

Population Models: Part II

Markov Chains and Diffusion Approximations

Phil. Pollett

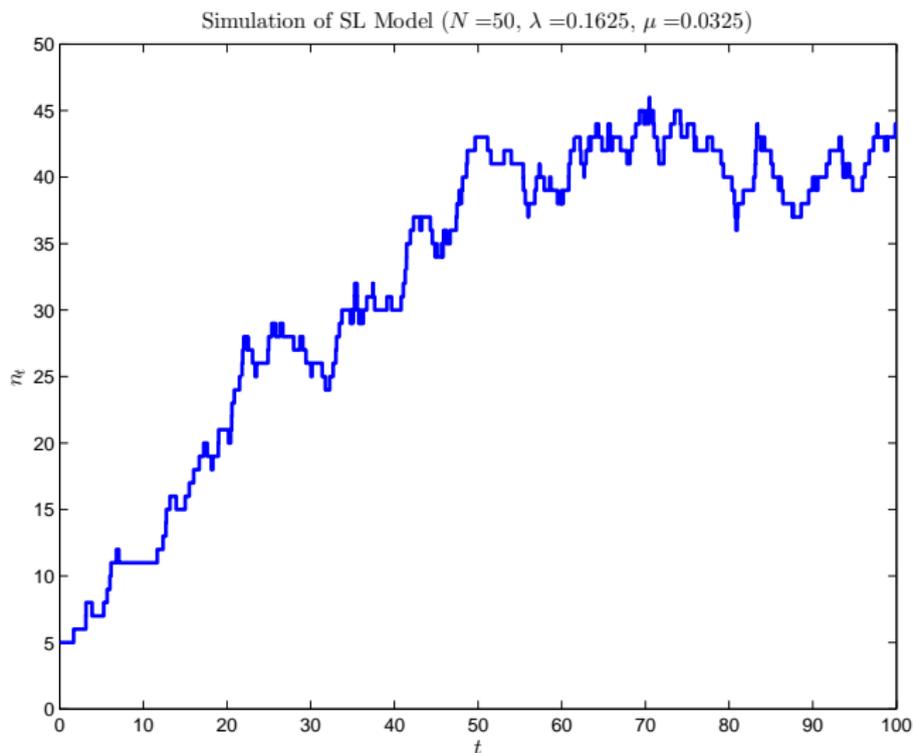
UQ School of Mathematics and Physics

Mathematics Enrichment Seminars

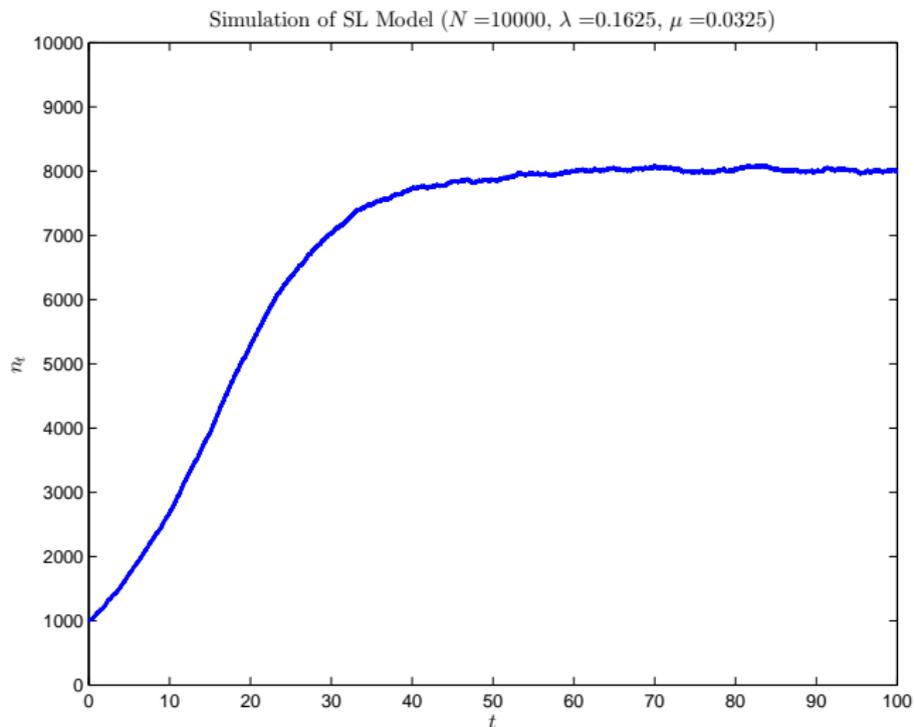
24 August 2016



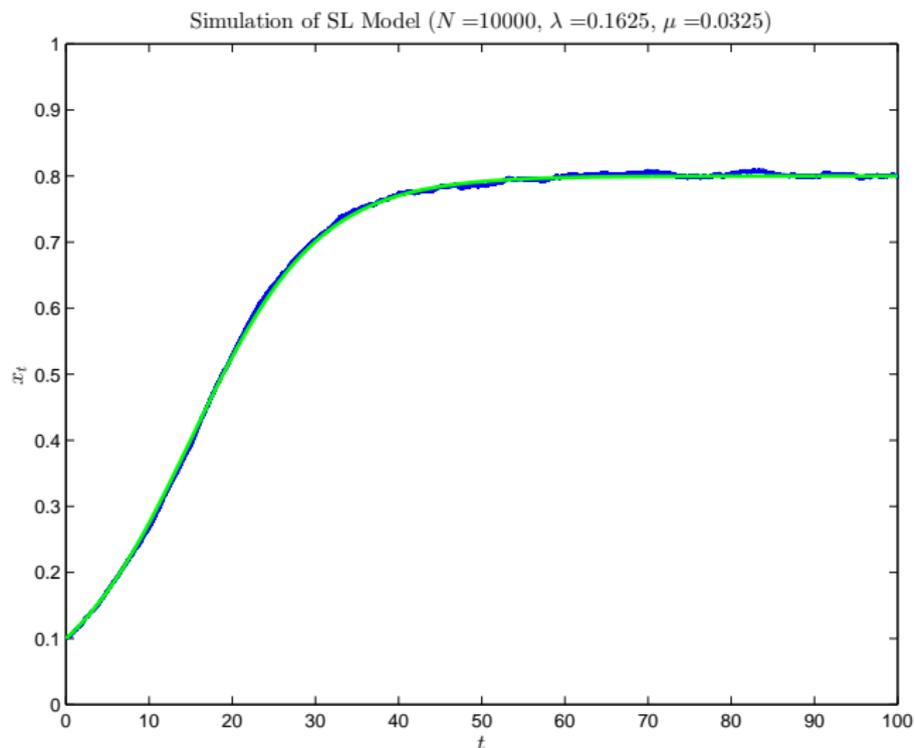
Executive Summary - Simulation of the SL Model ($N = 50$)



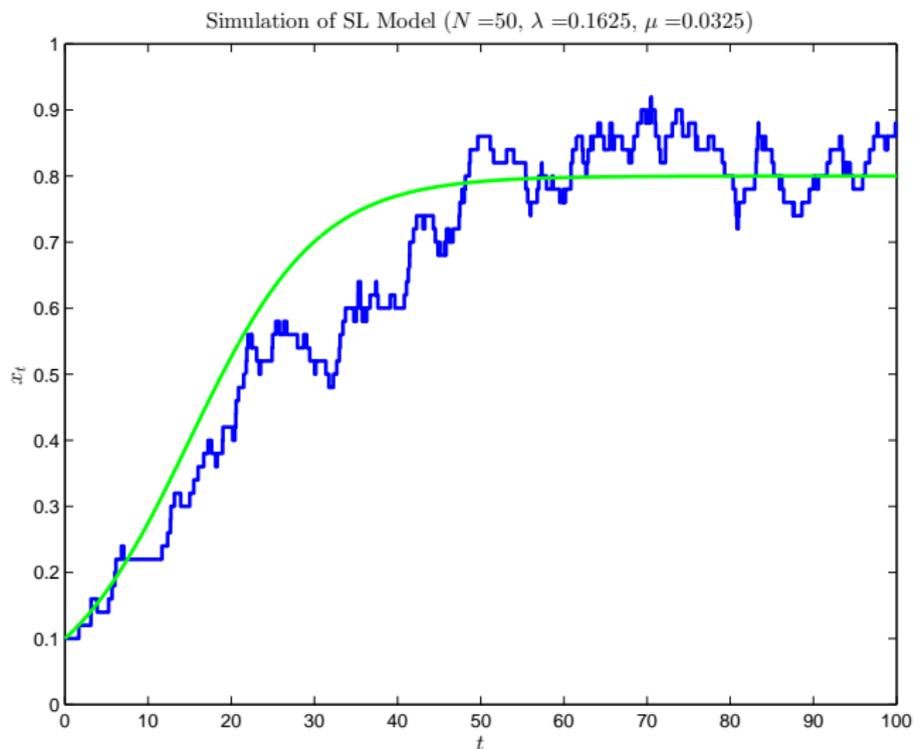
Executive Summary - Simulation of the SL Model (N large)



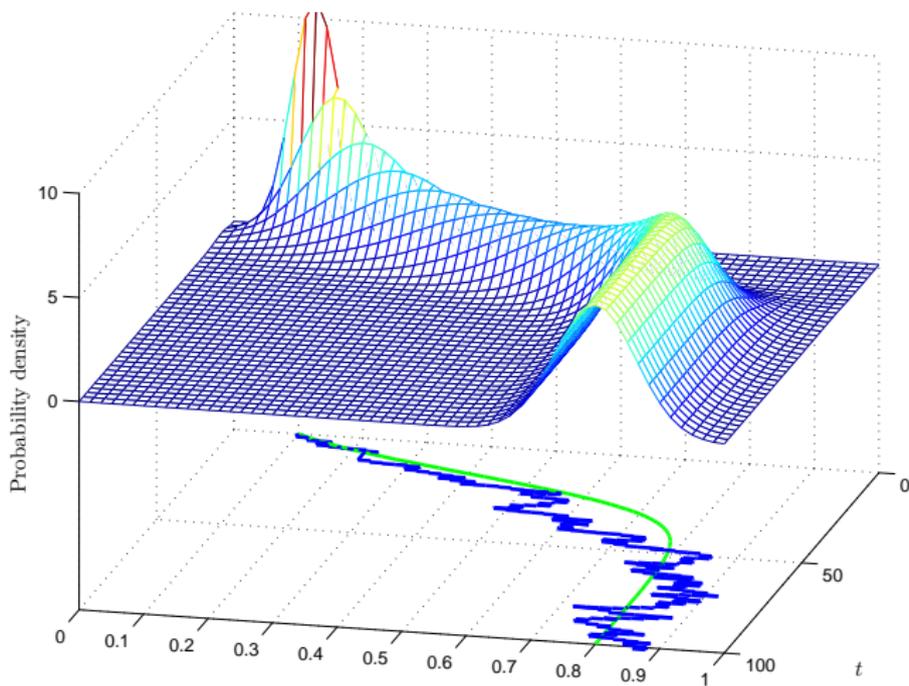
Executive Summary - Solution to deterministic model



