

# Balancing Reduces Asymptotic Variance of Outputs

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# Outline

- 1 Asymptotic Variance Rate
- 2 M/M/1
- 3 M/M/1/K
- 4 GI/G/1
- 5 GI/G/1/K
- 6 Discussion

# Outline

1 Asymptotic Variance Rate

2 M/M/1

3 M/M/1/K

4 GI/G/1

5 GI/G/1/K

6 Discussion

# Asymptotic Variance Rate

- $D(t) \equiv$  Number of outputs during  $[0, t]$
- Often,  $\exists \bar{V} : \text{Var}(D(t)) = \bar{V}t + o(t)$
- $\bar{V} \equiv$  Asymptotic Variance Rate

## Examples

- $D(t)$  Poisson with rate  $\lambda$ :  $\bar{V} = \lambda$ 
  - E.g. stationary M/M/1 with arrival rate  $\lambda$
- $D(t)$  renewal with inter-renewal mean  $1/\lambda$  and SCV  $c^2$ :  $\bar{V} = \lambda c^2$ 
  - E.g. M/M/1/1 (no waiting room) with arrival = service rate =  $\lambda$ 
    - $D(t)$  renewal with Erlang(2,  $\lambda$ )
    - $\text{Var}(D(t)) = \frac{1}{4}\lambda t + \frac{1}{8} - \frac{1}{8}e^{-2\lambda t}$  (stationary case)
- M/G/1 with infinite variance Pareto service times:  
 $\bar{V}$  does not exist (Daley, Vesilo, 1997)

# Underloaded Queues ( $\lambda < \mu$ ): Typically $\bar{V} = \bar{V}_{in}$

## GI/G/1

Proof:

- 1 Assume  $Q(0) = 0$ :  $D(t) = A(t) - Q(t)$
- 2  $\text{Var}(D(t)) = \text{Var}(A(t)) + \text{Var}(Q(t)) - 2\text{Cov}(A(t), Q(t))$
- 3 Assume  $\text{Var}(Q(t)) < \infty$  and  $\bar{V}$  exists:  
 $\text{Cov}(A(t), Q(t)) \leq \sqrt{\text{Var}(A(t))\text{Var}(Q(t))} = O(\sqrt{t})$
- 4 Divide by  $t$  and take  $t \rightarrow \infty$

## GI/G/1/K

Intuition:

- 1 For  $K$  large and  $\lambda \ll \mu$ , queue is almost never full
- 2 Similar to GI/G/1 ( $\bar{V} \approx \bar{V}_{in}$ )

# Overloaded Queues ( $\lambda > \mu$ ): Typically $\bar{V} = \bar{V}_{\text{service}}$

## GI/G/1

Intuition:

- 1 After a finite time,  $\tau$ , queue never empties again
- 2 For  $t > \tau$ : output is a renewal process of services

## GI/G/1/K

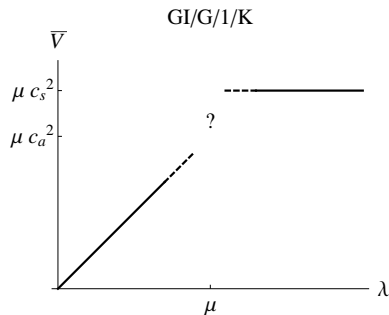
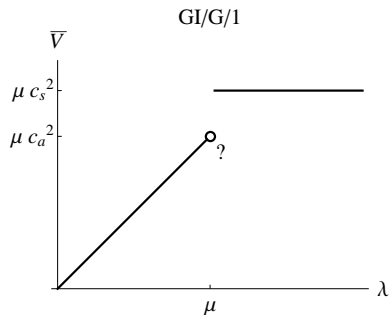
Intuition:

- 1 For  $K$  large and  $\lambda \gg \mu$ , queue is almost never empty
- 2 Output is "almost" renewal ( $\bar{V} \approx \bar{V}_{\text{service}}$ )

# Intermediate Summary

$$\bar{V}_{GI/G/1} = \begin{cases} \lambda c_a^2 & \lambda < \mu \\ ? & \lambda = \mu \\ \mu c_s^2 & \lambda > \mu \end{cases},$$

$$\bar{V}_{GI/G/1/K} = \begin{cases} \lambda c_a^2 & \lambda \ll \mu \\ ? & \lambda \approx \mu \\ \mu c_s^2 & \lambda \gg \mu \end{cases}$$



