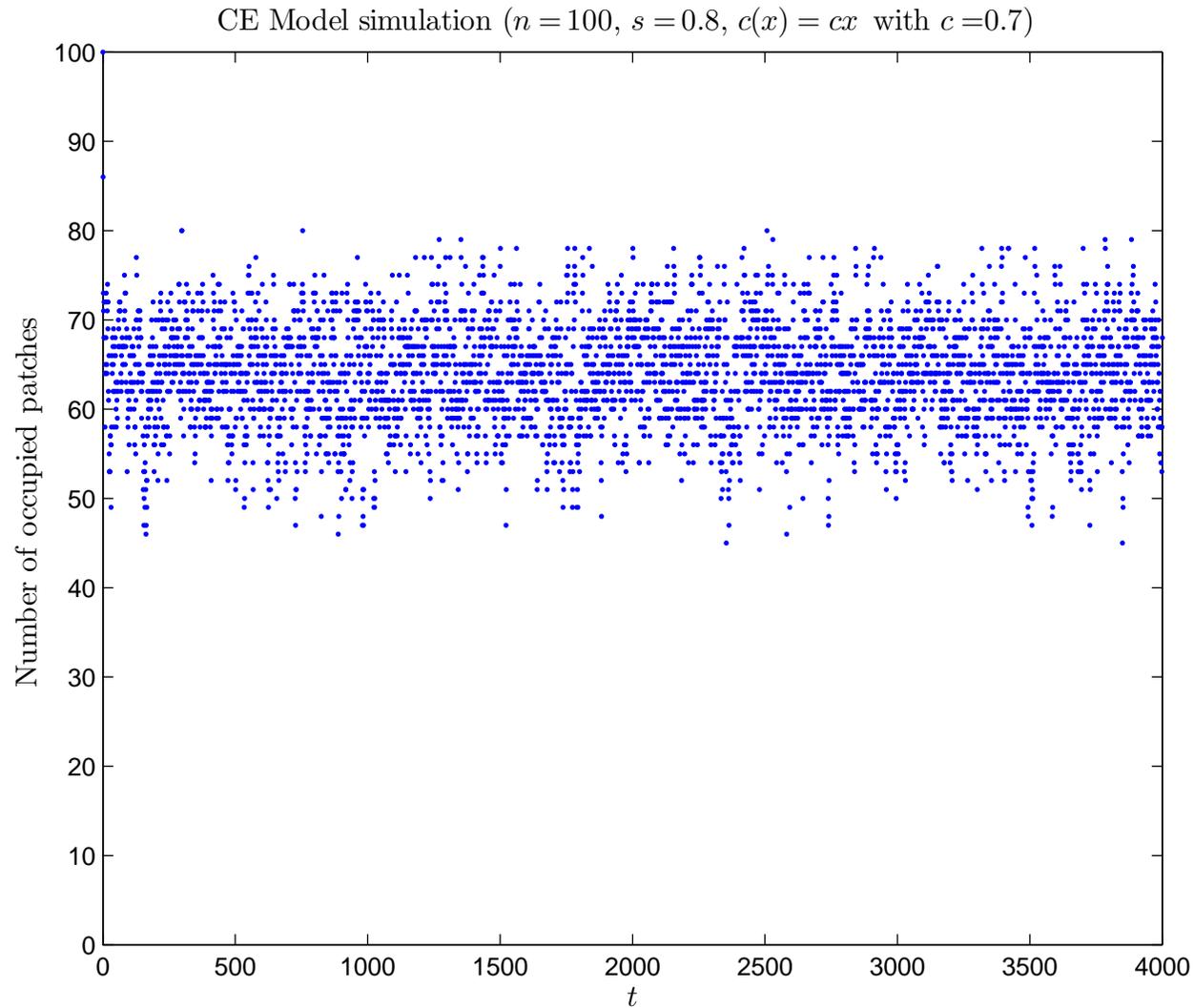
CE Model - Quasi stationarity



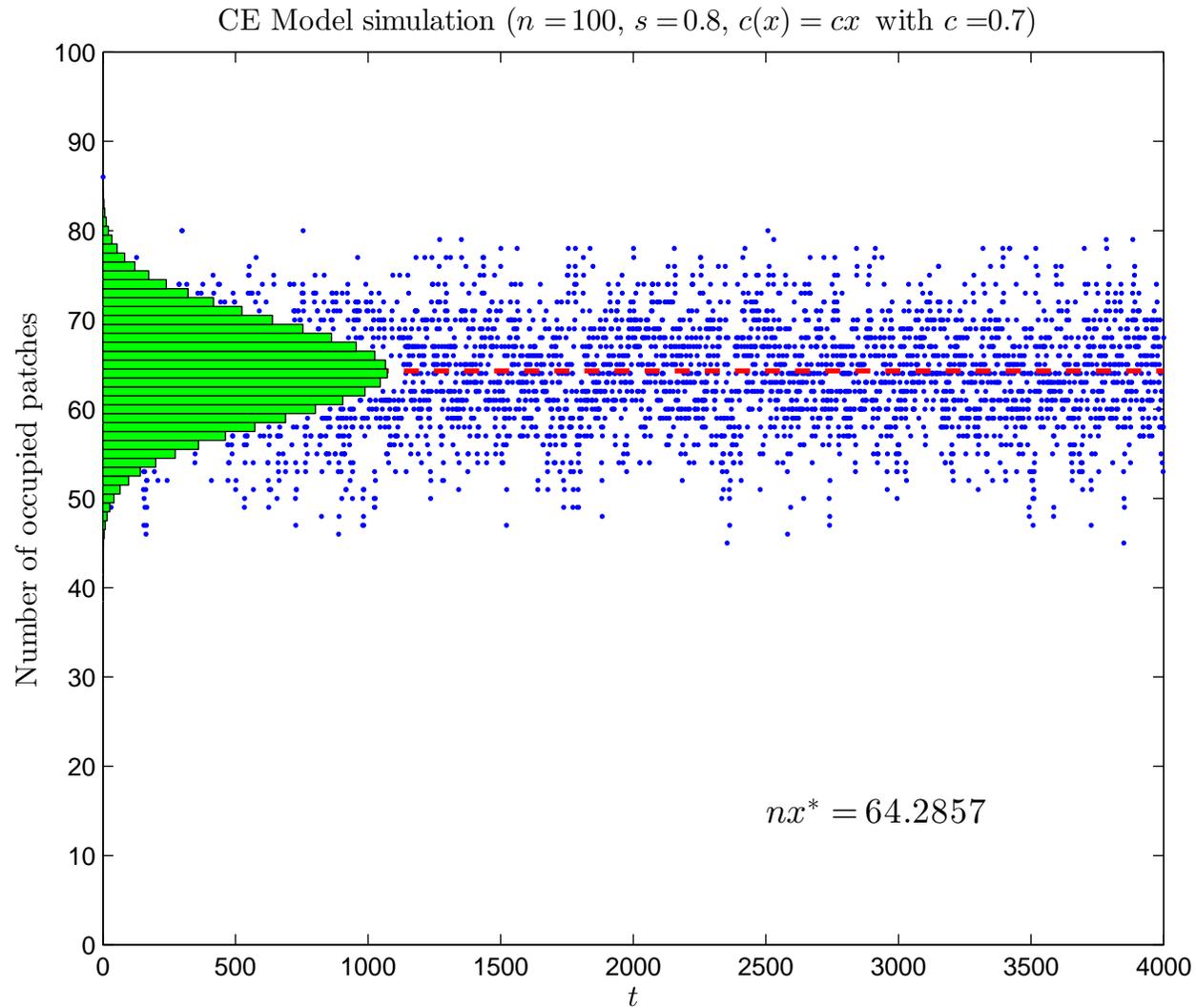
CE Model - Gaussian approximation



CE Model - Quasi stationarity



CE Model - Gaussian approximation



SPOM - general case

Return now to the general case, where patch survival probabilities (s_i) are *random* and *patch dependent*, and we keep track of which patches are occupied ...

$$X_{i,t+1}^{(n)} \stackrel{d}{=} \text{Bin}\left(X_{i,t}^{(n)} + \text{Bin}\left(1 - X_{i,t}^{(n)}, c\left(\frac{1}{n} \sum_{j=1}^n X_{j,t}^{(n)}\right)\right), s_i\right).$$

Our approach - Point Processes

Treat the collection of patch survival probabilities and those of *occupied patches* at time t as point processes on $[0, 1]$.

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Define sequences (σ_n) and $(\mu_{n,t})$ of random measures by

$$\sigma_n(B) = \#\{s_i \in B\}/n, \quad B \in \mathcal{B}([0, 1]),$$

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Think of σ as being the distribution of survival probabilities. In the earlier simulation σ was a Beta(25.2, 19.8) distribution.

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Our approach - Point Processes