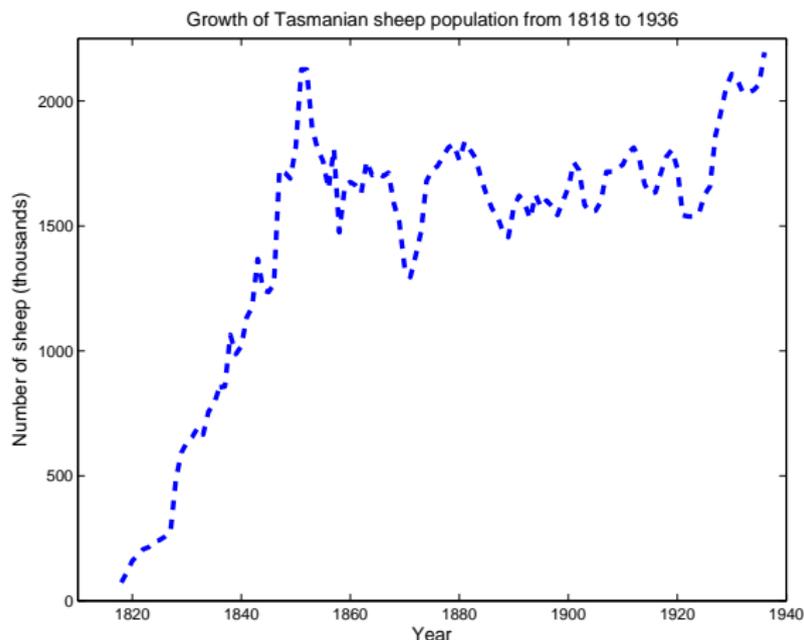
Executive Summary - Solution to deterministic model



Executive Summary - Normal approximation



Part I Recap: Sheep in Tasmania



Davidson, J. (1938) On the growth of the sheep population in Tasmania. *Trans. Roy. Soc. Sth. Austral.* 62, 342–346.

Part I Recap: The Verhulst-Pearl Model (Logistic Model)

We started with a simple deterministic model for n_t , the number in our population at time t :

$$\frac{dn}{dt} = rn \left(1 - \frac{n}{K}\right),$$

with r being the growth rate with unlimited resources and K being the “natural” population size (the *carrying capacity*).

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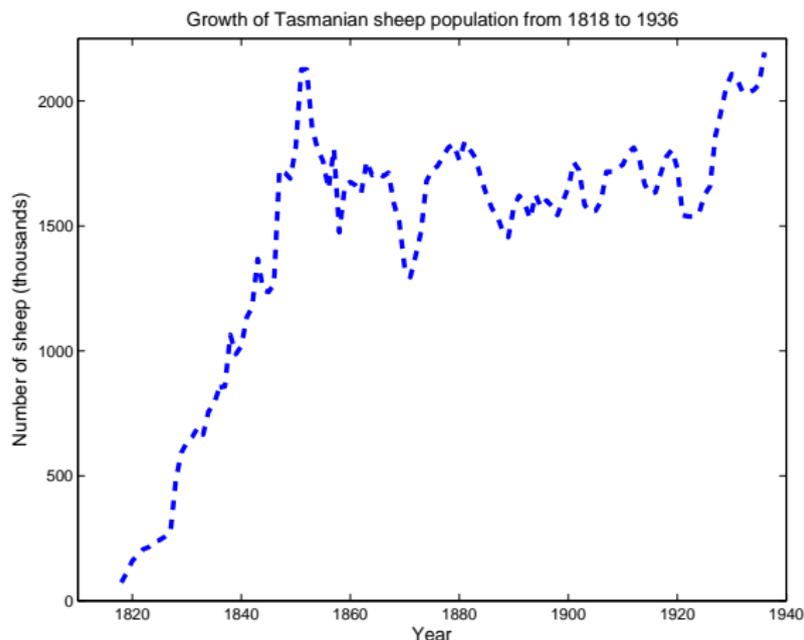
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Integration gives

$$n_t = \frac{K}{1 + \left(\frac{K-n_0}{n_0}\right) e^{-rt}}, \quad t \geq 0.$$

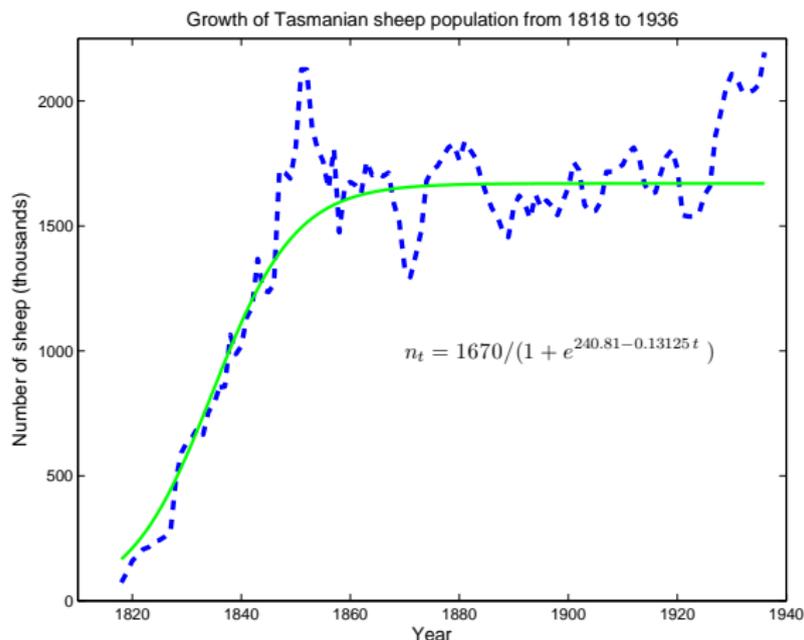
(As covered in MATH1052!)

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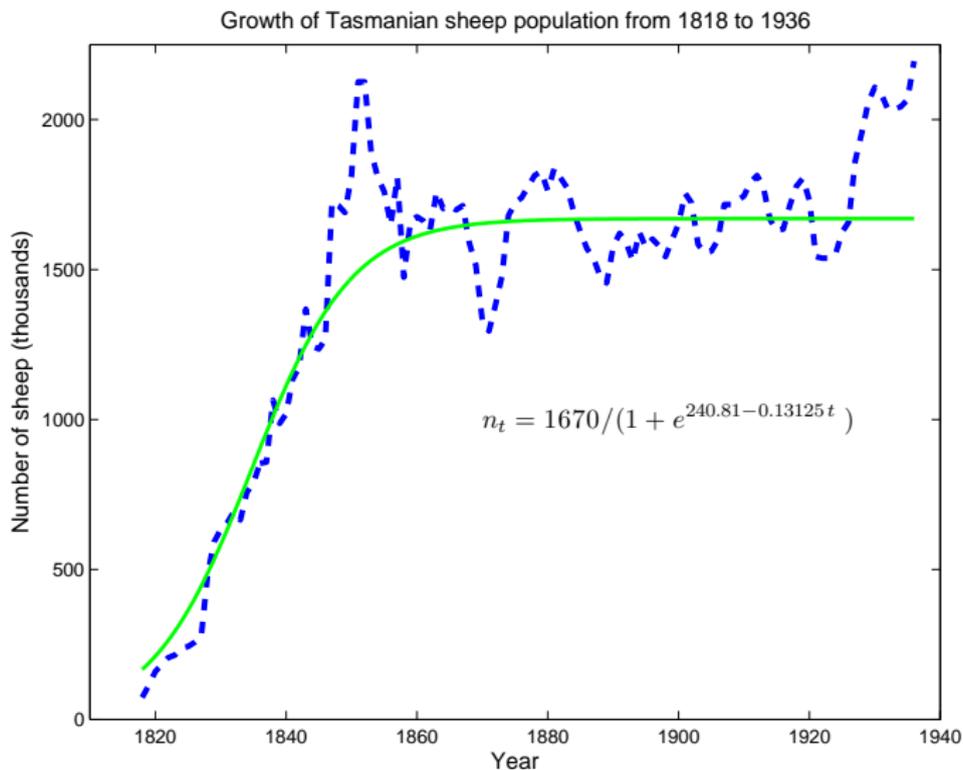
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In Matlab ...

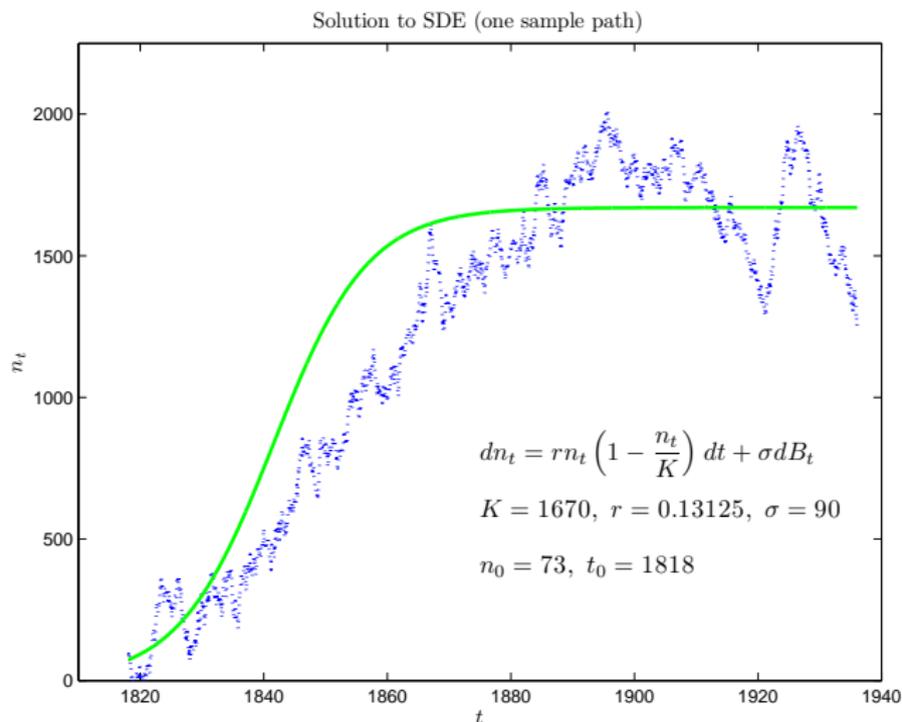
```
n = n + r*n*(1-n/K)*h + sigma*sqrt(h)*randn;
```