# Flow Rate and IDC

- Flow rate:

$$\lambda^* = \lim_{t \rightarrow \infty} \frac{E[D(t)]}{t}$$

- Index of dispersion of counts (IDC):

$$IDC = \lim_{t \rightarrow \infty} \frac{\text{Var}(D(t))}{E[D(t)]} = \frac{\bar{V}}{\lambda^*}$$

## Examples

- IDC of a Poisson process is 1
- IDC of a renewal process is  $c^2$

For GI/G/1 and GI/G/1/K with  $\lambda \neq \mu$ :

The IDC is "renewal like"



# Preview of Results

We shall show:

**BRAVO: Balancing Reduces Asymptotic Variance of Outputs**

M/M/1	GI/G/1	M/M/1/K	GI/G/1/K
$\lambda 2 \left(1 - \frac{2}{\pi}\right)$	$\lambda (c_a^2 + c_s^2) \left(1 - \frac{2}{\pi}\right)$	$\lambda \frac{2K^2 + K}{3(K+1)^2}$	$\lambda \left( (c_a^2 + c_s^2) \frac{1}{3} + o_K(1) \right)$

## Notes

- When  $\lambda = \mu$  the IDC is NOT "renewal like"
- $\lim_{K \rightarrow \infty} \bar{V}_{GI/G/1/K} \neq \bar{V}_{GI/G/1}$
- Currently: Complete proofs for M/M cases only
- Currently: Lacking good intuition for BRAVO

# Outline

1 Asymptotic Variance Rate

2 **M/M/1**

3 M/M/1/K

4 GI/G/1

5 GI/G/1/K

6 Discussion

## Theorem

For the M/M/1 queue with  $\lambda = \mu$ :

$$\bar{V} = \lambda^2 \left(1 - \frac{2}{\pi}\right)$$

## Notes

- $2\left(1 - \frac{2}{\pi}\right) = 0.72676\dots$
- Explicit expression for  $\text{Var}(D(t))$  in terms of Bessel functions
- From this explicit expression:

$$\text{Var}(D(t)) = \bar{V}t + \frac{1}{\sqrt{\pi}} (\lambda t)^{1/2} + \left(\frac{1}{4} - \frac{1}{2\pi}\right) + o_t(1)$$

# M/M/1 BRAVO: Outline of Derivation

- Denote:  $X_\alpha \sim \exp(\alpha)$ ,  $\phi_\alpha(z) := E[z^{D(X_\alpha)}] = \int_0^\infty \alpha e^{-\alpha t} E[z^{D(t)}] dt$
- Denote  $\phi_\alpha^1 := \frac{\phi'_\alpha(1)}{\alpha}$  and  $\phi_\alpha^2 := \frac{\phi''_\alpha(1)}{\alpha}$ :

$$E[D(t)] = \mathcal{L}^{-1}(\phi_\alpha^1), \quad E[D(t)^2] = \mathcal{L}^{-1}(\phi_\alpha^2 + \phi_\alpha^1)$$

- Known:

—  $\phi_\alpha(z) = \frac{\alpha}{\mu(1-z)+\alpha} \left( 1 + r_\alpha(z) \frac{1-z}{z-r_\alpha(z)} \right)$ ,  $r_\alpha(z) = \frac{\lambda+\mu+\alpha - \sqrt{(\lambda+\mu+\alpha)^2 - 4\lambda\mu z}}{2\lambda}$

— Inverse transforms of  $\phi_\alpha^1$  and  $\phi_\alpha^2$

- Set  $\lambda = \mu = 1$  and combine:

$$\begin{aligned} \text{Var}((D(t))) &= \frac{1}{4} e^{-4t} \left( e^{4t} (1 + 8t) - (1 + 4t)^2 I_0(2t)^2 - 4e^{2t} t I_1(2t) \right. \\ &\quad \left. - 16t^2 I_1(2t)^2 - 4t I_0(2t) \left( e^{2t} + (2 + 8t) I_1(2t) \right) \right), \end{aligned}$$

with  $I_j(2t) = \sum_{n=0}^{\infty} \frac{t^{j+2n}}{(j+n)! \cdot n!}$

# M/M/1 BRAVO: Relation to the Y-Intercept

## Theorem

For M/G/1 with  $\lambda < \mu$ , starting empty and third service moment finite:

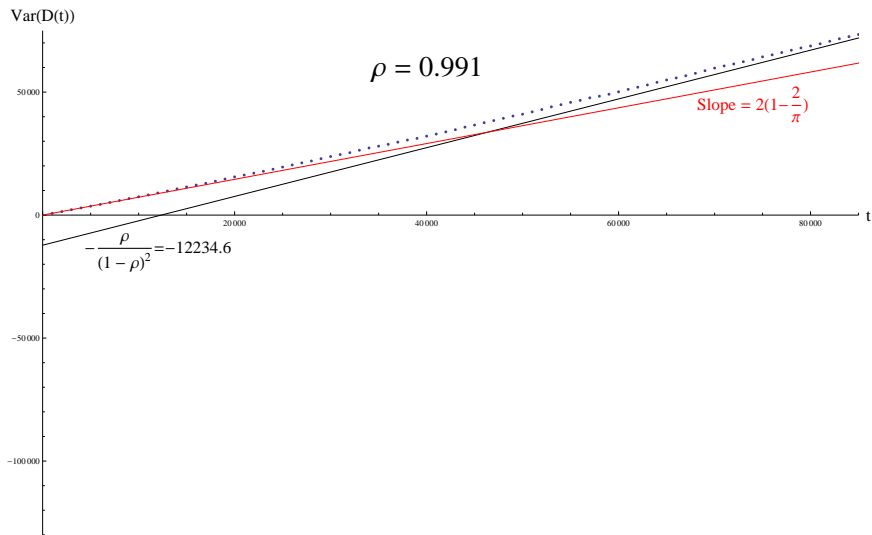
$$\text{Var}(D(t)) = \lambda t - (1 - L_0) \frac{\rho}{(1 - \rho)^2} + o_t(1)$$

- $L_0$  expression of  $\rho$  and first 3 service moments
- For M/M/1,  $L_0 = 0$

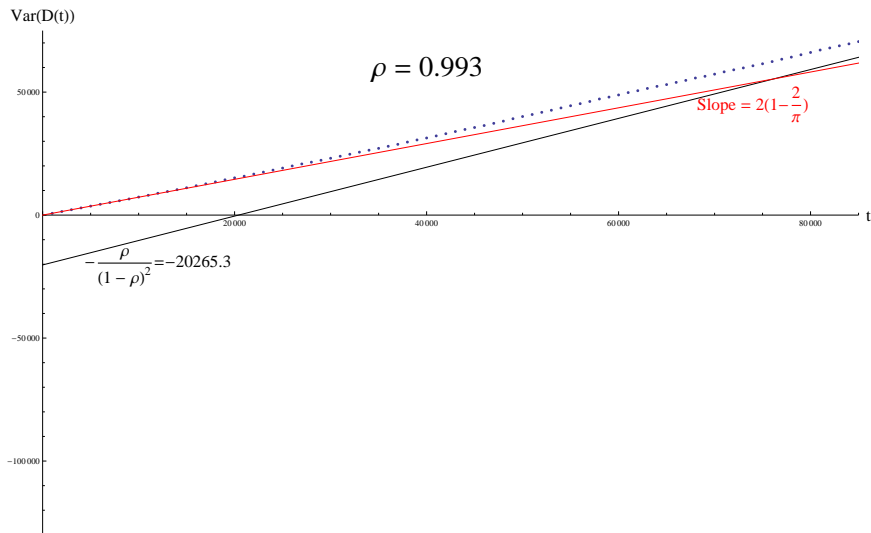
## Example: Look at the curves $\text{Var}(D(t))$

- M/M/1 starting empty,  $\mu = 1$ ,  $\rho = 0.991, 0.993, 0.995, 0.997$
- Linear asymptote slope  $\approx 1$ ,  
y-intercepts  $\approx -10^4, -2 \times 10^4, -4 \times 10^4, -10^5$
- Simulate  $D(t)$ :  
 $3 \times 10^4$  repetitions,  $10^5$  time units, sample variance every 1000