Equivalently, we may define (σ_n) and $(\mu_{n,t})$ by

$$\int h(s)\sigma_n(ds) = \frac{1}{n} \sum_{i=1}^n h(s_i)$$
$$\int h(s)\mu_{n,t}(ds) = \frac{1}{n} \sum_{i=1}^n X_{i,t}^{(n)} h(s_i),$$

for h in $C^+([0, 1])$, the class of continuous functions that map $[0, 1]$ to $[0, \infty)$.

Our approach - Point Processes

Equivalently, we may define (σ_n) and $(\mu_{n,t})$ by

$$\int h(s)\sigma_n(ds) = \frac{1}{n} \sum_{i=1}^n h(s_i)$$

$$\int h(s)\mu_{n,t}(ds) = \frac{1}{n} \sum_{i=1}^n X_{i,t}^{(n)} h(s_i),$$

for h in $C^+([0, 1])$, the class of continuous functions that map $[0, 1]$ to $[0, \infty)$. For example ($h \equiv 1$),

$$\int \mu_{n,t}(ds) = \frac{1}{n} \sum_{i=1}^n X_{i,t}^{(n)} \quad (\text{proportion occupied}).$$

A measure-valued difference equation

Theorem Suppose that $\sigma_n \xrightarrow{d} \sigma$ and $\mu_{n,0} \xrightarrow{d} \mu_0$ for some non-random measures σ and μ_0 . Then, $\mu_{n,t} \xrightarrow{d} \mu_t$ for all $t = 1, 2, \dots$, where μ_t is defined by the following recursion: for $h \in C^+([0, 1])$,

$$\int h(s) \mu_{t+1}(ds) = (1 - c_t) \int sh(s) \mu_t(ds) + c_t \int sh(s) \sigma(ds),$$

where $c_t = c(\mu_t([0, 1])) = c(\int \mu_t(ds))$.