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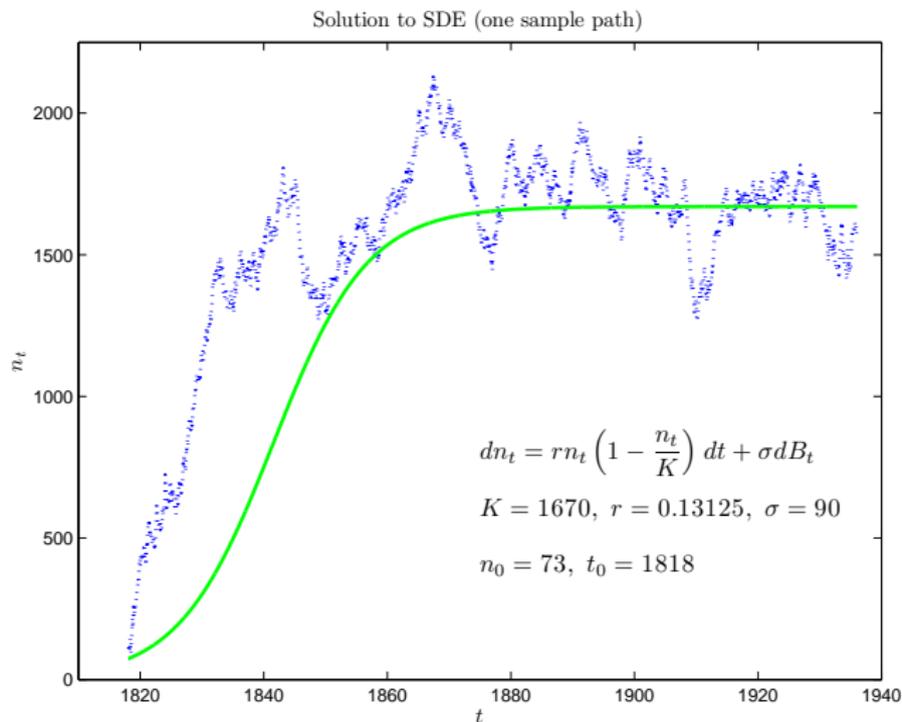
Part I Recap: Solution to SDE (Run 1)



(Solution to the deterministic model is in green)



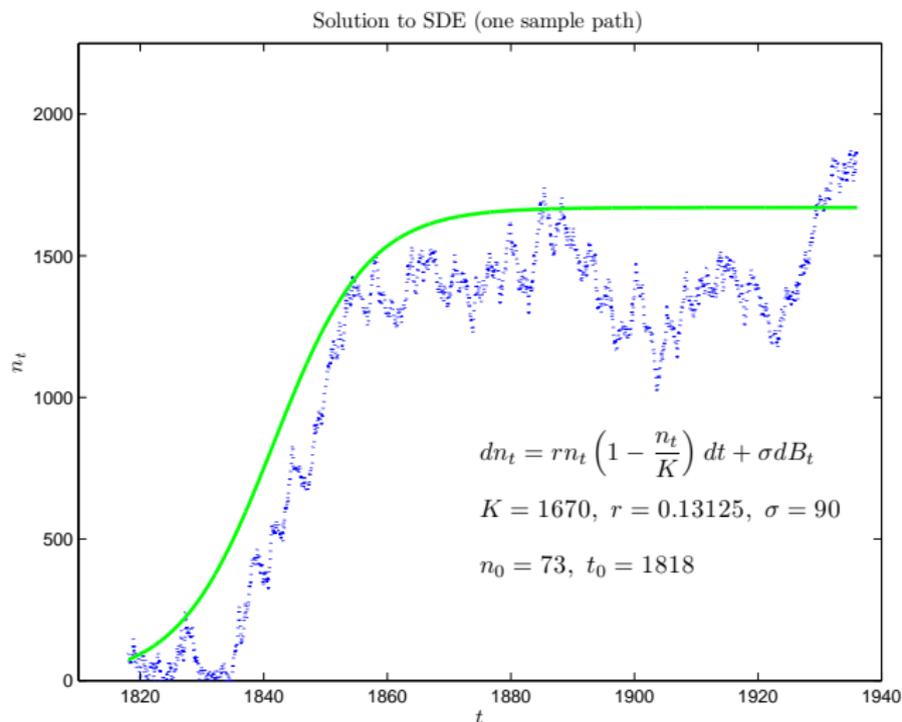
Part I Recap: Solution to SDE (Run 2)



(Solution to the deterministic model is in green)

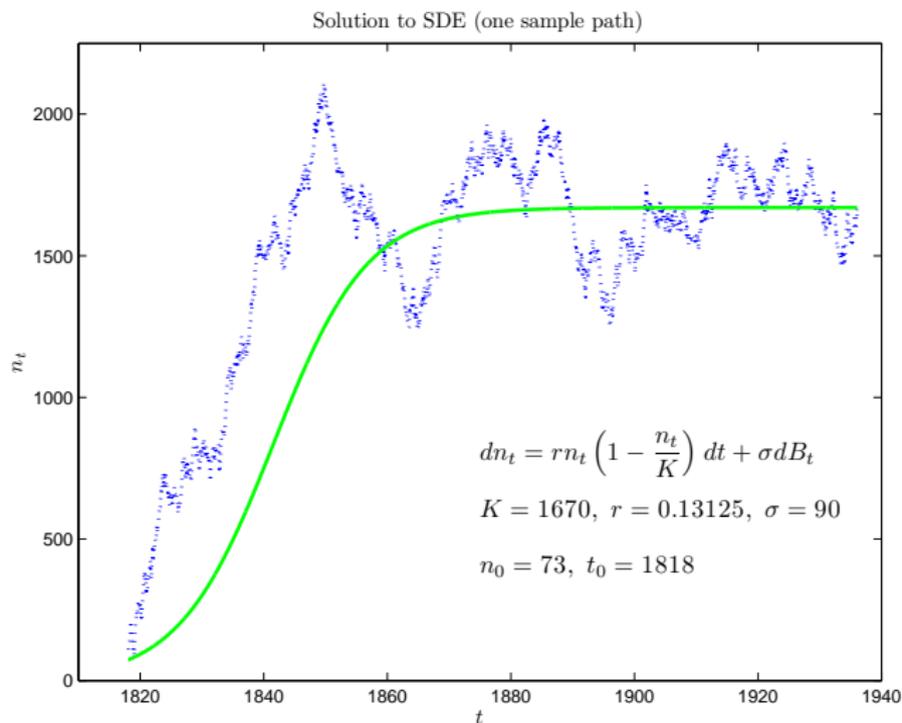


Part I Recap: Solution to SDE (Run 3)



(Solution to the deterministic model is in green)

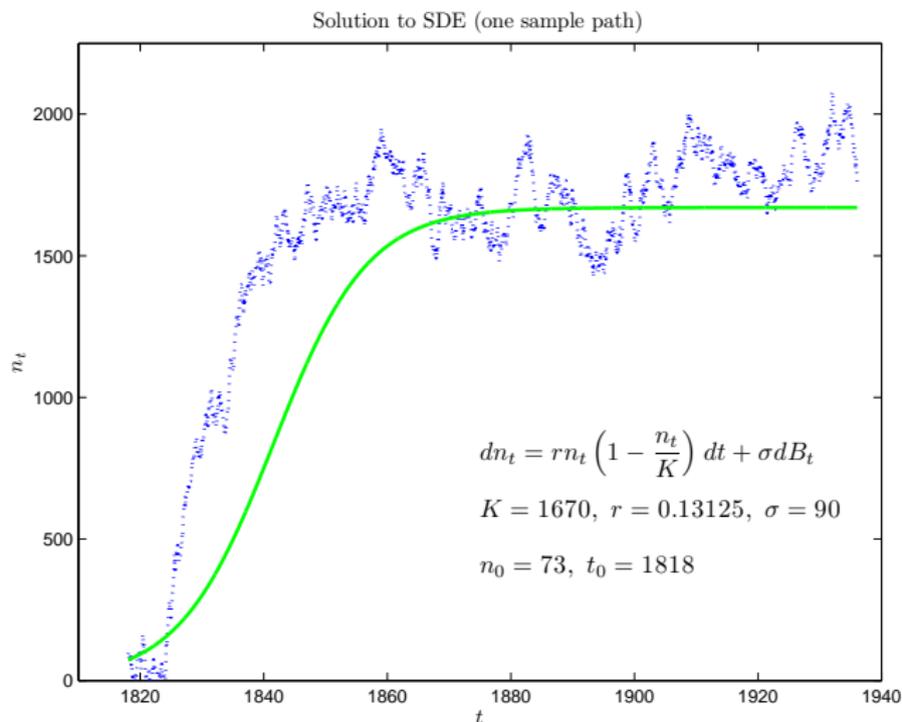
Part I Recap: Solution to SDE (Run 4)



(Solution to the deterministic model is in green)

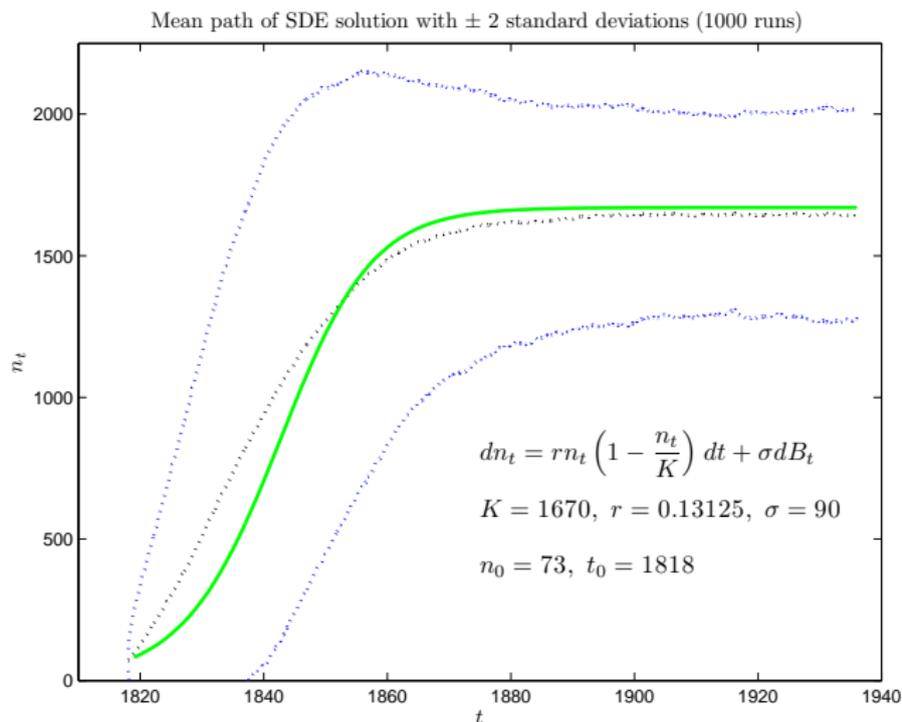


Part I Recap: Solution to SDE (Run 5)



(Solution to the deterministic model is in green)

Part I Recap: Solution to SDE



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My last slide from Part I

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Arguably a better approach is to use a continuous-time Markov chain to model n_t .