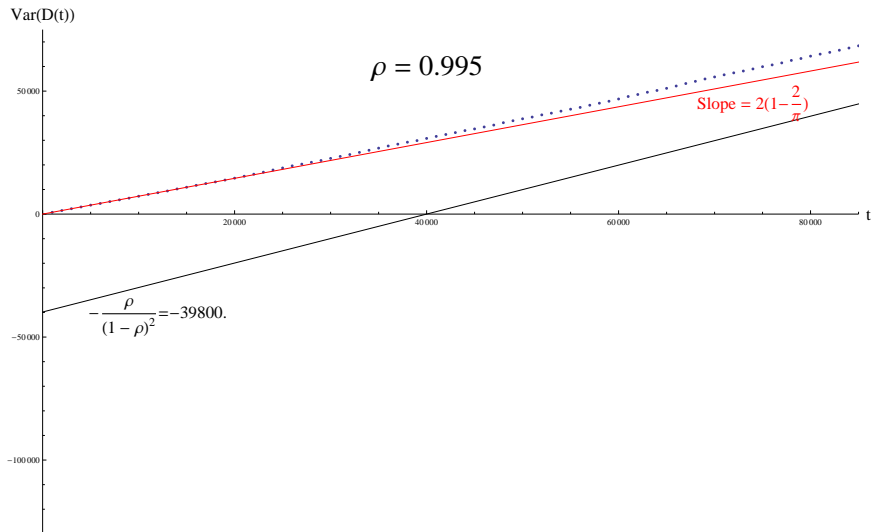
# M/M/1 BRAVO: Relation to the Y-Intercept



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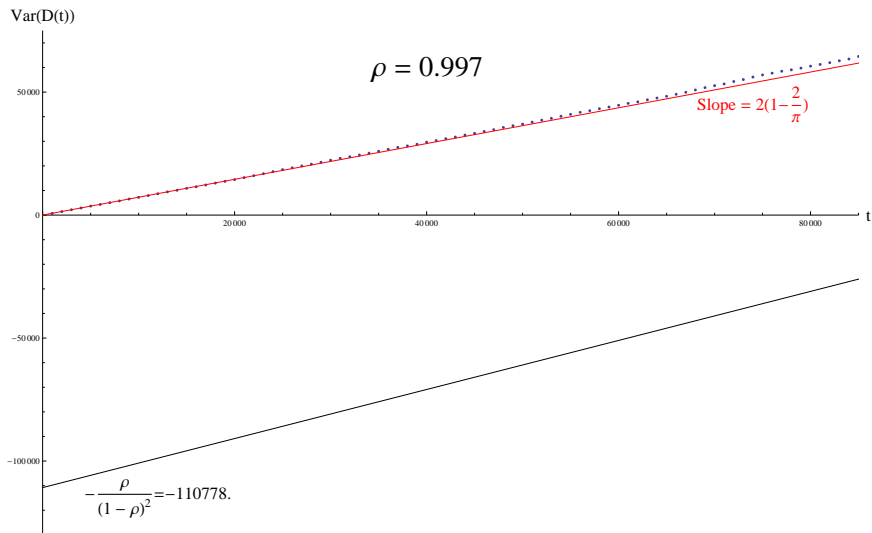


# M/M/1 BRAVO: Relation to the Y-Intercept





# M/M/1 BRAVO: Relation to the Y-Intercept



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2 M/M/1

**3 M/M/1/K**

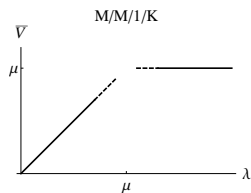
4 GI/G/1

5 GI/G/1/K

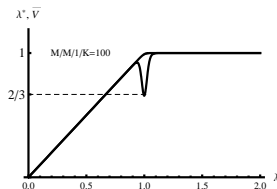
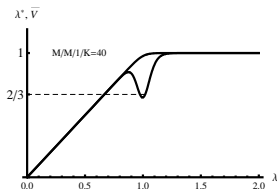
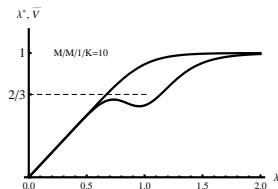
6 Discussion

# M/M/1/K BRAVO

We expect similar to M/M/1 when  $\lambda \ll \mu$  or  $\lambda \gg \mu$ :



For  $\lambda \approx \mu$  we get,  $\bar{V} \approx \lambda \frac{2}{3}$ :



## Theorem

For the M/M/1/K queue with  $\lambda = \mu$ :

$$\bar{V} = \lambda \frac{2K^2 + K}{3(K+1)^2}$$

## Notes

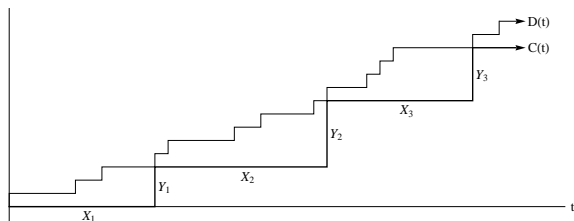
- Simple formula for any finite birth-death queue (depends on  $\pi$ )
- For "monotonic" finite birth-death queues,  $IDC < 1$
- BRAVO also appears in M/M/c/K queues (IDC approximately minimized when  $\lambda = \mu c$ )
- For stationary M/M/1/K,  $\lambda = \mu$ :

$$\text{Var}(D(t)) = \bar{V} t + \frac{1}{180} \frac{7K^4 + 28K^3 + 37K^2 + 18K}{(K+1)^2} + o_t(1)$$