Homogeneous case

When $\bar{\sigma}^{(k)} = (\bar{\sigma}^{(1)})^k$ for all k , that is the patch survival probabilities are the same, then it is possible to simplify

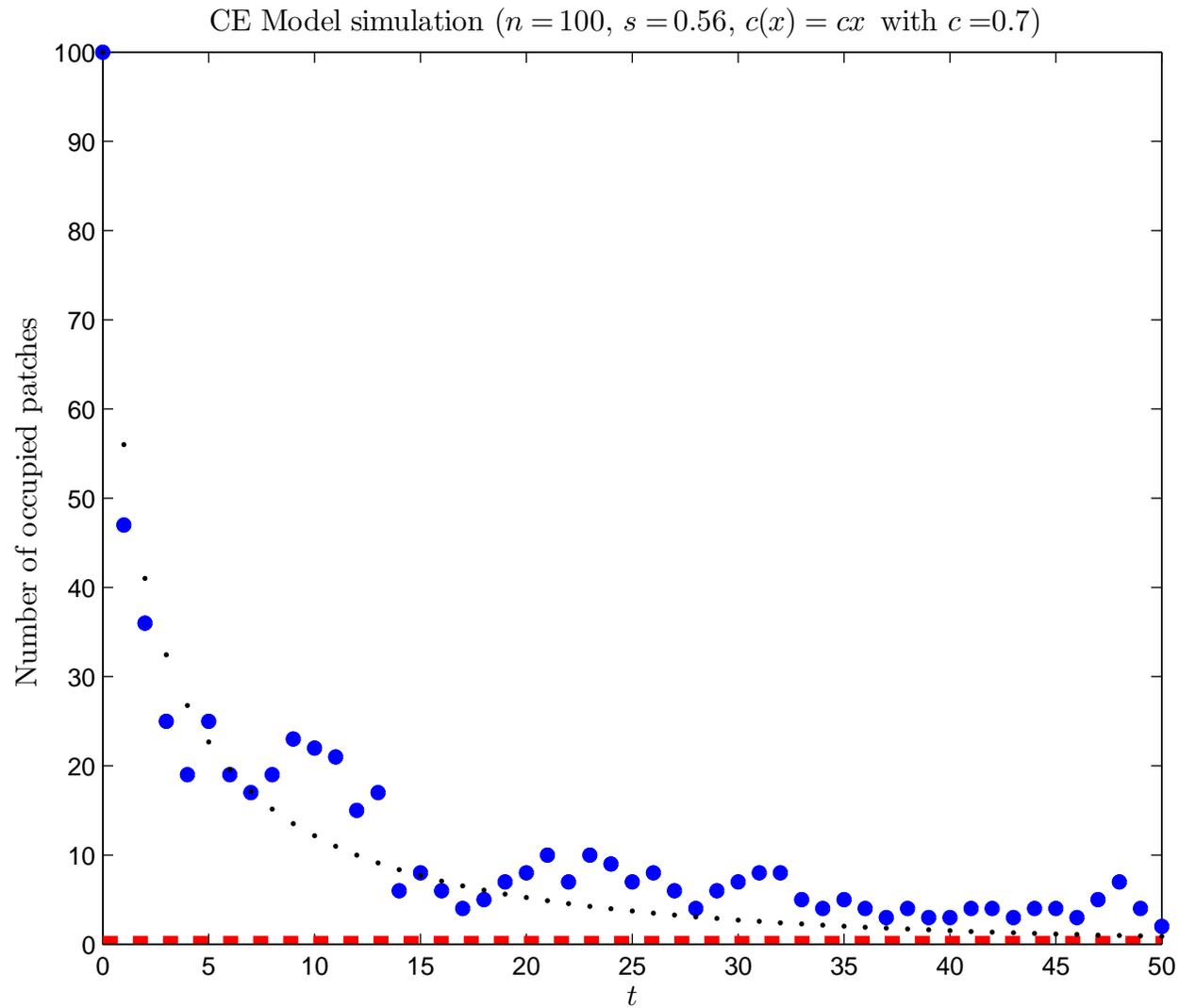
$$\bar{\mu}_{t+1}^{(k)} = (1 - \bar{\mu}_t^{(0)}) \bar{\mu}_t^{(k+1)} + \bar{\mu}_t^{(0)} \bar{\sigma}^{(k+1)},$$