This will be dealt with in Part II.

A different approach - a continuous time stochastic model

Let's start from scratch specifying a stochastic model with variation being an inherent ingredient: a *Markovian model*.

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We will suppose that n_t evolves (in continuous time) as a birth-death process with transitions

$$n \rightarrow n + 1 \quad \text{at rate} \quad \frac{\lambda}{N} n (N - n) \quad (\text{birth})$$

$$n \rightarrow n - 1 \quad \text{at rate} \quad \mu n \quad (\text{death})$$

where $\mu (> 0)$ is the per-capita death rate and $\lambda (> 0)$ is the birth rate (per-capita when the population is small). N is the *population ceiling*; n_t now takes values in the set $S = \{0, 1, \dots, N\}$.

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In the context of general population modelling it is called the *Stochastic Logistic Model* (for reasons that will become apparent soon), and can be traced back to William Feller:

Feller, W. (1939) Die grundlagen der volterraschen theorie des kampfes ums dasein in wahrscheinlichkeitsteoretischer behandlung. *Acta Biotheoretica* 5, 11–40.



The Stochastic Logistic Model

In the epidemiological context it is known as the *SIS (Susceptible-Infectious-Susceptible) Model*, and was introduced by Weiss and Dishon to study infections, in a closed population of N individuals, that do not confer any long lasting immunity.

Weiss, G.H. and Dishon, M. (1971) On the asymptotic behavior of the stochastic and deterministic models of an epidemic. *Mathematical Biosciences* 11, 261–265.

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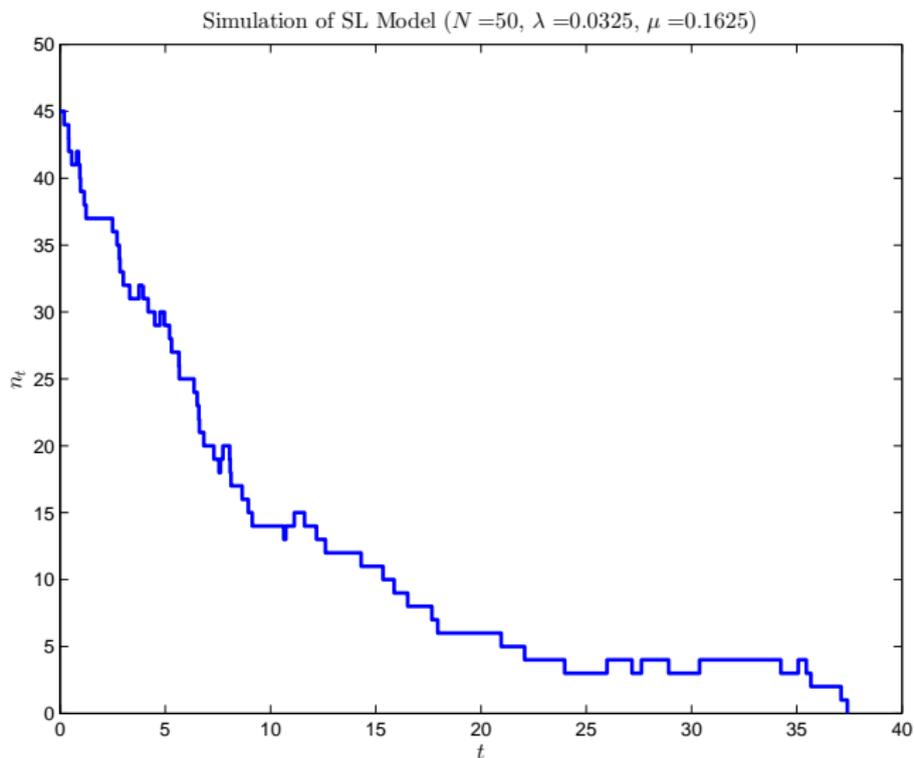
The transitions have the interpretation

$$n \rightarrow n + 1 \quad \text{at rate} \quad \frac{\lambda}{N} n (N - n) \quad \text{(infection)}$$

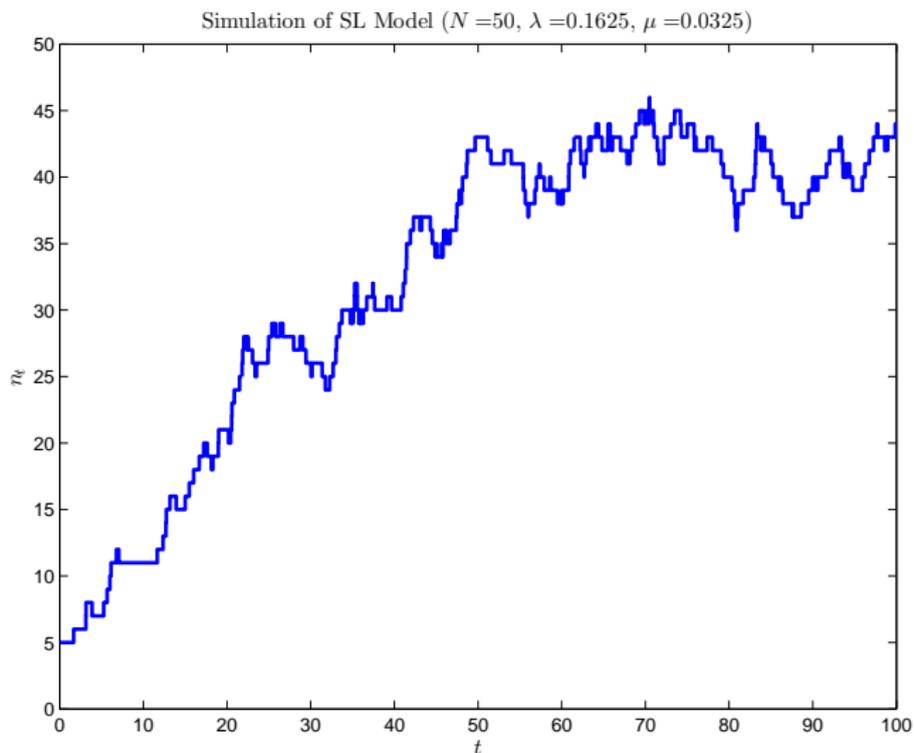
$$n \rightarrow n - 1 \quad \text{at rate} \quad \mu n \quad \text{(recovery)}$$

with μ being the per-capita recovery rate and λ being the per-proximate encounter transmission rate.

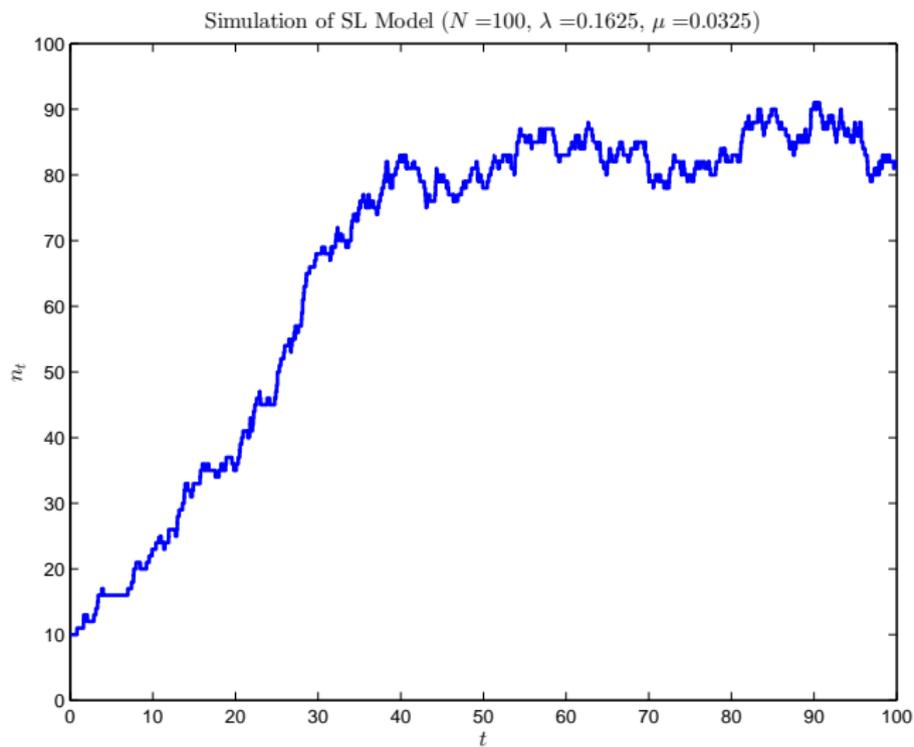
Simulation of the SL Model - extinction



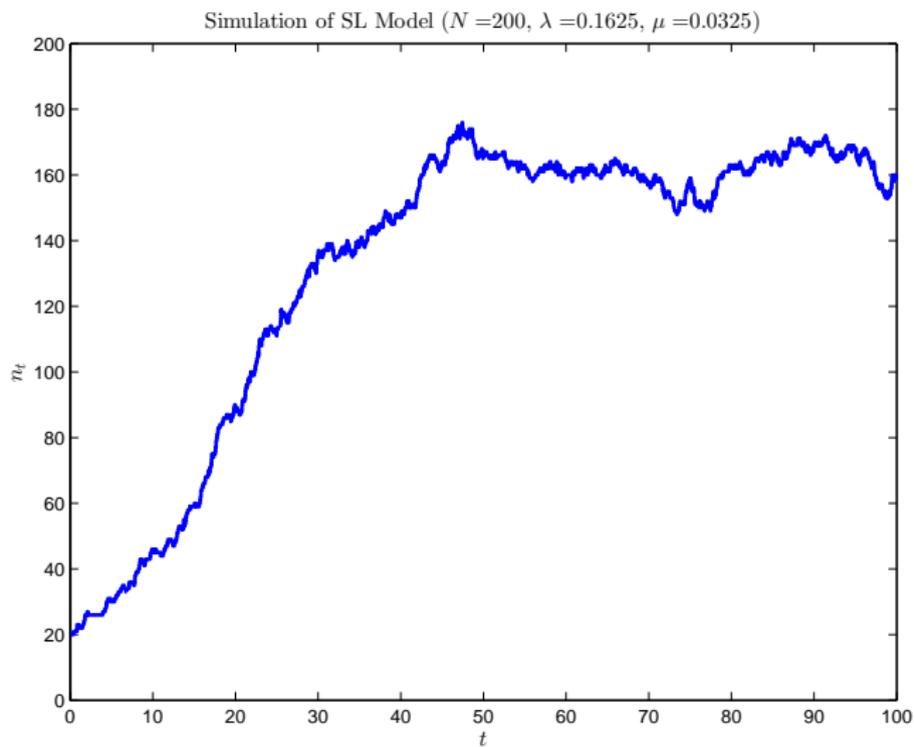
Simulation of the SL Model - persistence