# M/M/1/K BRAVO Derivation (New Simple)

## "Embed" $D(t)$ in a Renewal-Reward Process, $C(t)$

- 1  $(X_n, Y_n) \equiv$  (busy cycle, number served) in cycle  $n$
- 2  $N(t) = \inf\{n : \sum_{i=1}^n X_i > t\}$ ,  $C(t) = \sum_{i=1}^{N(t)} Y_i$
- 3 Variance rates of  $C(t)$  and  $D(t)$  are equal
- 4 Known:
  - Variance rate of  $C(t)$  is  $\frac{1}{E[X]} \text{Var}(Y - \frac{E[Y]}{E[X]} X)$
  - For M/M/1/K: Explicit joint transform of  $(X, Y)$



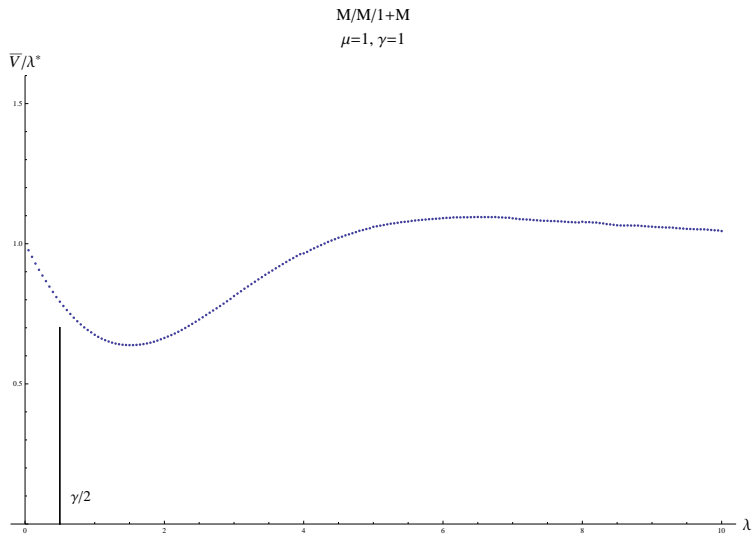
# Estimate $\bar{V}$ by Regenerative Simulation

- Assume  $\text{Var}(D(t)) = \bar{V}t + \bar{B} + o_t(1)$ . Bias of  $\frac{\text{Var}(D(T))}{T}$  is  $\bar{B}/T$
- Alternative: estimate moments of  $(X, Y)$  and plug in  $\frac{1}{E[X]} \text{Var}(Y - \frac{E[Y]}{E[X]}X)$

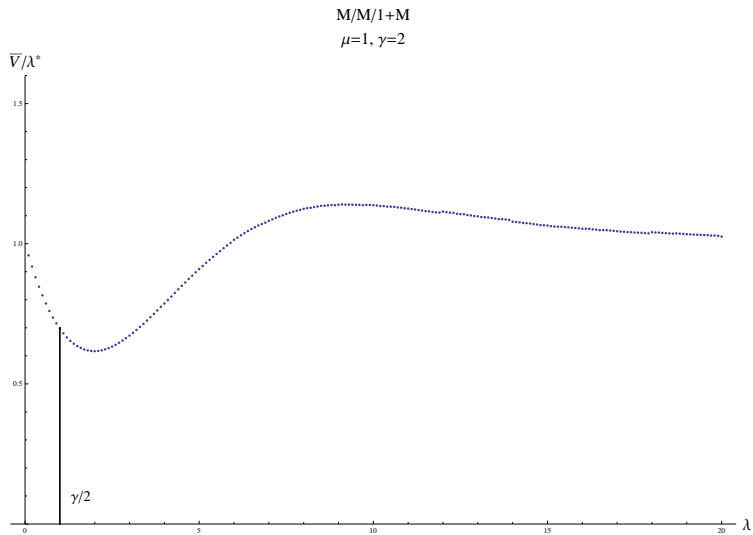
## Example: M/M/1+M

- Arrivals:  $\lambda$ , service:  $\mu$ , patience:  $\gamma$
- $D(t)$  is a "thinned" version of  $A(t)$ 
  - For i.i.d. thinning, IDC is 1
  - What about this case...
- Simulation details:
  - $\mu = 1, \gamma = 1, 2, 4, 8, \dots, 256$
  - $\lambda$  varies over 200 points in  $(0, \gamma/20)$
  - Regeneration state: approximately maximize  $\pi_i$
  - $2 \times 10^6$  seeds, keep seed constant while varying  $\lambda$
  - Running time: 15 minutes