We can show by induction that $\mu_t^{(k)} = (\bar{\sigma}^{(1)})^k x_t$, where

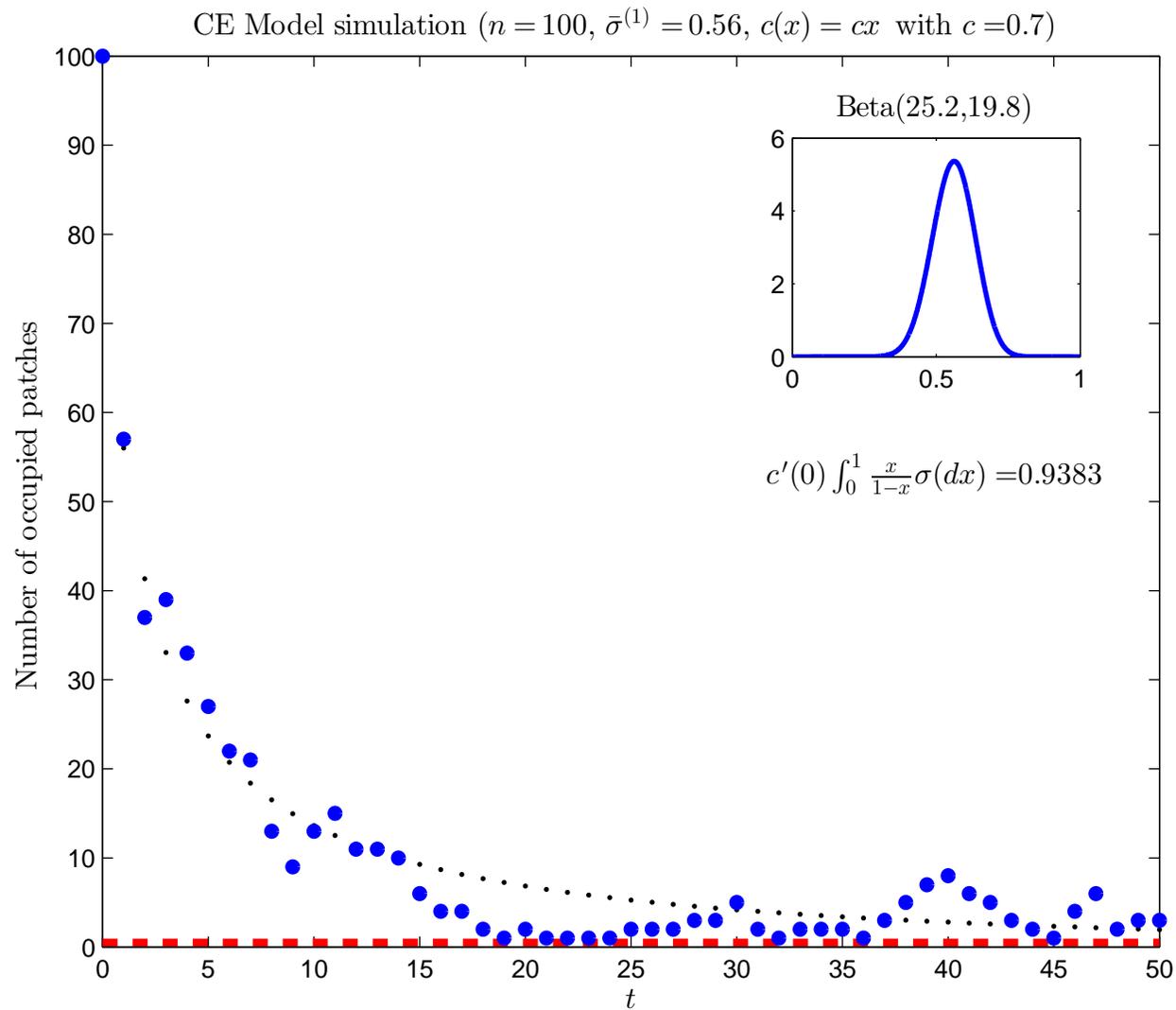
$$x_{t+1} = \bar{\sigma}^{(1)} (x_t + (1 - x_t) c(x_t)).$$

Compare this with the earlier result.