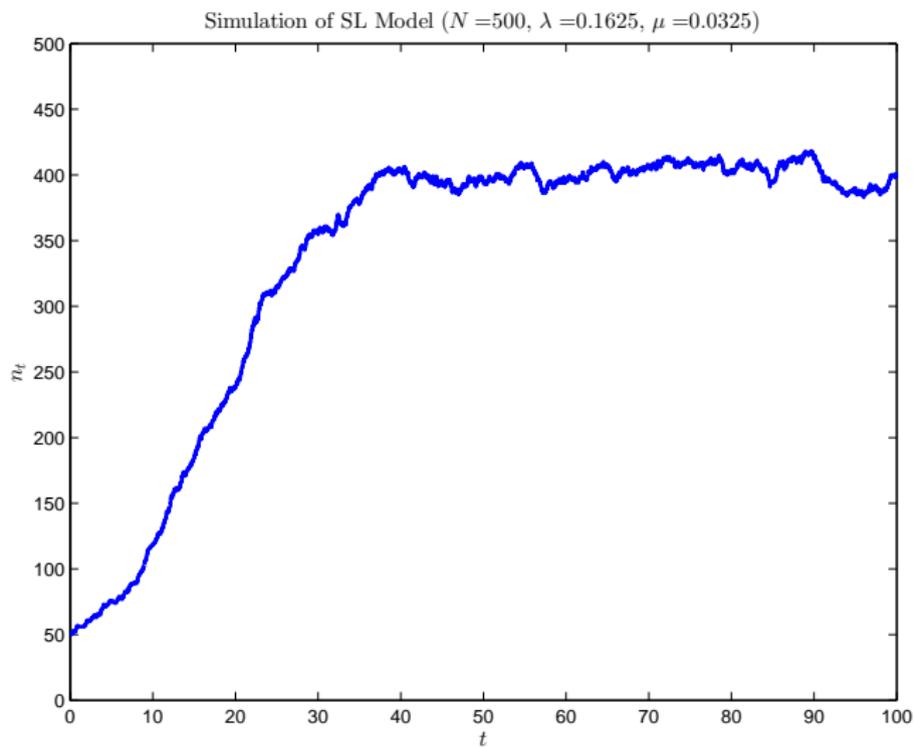
Simulation of the SL Model - N growing



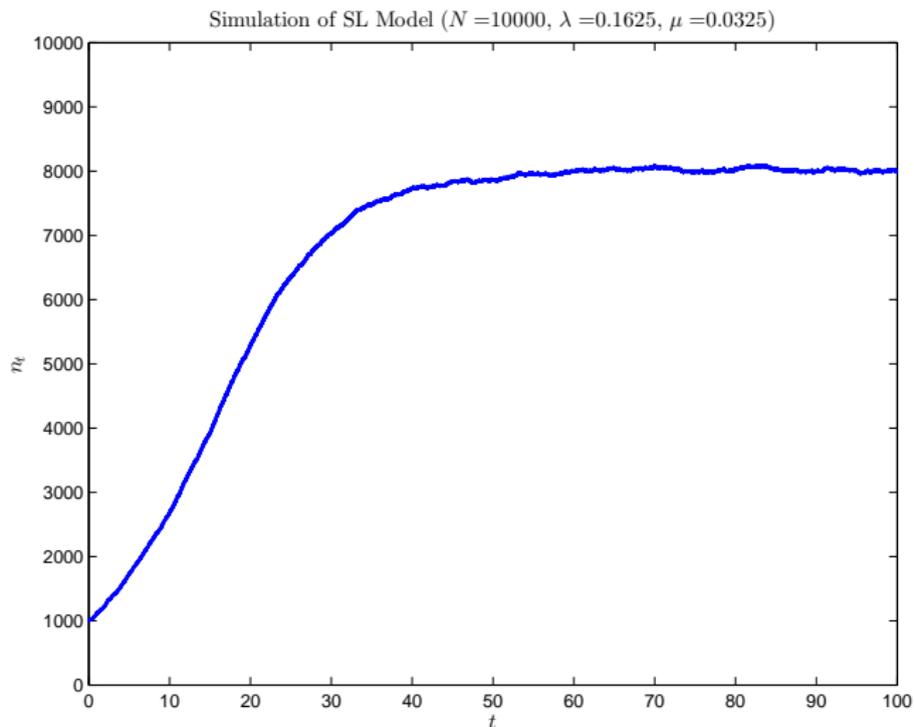
Simulation of the SL Model - N growing



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The key to answering Question 1 is *density dependence*, a property that is shared by the deterministic and stochastic logistic models.

Density dependence

The Verhulst-Pearl Model

$$\frac{dn}{dt} = rn \left(1 - \frac{n}{K}\right)$$

can be written

$$\frac{1}{N} \frac{dn}{dt} = r \frac{n}{N} \left(1 - \frac{N}{K} \frac{n}{N}\right).$$

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So, letting $x_t = n_t/N$ be the “population density” at time t , we see that

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BTW: How could ODEs possibly be useful for modelling integer-valued quantities such as a population size? *Scaling like this helps explain why.*

Density dependence in Markovian models

A stochastic process $(n_t, t \geq 0)$ in continuous time taking values in $S \subseteq \mathbb{Z}^k$, called a *Markov chain*, is characterized by its transition rates $Q = (q_{nm}, n, m \in S)$; q_{nm} , for $m \neq n$, represents the rate at which the process moves from state n to state m .

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$$q_{n, n+l} = N f_l \left(\frac{n}{N} \right), \quad l \neq 0, \quad l \in \mathbb{Z}^k.$$

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The mean path of a density dependent Markovian model

Consider the *forward equations* for the *state probabilities* $p_n(t) := \Pr(n_t = n)$ (in statistical mechanics, the *master equation*):

$$p'_n(t) = -q_n p_n(t) + \sum_{m \neq n} p_m(t) q_{mn}, \quad n \in S,$$

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So, if $q_{n,n+\ell} = N f_\ell(n/N)$ (density dependence), then

$$\begin{aligned} \frac{d}{dt} \mathbb{E}(n_t) &= - \sum_n \sum_{\ell \neq 0} N f_\ell(n/N) n p_n(t) + \sum_m p_m(t) \sum_{\ell \neq 0} (m + \ell) N f_\ell(m/N) \\ &= \sum_m p_m(t) N \sum_{\ell \neq 0} \ell f_\ell(m/N) = N \mathbb{E} \left(\sum_{\ell \neq 0} \ell f_\ell(n_t/N) \right). \end{aligned}$$



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So, for an arbitrary density dependent Markovian model, we may write

$$\frac{d}{dt} \mathbb{E}(n_t) = N \mathbb{E} \left(F \left(\frac{n_t}{N} \right) \right),$$

where $F : E \rightarrow \mathbb{R}$ is given by

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Or, setting $X_t = n_t/N$ (the “*density process*”),

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$$\frac{d}{dt} \mathbb{E}(X_t) = F(\mathbb{E}(X_t)).$$

But, it's an obvious candidate for our (limiting, we hope) deterministic model for the population density.

The Stochastic Logistic Model is density dependent

For the Stochastic Logistic Model we have $S = \{0, 1, \dots, N\}$ with

$$q_{n,n+1} = \frac{\lambda}{N} n (N - n) = N\lambda \frac{n}{N} \left(1 - \frac{n}{N}\right) \quad \text{and} \quad q_{n,n-1} = \mu n = N\mu \frac{n}{N}.$$

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Therefore, $f_{+1}(x) = \lambda x (1 - x)$ and $f_{-1}(x) = \mu x$, $x \in E := [0, 1]$, and so

$$F(x) = \sum_{\ell \neq 0} \ell f_{\ell}(x) = f_{+1}(x) - f_{-1}(x) = \lambda x (1 - \rho - x), \quad x \in E,$$

where $\rho = \mu/\lambda$.

The Stochastic Logistic Model is density dependent

For the Stochastic Logistic Model we have $S = \{0, 1, \dots, N\}$ with

$$q_{n,n+1} = \frac{\lambda}{N} n (N - n) = N\lambda \frac{n}{N} \left(1 - \frac{n}{N}\right) \quad \text{and} \quad q_{n,n-1} = \mu n = N\mu \frac{n}{N}.$$

Therefore, $f_{+1}(x) = \lambda x (1 - x)$ and $f_{-1}(x) = \mu x$, $x \in E := [0, 1]$, and so

$$F(x) = \sum_{\ell \neq 0} \ell f_{\ell}(x) = f_{+1}(x) - f_{-1}(x) = \lambda x (1 - \rho - x), \quad x \in E,$$

where $\rho = \mu/\lambda$. Now compare $F(x)$ with the right-hand side of the Verhulst-Pearl Model for the density process:

$$\frac{dx}{dt} = rx \left(1 - \frac{x}{E}\right), \quad \text{where} \quad E = K/N. \quad (1)$$

If $K \sim \beta N$ for N large, so that $K/N \rightarrow \beta$, then we may identify β with $1 - \rho$ and r with $\lambda\beta$, and discover that (1) can be rewritten as $dx/dt = F(x)$.

What about convergence?

Recall that $(n_t, t \geq 0)$ is a continuous-time Markov chain taking values in $S \subseteq \mathbb{Z}^k$ with transition rates $Q = (q_{nm}, n, m \in S)$, and we have identified a quantity N , usually related to the size of the system being modelled.

The model is assumed to be *density dependent*: there is a subset E of \mathbb{R}^k and a continuous function $f : \mathbb{Z}^k \times E \rightarrow \mathbb{R}$, such that

$$q_{n,n+\ell} = Nf_\ell \left(\frac{n}{N} \right), \quad \ell \neq 0, \ell \in \mathbb{Z}^k.$$

We set $F(x) = \sum_{\ell \neq 0} \ell f_\ell(x)$, $x \in E$.

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Now formally define the *density process* $(X_t^{(N)})$ by $X_t^{(N)} = n_t/N$, $t \geq 0$. We hope that $(X_t^{(N)})$ becomes more deterministic as N gets large.

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To simplify the statement of results, I'm going to assume that the state space S is finite.

A law of large numbers

The following *functional law of large numbers* establishes convergence of the family $(X_t^{(N)})$ to the unique trajectory of an appropriate approximating deterministic model.

Theorem (Kurtz*) Suppose F is Lipschitz continuous¹. If $\lim_{N \rightarrow \infty} X_0^{(N)} = x_0$, then the density process $(X_s^{(N)})$ converges in probability uniformly on $[0, t]$ to (x_s) , the unique (deterministic) trajectory satisfying

$$\frac{d}{ds}x_s = F(x_s), \quad x_s \in E, \quad s \in [0, t].$$

Kurtz, T. (1970) Solutions of ordinary differential equations as limits of pure jump Markov processes, *J. of Appl. Probab.* 7, 49–58.