# Regenerative Simulation for M/M/1 + M

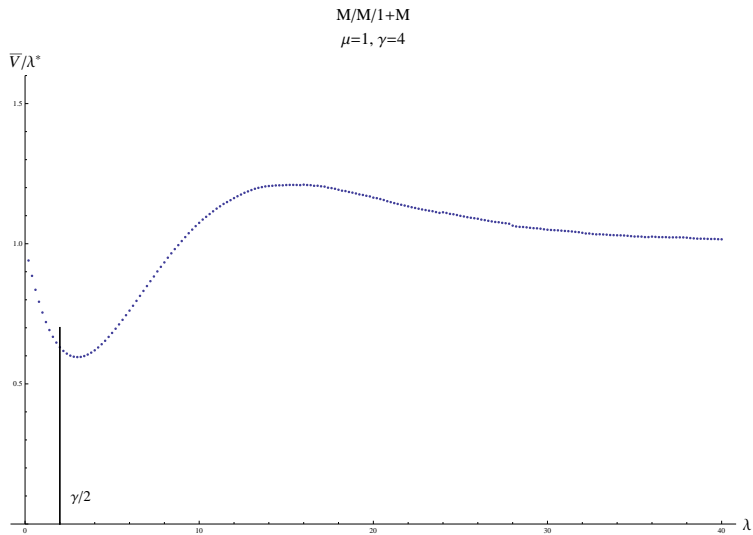


# Abandonments: M/M/1 + M

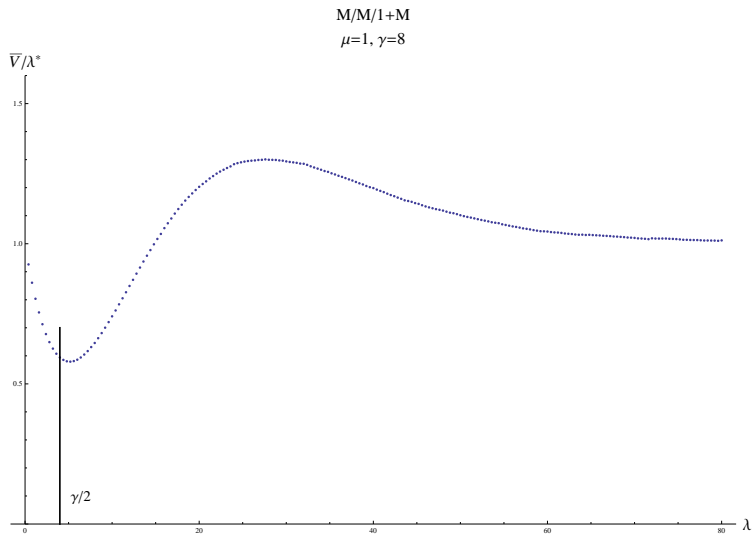




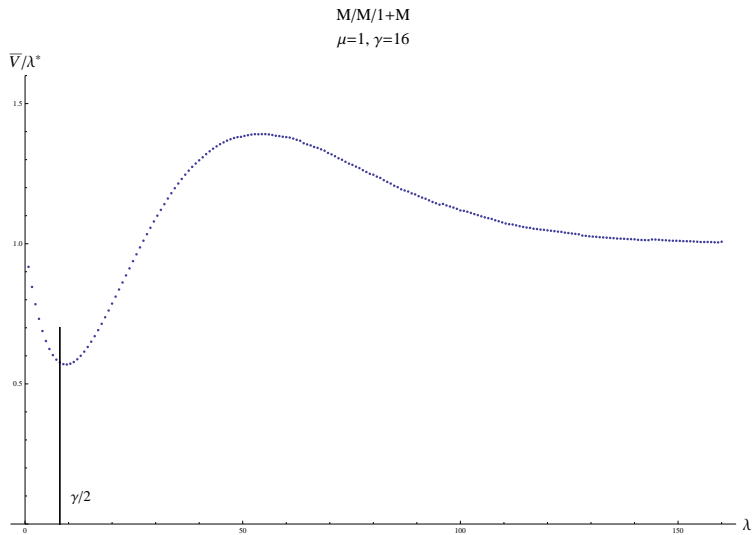