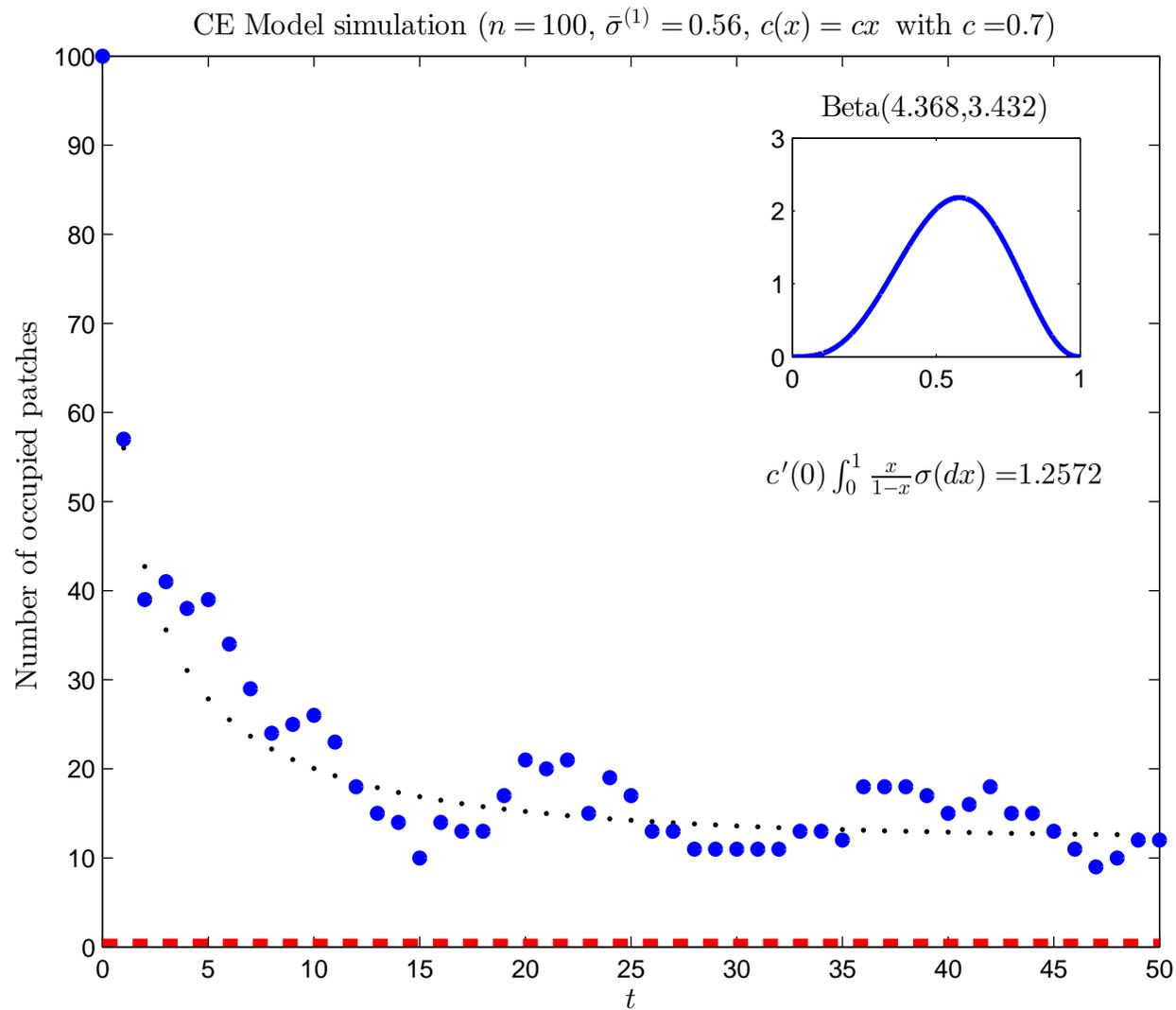
CE Model (homogeneous) - Evanescence



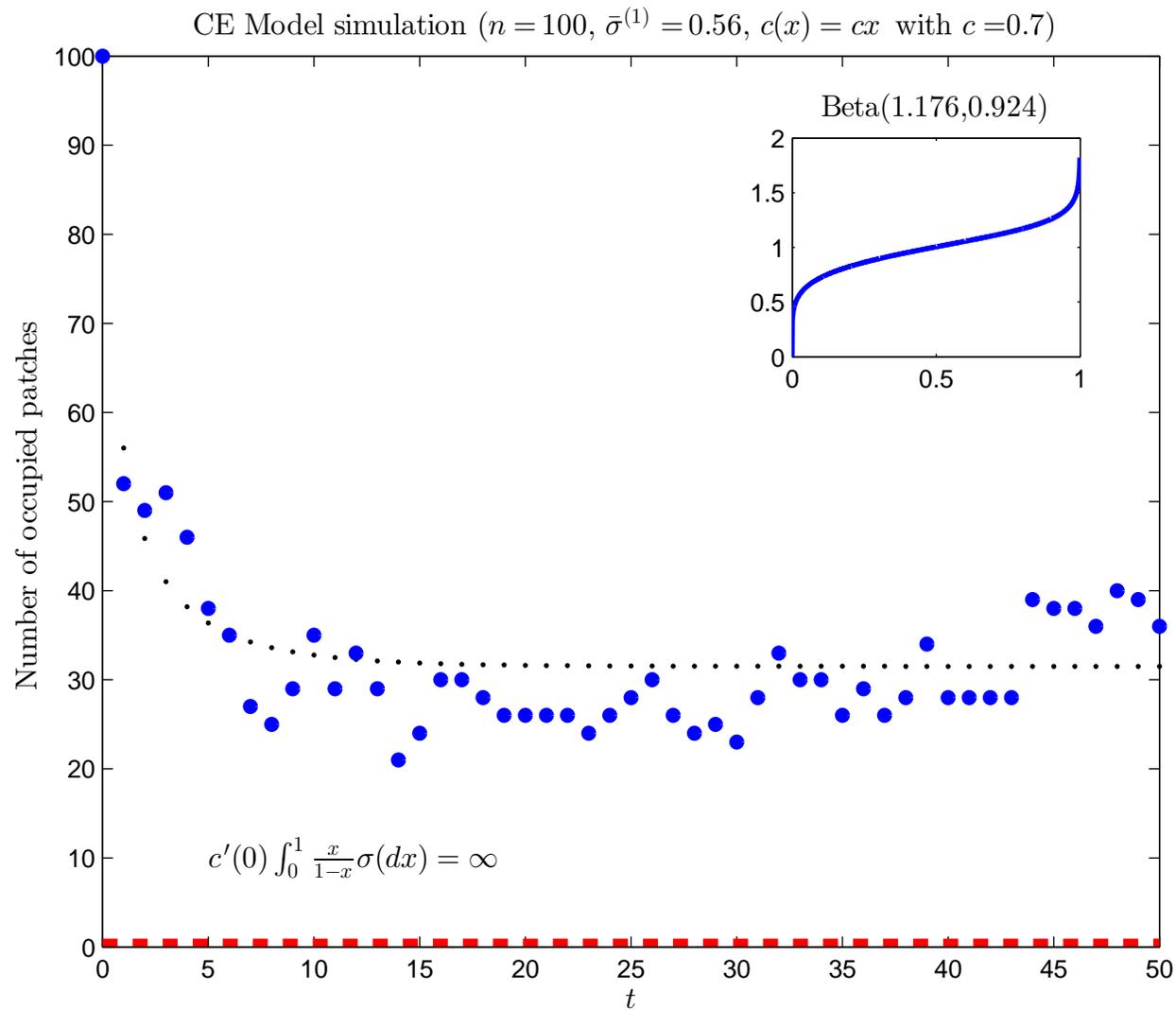
CE Model - Evanescence



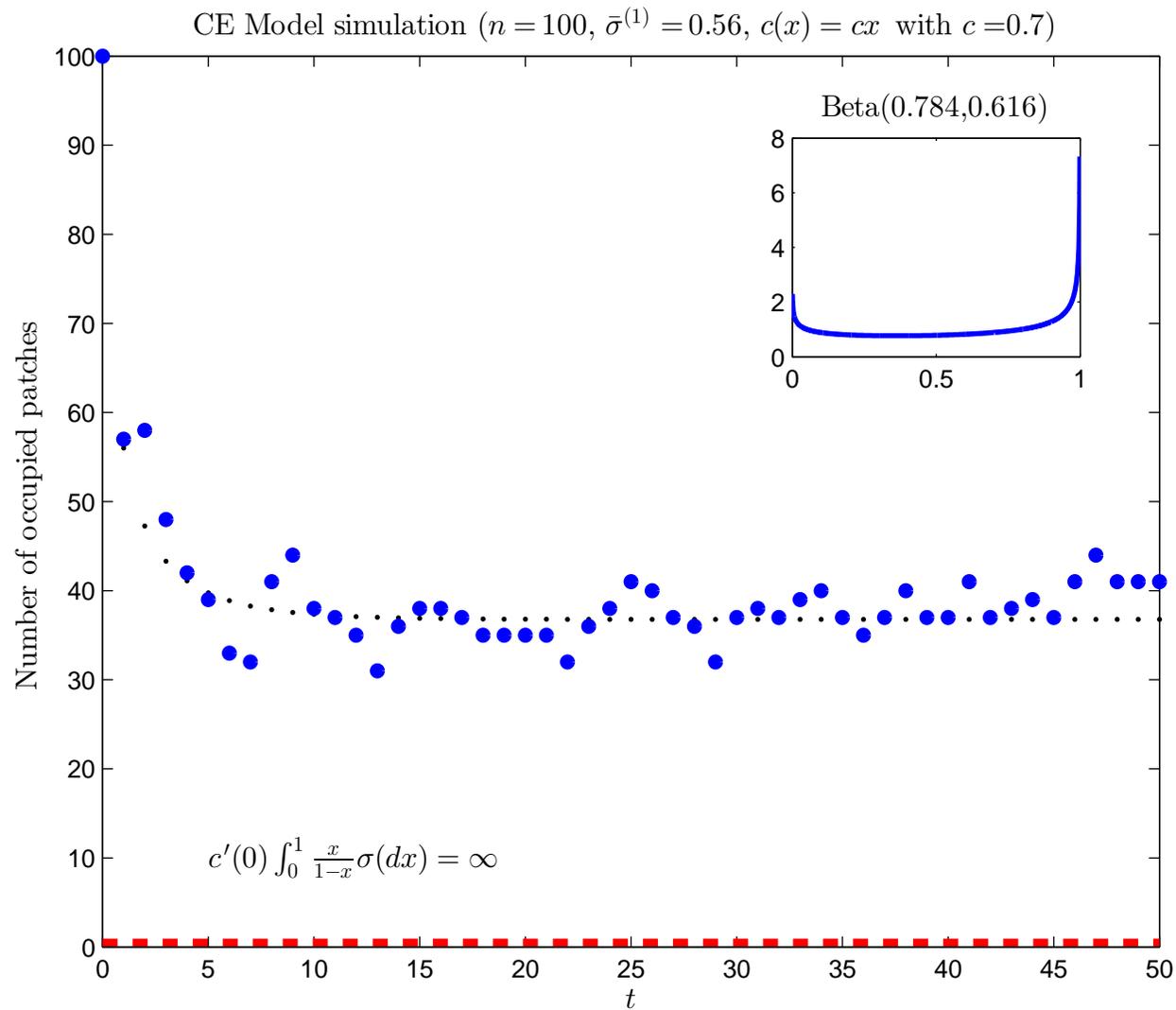
CE Model - Quasi stationarity



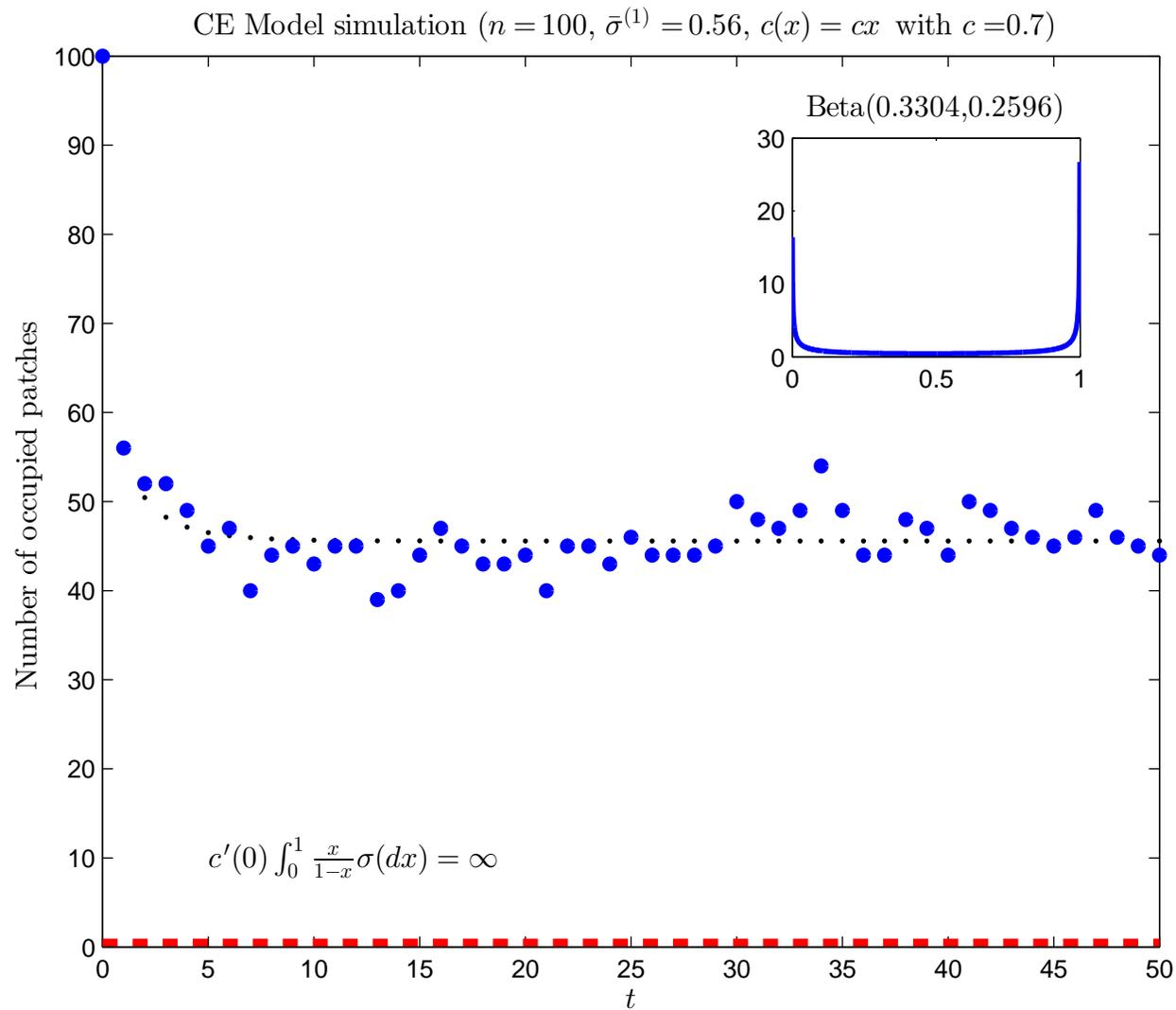
CE Model - Quasi stationarity



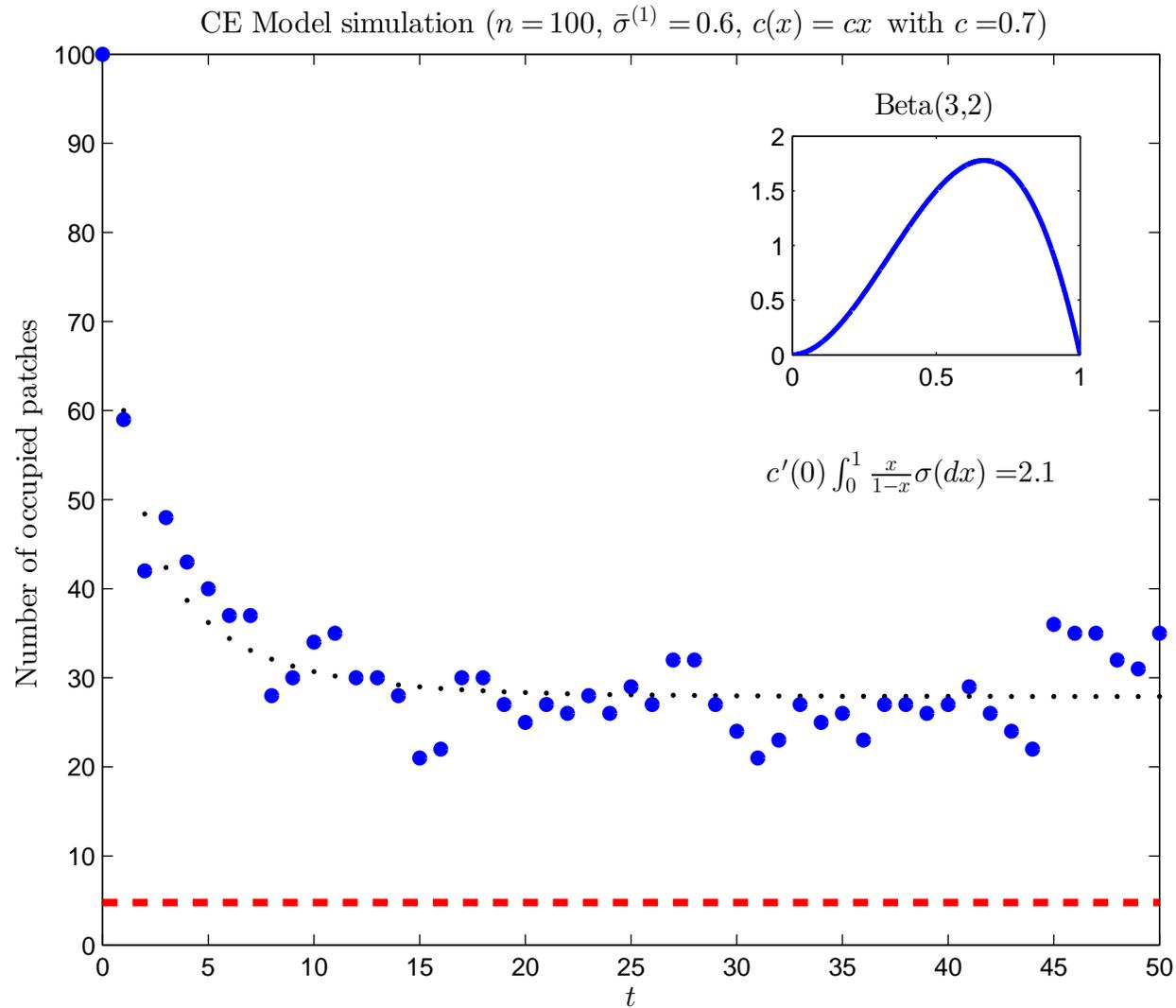
CE Model - Quasi stationarity



CE Model - Quasi stationarity



CE Model - Quasi stationarity



Our recursion is

$$\int h(s)\mu_{t+1}(ds) = (1 - c_t) \int sh(s)\mu_t(ds) + c_t \int sh(s)\sigma(ds).$$

Extra - equilibria

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$$\int h(s)\mu_{t+1}(ds) = (1 - c_t) \int sh(s)\mu_t(ds) + c_t \int sh(s)\sigma(ds).$$