¹For some $M > 0$, $|F(x) - F(y)| \leq M|x - y|$ for all $x \in E$.

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(If S is an infinite set, we have the additional conditions $\sup_{x \in E} \sum_{\ell \neq 0} |\ell| f_\ell(x) < \infty$ and $\lim_{d \rightarrow \infty} \sum_{|\ell| > d} |\ell| f_\ell(x) = 0, x \in E.$)

¹For some $M > 0$, $|F(x) - F(y)| \leq M|x - y|$ for all $x \in E$.

A law of large numbers

Convergence in probability uniformly on $[0, t]$ means that, for every $\epsilon > 0$,

$$\lim_{N \rightarrow \infty} \Pr \left(\sup_{s \in [0, t]} |X_s^{(N)} - x_s| > \epsilon \right) = 0.$$

A law of large numbers

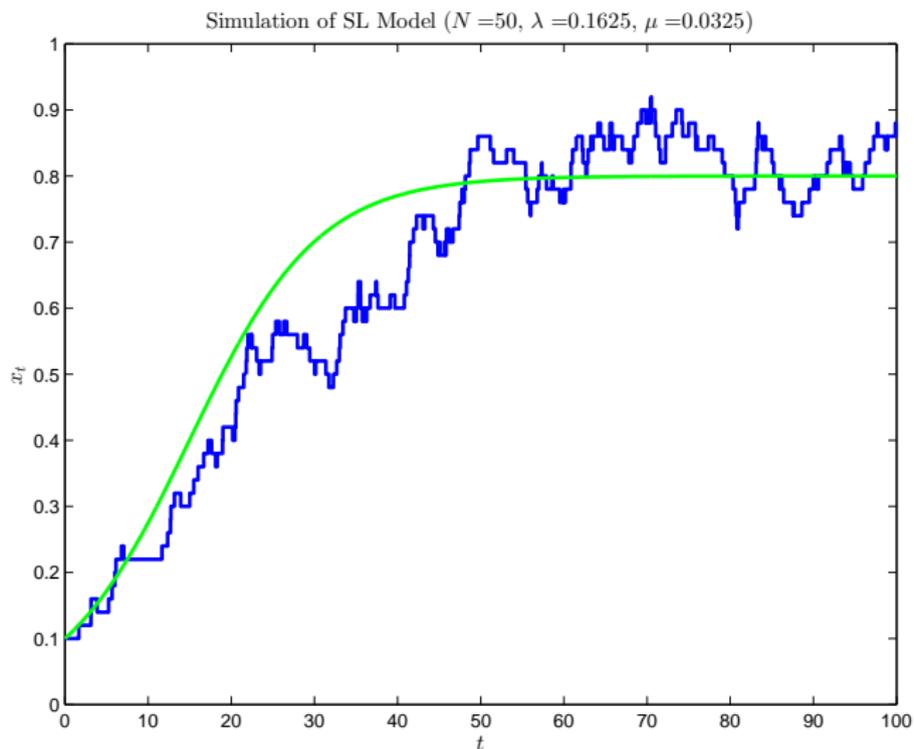
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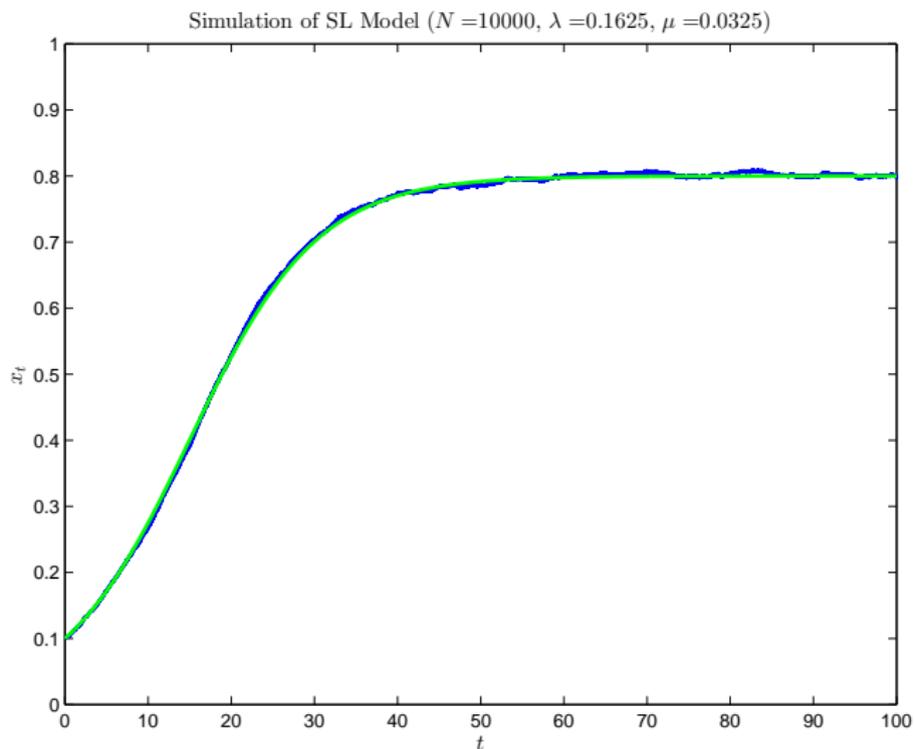
For the Stochastic Logistic Model, it is easy to check that $F(x) = \lambda x(1 - \rho - x)$ is Lipschitz continuous on $E = [0, 1]$. So, provided $X_0^{(N)} \rightarrow x_0$ as $N \rightarrow \infty$, the population density $(X_t^{(N)})$ converges (uniformly in probability on finite time intervals) to the solution (x_t) of the deterministic model

$$\frac{dx}{dt} = \lambda x(1 - \rho - x) \quad (x_t \in E).$$

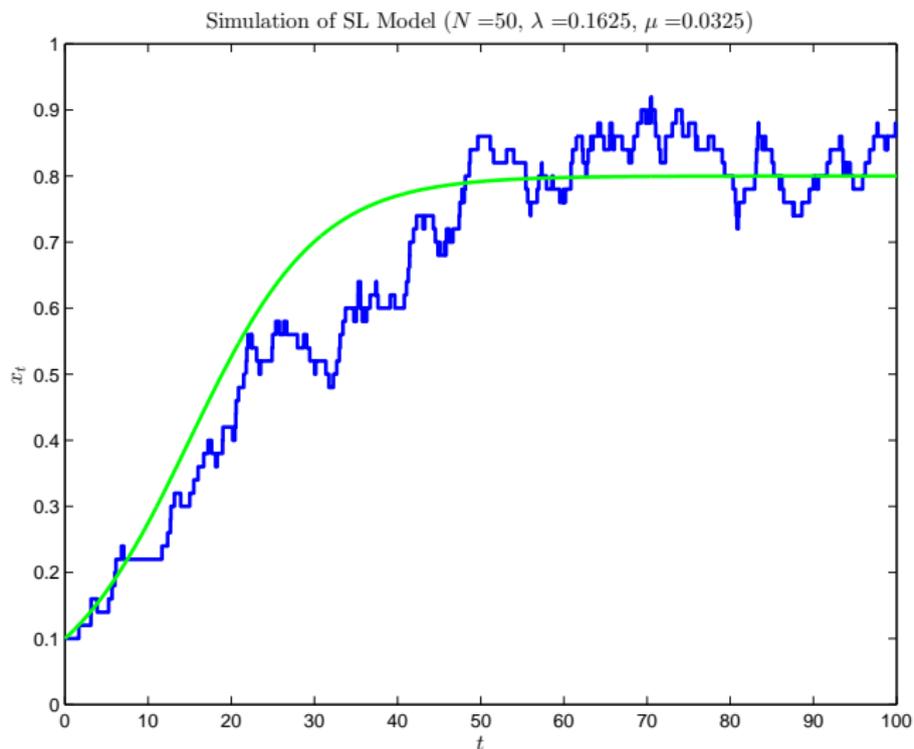
Simulation of the SL Model with x_t ($N = 50$)



Simulation of the SL Model with $x_t - N$ large



Simulation of the SL Model with x_t



Fluctuations about the deterministic trajectory

In a later paper Kurtz* proved a *functional central limit law* which establishes that, for large N , the fluctuations about the deterministic trajectory follow a *Gaussian diffusion*, provided that some mild extra conditions are satisfied.

He considered the family of processes $\{(Z_t^{(N)})\}$, indexed by N , and defined by

$$Z_s^{(N)} = \sqrt{N} (X_s^{(N)} - x_s), \quad s \in [0, t].$$

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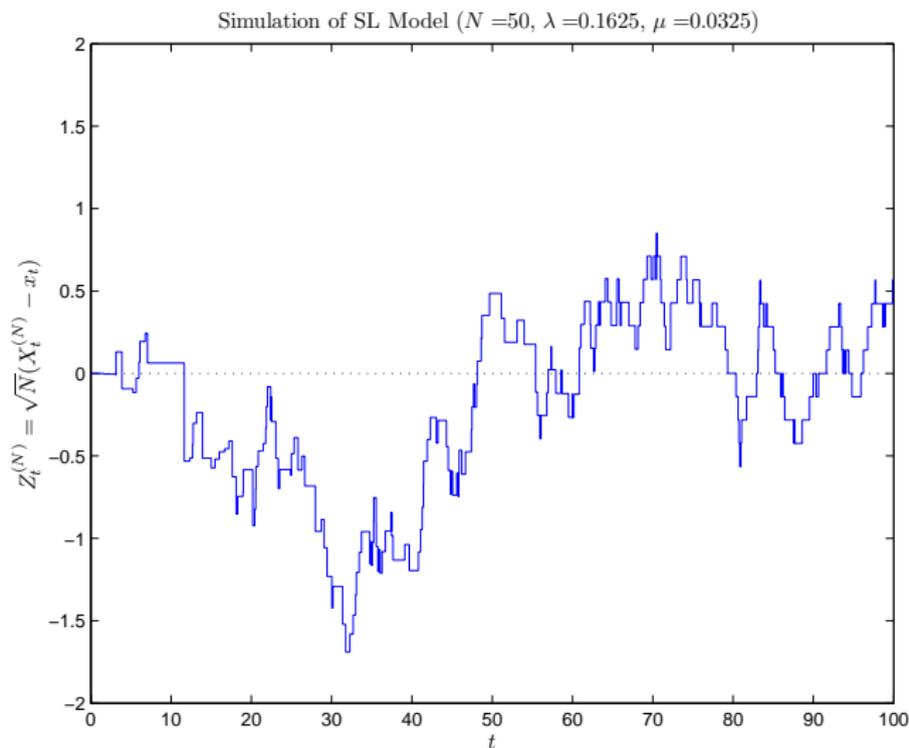
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Keep in mind the *Central Limit Theorem*. As applied to coin tossing (de Moivre ($\simeq 1733$)), if p_N is the proportion of “Heads” after N tosses of a fair coin,