# Abandonments: M/M/1 + M



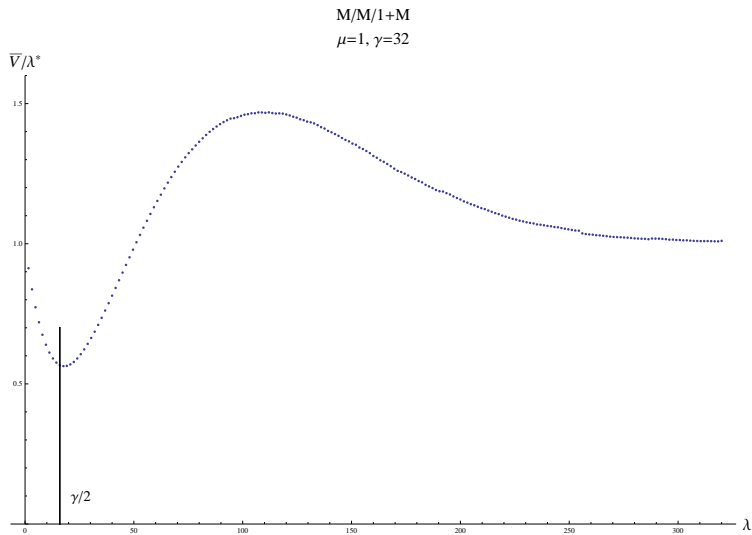
# Abandonments: M/M/1 + M



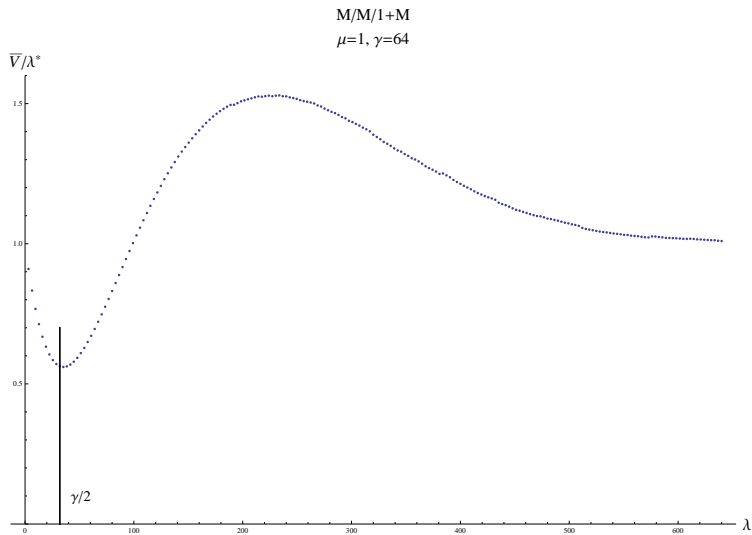
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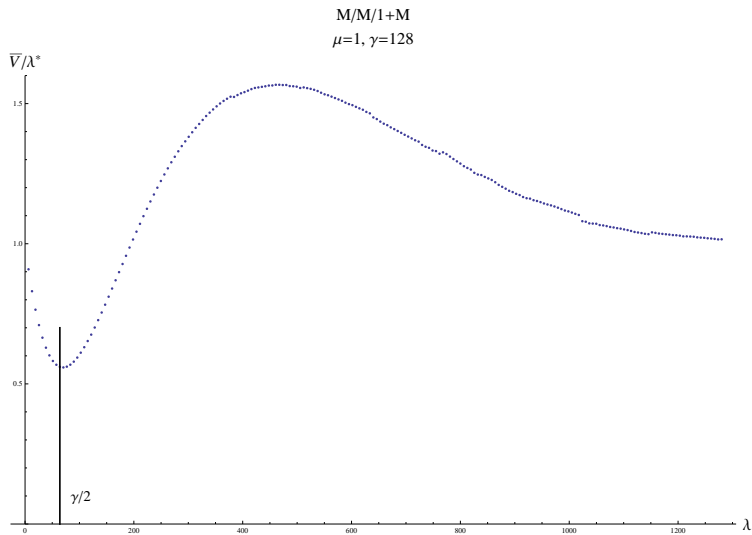
# Abandonments: M/M/1 + M