Let \mathcal{M} be the set of measures that are absolutely continuous with respect to σ and whose Radon-Nikodym derivative is bounded by 1, σ - a.e.

We shall be interested in the behaviour of solutions to our recursion starting with $\mu_0 \in \mathcal{M}$.

Extra - equilibria

"Differentiating" with respect to σ , we see that our recursion can be written

$$\frac{\partial \mu_{t+1}}{\partial \sigma} = s \frac{\partial \mu_t}{\partial \sigma} + sc_t \left(1 - \frac{\partial \mu_t}{\partial \sigma} \right).$$

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It will be clear that $\mu_0 \in \mathcal{M}$ implies that $\mu_t \in \mathcal{M}$ for all t .

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It will be clear that $\mu_0 \in \mathcal{M}$ implies that $\mu_t \in \mathcal{M}$ for all t .

Furthermore, a measure $\mu_\infty \in \mathcal{M}$ will be an equilibrium point of our recursion if it satisfies

$$\frac{\partial \mu_\infty}{\partial \sigma} = s \frac{\partial \mu_\infty}{\partial \sigma} + s c_\infty \left(1 - \frac{\partial \mu_\infty}{\partial \sigma} \right),$$

where $c_\infty = c(\mu_\infty([0, 1]))$.

Extra - equilibria

Theorem Suppose that $c(0) = 0$ and $c'(0) < \infty$. Let ψ^* be a solution to the equation

$$\psi = R_\sigma(\psi) := \int \frac{sc(\psi)}{1-s+sc(\psi)} \sigma(ds). \quad (1)$$

The fixed points of our recursion are given by

$$\mu_\infty(ds) = \frac{sc(\psi^*)}{1-s+sc(\psi^*)} \sigma(ds).$$

Equation (1) has the unique solution $\psi^* = 0$ if and only if

$$c'(0) \int \frac{s}{1-s} \sigma(ds) \leq 1.$$

Otherwise, there are two solutions, one of which is $\psi^* = 0$.

Theorem If $\psi^* = 0$ is the only solution to Equation (1), then, for all $\mu_0 \in \mathcal{M}$, $\mu_t \rightarrow 0$. If Equation (1) has a non-zero solution, then, for all $\mu_0 \in \mathcal{M}$ such that $\int \mu_{0,j}(ds) > 0$ for some j , $\mu_t \rightarrow \mu_\infty$.