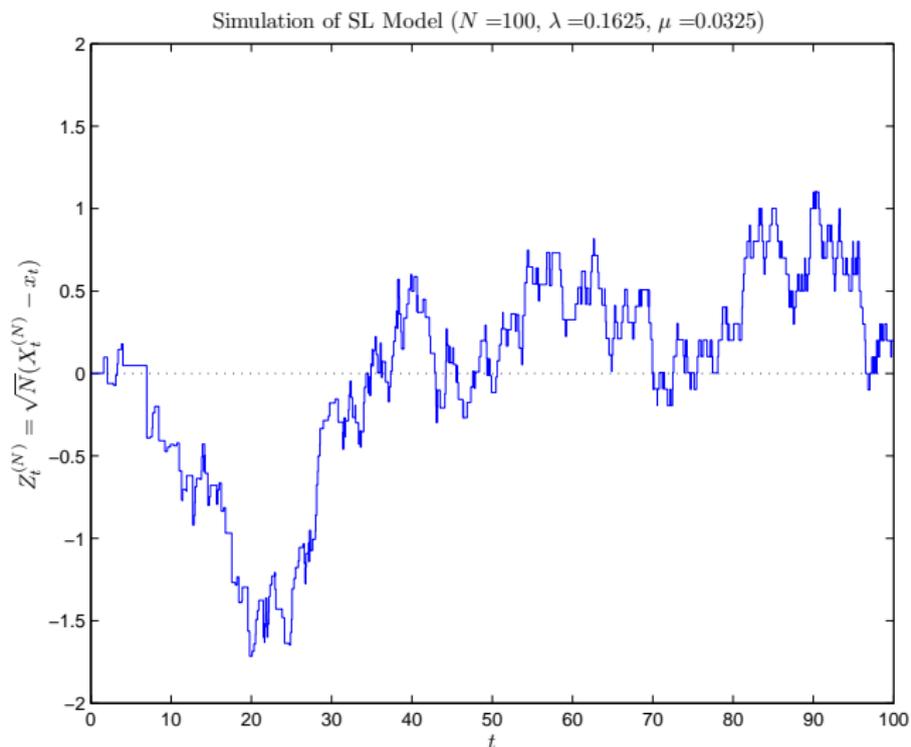
$$\sqrt{N} (p_N - \frac{1}{2}) \xrightarrow{D} Z \sim N(0, \frac{1}{4}), \quad \text{as } N \rightarrow \infty.$$

(STAT1201: the normal approximation to the binomial distribution.)

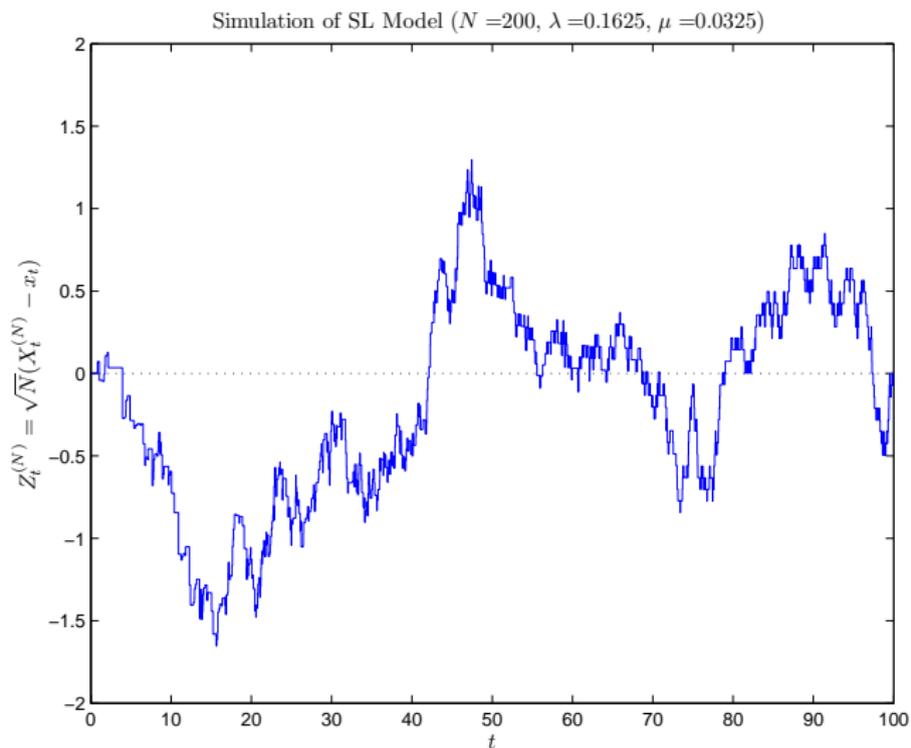
Scaled fluctuations in the SL Model ($N = 50$)



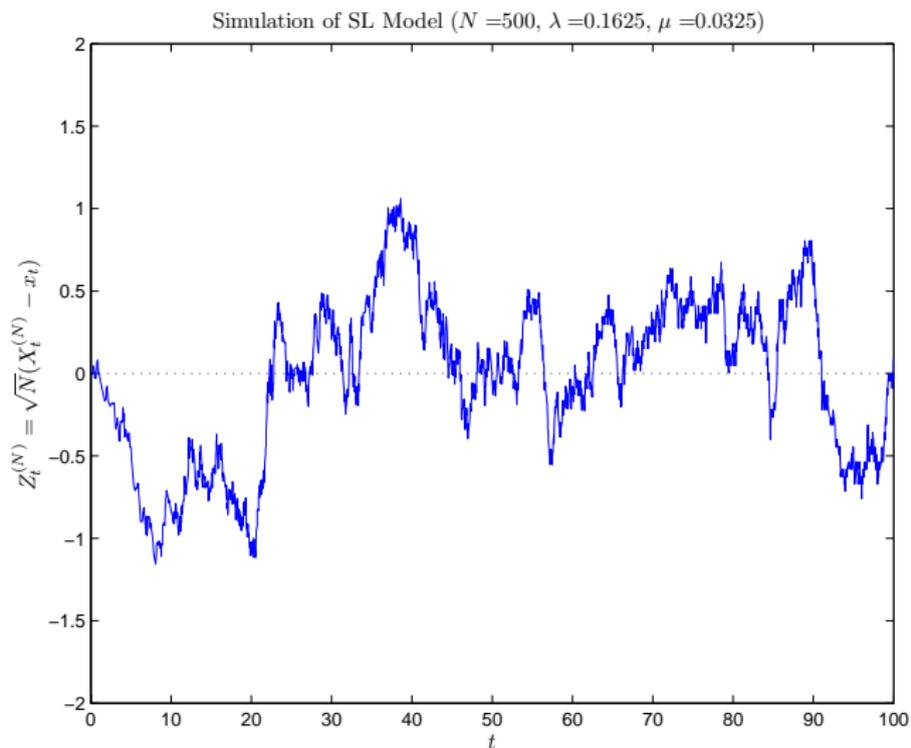
Scaled fluctuations in the SL Model ($N = 100$)



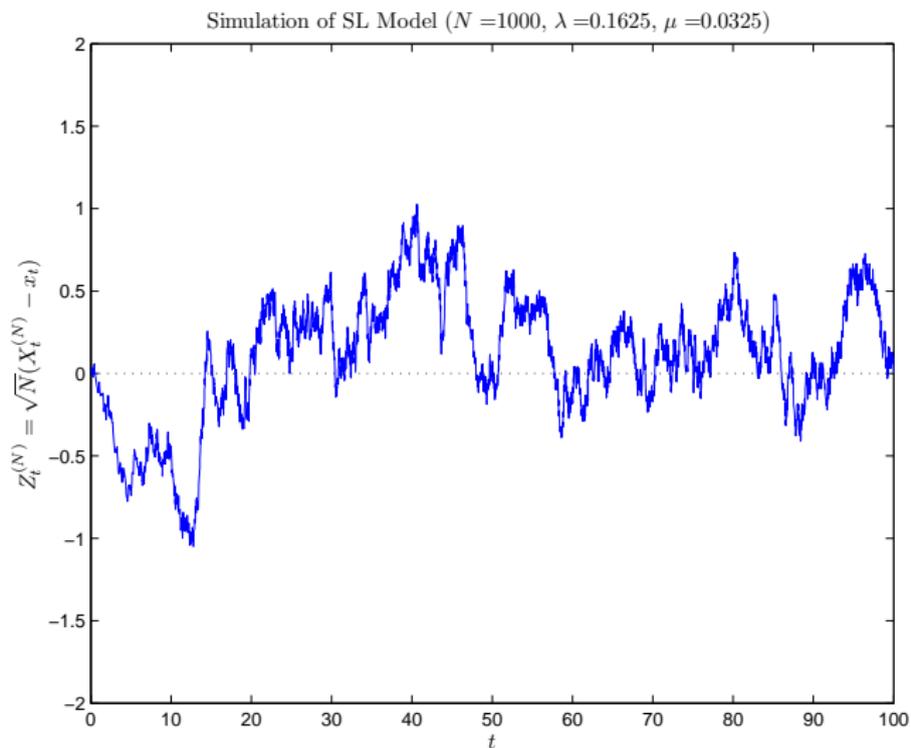
Scaled fluctuations in the SL Model ($N = 200$)



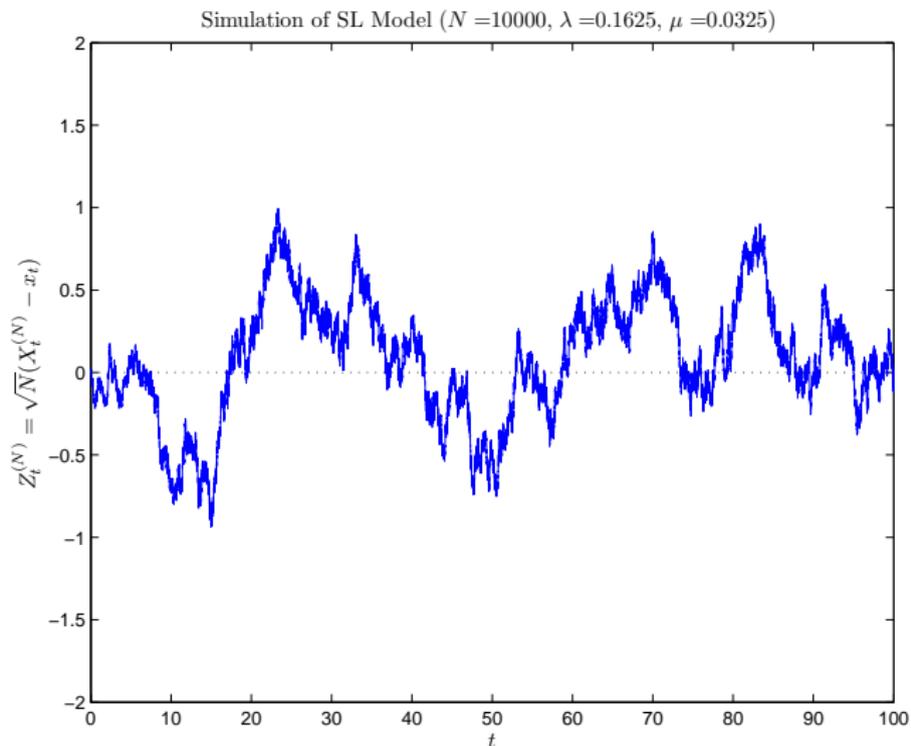
Scaled fluctuations in the SL Model ($N = 500$)



Scaled fluctuations in the SL Model ($N = 1000$)



Scaled fluctuations in the SL Model ($N = 10\,000$)



A central limit law

Theorem Suppose that F is Lipschitz continuous and has uniformly continuous first derivatives on E , and that the $k \times k$ matrix $G(x)$, defined for $x \in E$ by $G_{ij}(x) = \sum_{\ell \neq 0} \ell_i \ell_j f_\ell(x)$, is uniformly continuous on E .

Let (x_t) be the unique deterministic trajectory starting at x_0 and suppose that $\lim_{N \rightarrow \infty} \sqrt{N} (X_0^{(N)} - x_0) = z$.

Then, $\{(Z_t^{(N)})\}$ converges weakly in $D[0, t]$ (the space of right-continuous, left-hand limits functions on $[0, t]$) to a Gaussian diffusion (Z_t) with initial value $Z_0 = z$ and with mean given by $\mu_s := \mathbb{E}(Z_s) = M_s z$, where $M_s = \exp(\int_0^s B_u du)$ and $B_s = \nabla F(x_s)$, and covariance given by

$$\Sigma_s := \text{Cov}(Z_s) = M_s \left(\int_0^s M_u^{-1} G(x_u) (M_u^{-1})^T du \right) M_s^T.$$

A central limit law

The functional central limit theorem tells us that, for large N , the scaled density process $Z_t^{(N)}$ can be approximated *over finite time intervals* by the Gaussian diffusion (Z_t) .

In particular, for all $t > 0$, $X_t^{(N)}$ has an approximate normal distribution with $\text{Cov}(X_t^{(N)}) \simeq \Sigma_t/N$.

We usually take $X_0^{(N)} = x_0$, for all N , thus giving $\mathbb{E}(X_t^{(N)}) \simeq x_t$.

A central limit law for the SL Model

For the SL Model we have $F(x) = \lambda x(1 - \rho - x)$, and the solution to $dx/dt = F(x)$ is

$$x_t = \frac{(1-\rho)x_0}{x_0 + (1-\rho-x_0)e^{-\lambda(1-\rho)t}}.$$

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Then, $\mathbb{E}(X_t^{(N)}) \simeq x_t$. We also have $F'(x) = \lambda(1 - \rho - 2x)$ and

$$G(x) = \sum_{\ell} \ell^2 f_{\ell}(x) = \lambda x(1 + \rho - x) = F(x) + 2\mu x,$$

giving

$$M_t = \exp\left(\int_0^t F'(x_s) ds\right) = \frac{(1-\rho)^2 e^{-\lambda(1-\rho)t}}{(x_0 + (1-\rho-x_0)e^{-\lambda(1-\rho)t})^2}.$$

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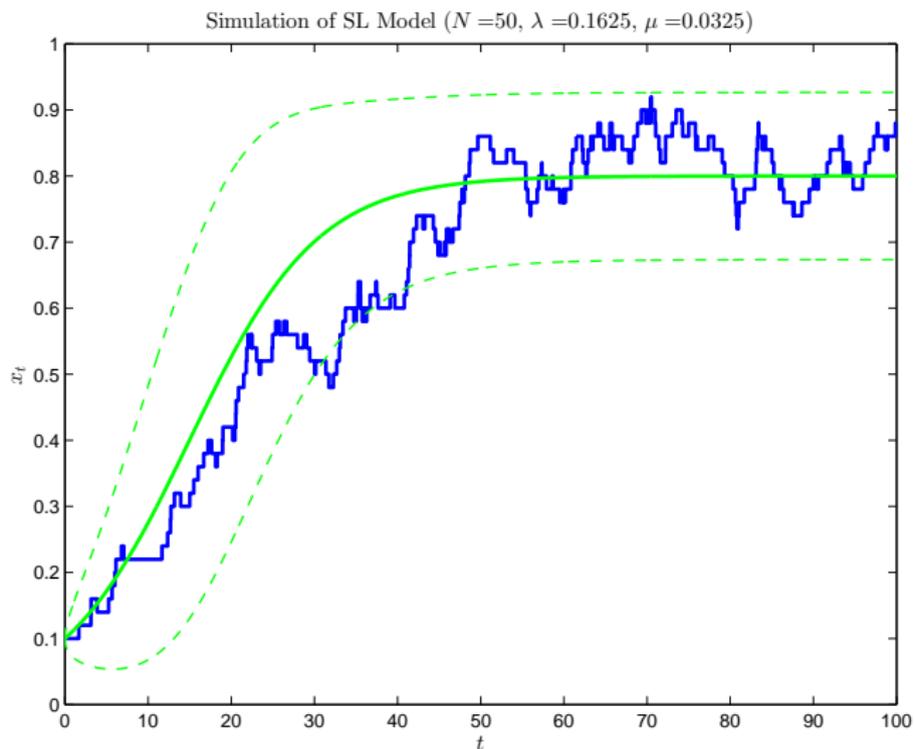
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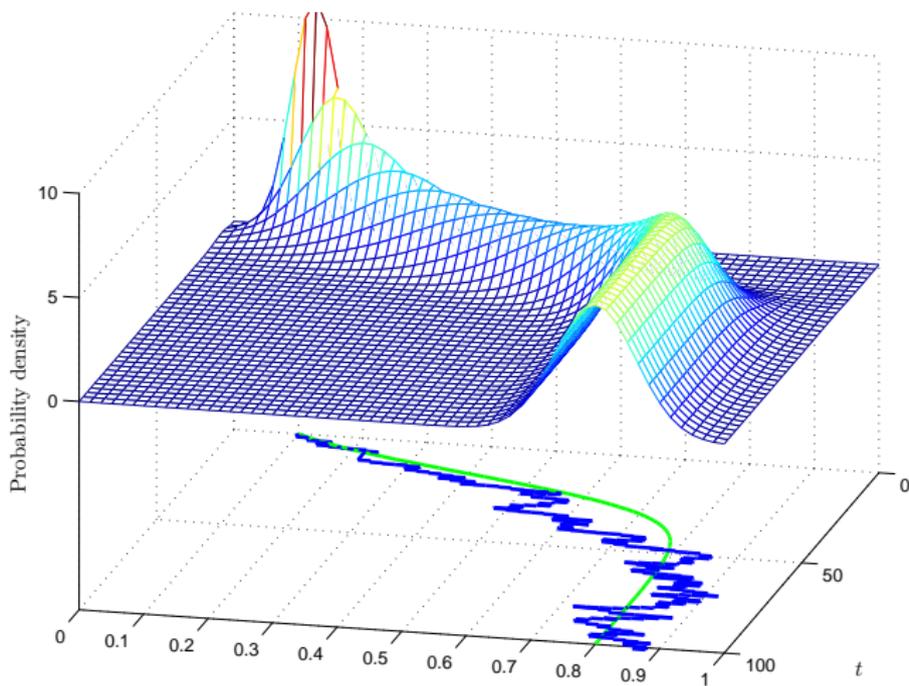
$$\begin{aligned} v_t = & x_0 \left(\rho x_0^3 + x_0^2(1+5\rho)(1-\rho-x_0)e^{-\lambda(1-\rho)t} + 2x_0(1+2\rho)(1-\rho-x_0)^2(\lambda(1-\rho)t)e^{-2\lambda(1-\rho)t} \right. \\ & - \left. ((1-\rho-x_0)[3\rho x_0^2 + (2+\rho)(1-\rho)x_0 - ((1+2\rho)(1-\rho)^2] + \rho(1-\rho)^3)e^{-2\lambda(1-\rho)t} \right. \\ & \left. - (1+\rho)(1-\rho-x_0)^3 e^{-3\lambda(1-\rho)t} \right) / \left(x_0 + (1-\rho-x_0)e^{-\lambda(1-\rho)t} \right)^4. \end{aligned}$$

Then, $\text{Var}(X_t^{(N)}) \simeq v_t/N$.

Simulation of the SL Model with $x_t \pm 2\sqrt{v_t/N}$



Simulation of the SL Model with Normal approximation



The OU approximation

If the initial point x_0 of the deterministic trajectory is chosen to be an equilibrium point of the deterministic model, we can be far more precise about the approximating diffusion.

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Corollary If x_{eq} satisfies $F(x_{\text{eq}}) = 0$, then, under the conditions of the theorem, the family $\{(Z_t^{(N)})\}$, defined by

$$Z_s^{(N)} = \sqrt{N}(X_s^{(N)} - x_{\text{eq}}), \quad 0 \leq s \leq t, \quad s \in [0, t].$$

converges weakly in $D[0, t]$ to an *Ornstein–Uhlenbeck (OU) process* (Z_t) with initial value $Z_0 = z$, local drift matrix $B = \nabla F(x_{\text{eq}})$ and local covariance matrix $G(x_{\text{eq}})$.



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In particular, Z_s is normally distributed with mean and covariance given by $\mu_s := \mathbb{E}(Z_s) = e^{Bs}z$ and

$$\Sigma_s := \text{Cov}(Z_s) = \int_0^s e^{Bu} G(x_{\text{eq}}) e^{B^T u} du.$$



The OU approximation

Note that

$$\Sigma_s = \int_0^s e^{Bu} G(x_{\text{eq}}) e^{B^T u} du = V_\infty - e^{Bs} V_\infty e^{B^T s},$$

where V_∞ , the stationary covariance matrix, satisfies

$$BV_\infty + V_\infty B^T + G(x_{\text{eq}}) = 0.$$

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For the SL Model, $v_t := \text{Var}(X_t^{(N)}) \simeq \rho(1 - e^{-2\lambda(1-\rho)t})/N$.

Finally, this brings us “full circle” to the approximating SDE

$$dn_t = -\alpha(n_t - K) dt + \sqrt{2N\alpha\rho} dB_t,$$

where $\alpha = \lambda(1 - \rho)$.

Simulation of the SL Model with $x_{\text{eq}} \pm 2\sqrt{v_t/N}$ (OU Approximation)