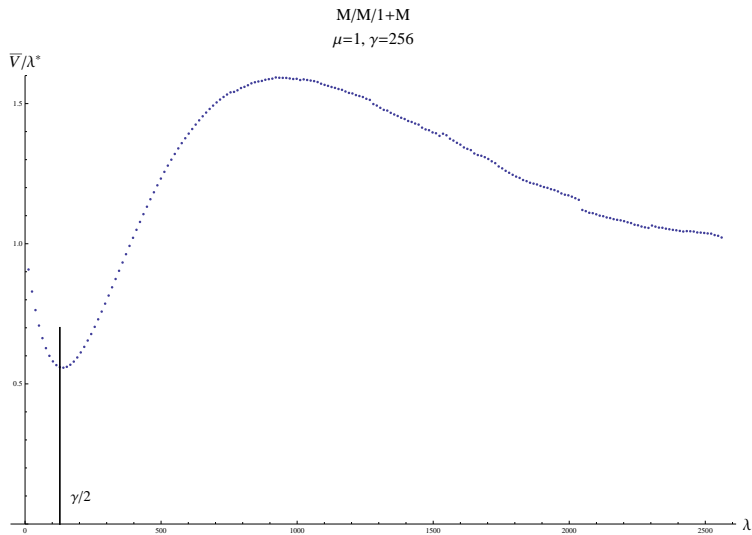
# Abandonments: M/M/1 + M



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# Abandonments: M/M/1 + M



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1 Asymptotic Variance Rate

2 M/M/1

3 M/M/1/K

**4 GI/G/1**

5 GI/G/1/K

6 Discussion



## Claim

For GI/G/1 with  $\lambda = \mu$ :

$$\bar{V} = \lambda(c_a^2 + c_s^2)\left(1 - \frac{2}{\pi}\right)$$

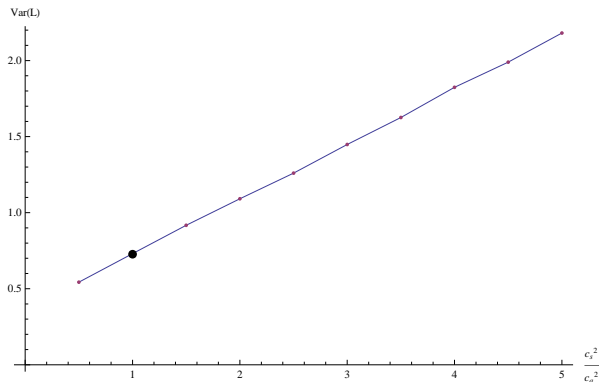
## Notes

- No full proof (yet)
- Moment conditions
- Holds also for  $GI/G/c$

## Outline of basis for claim

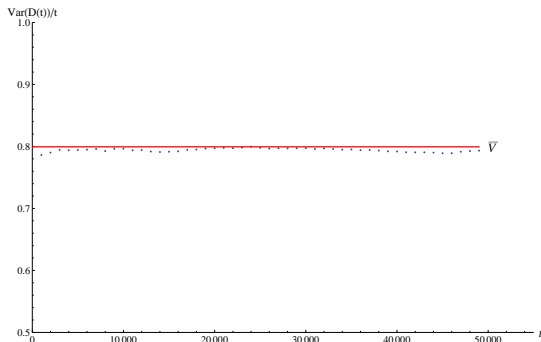
- 1 Iglehart and Whitt (1970):
  - GI/G/c,  $\lambda = \mu$
  - $\hat{D}^n(t) := (D(nt) - \lambda nt)n^{-1/2}$
  - $B_1(\cdot), B_2(\cdot)$  independent Brownian motions
  - Continuous mapping:  $g(U, V)(t) := V(t) + \inf_{0 \leq s \leq t} \{U(s) - V(s)\}$
  - $\hat{D}^n(\cdot) \Rightarrow g(c_a B_1(\cdot), c_s B_2(\cdot))$
- 2  $\frac{D(t) - \lambda t}{\sqrt{\lambda t}} \Rightarrow c_a \inf_{0 \leq t \leq 1} \{B_1(t) + \sqrt{\frac{c_s^2}{c_a^2}} B_2(1 - t)\}$ 
  - Uniform integrability conditions for first two moments
  - $\bar{V} = \lim_{t \rightarrow \infty} \frac{\text{Var}(D(t))}{t} = \lambda c_a^2 \text{Var}(\inf_{0 \leq t \leq 1} \{B_1(t) + \sqrt{\frac{c_s^2}{c_a^2}} B_2(1 - t)\})$
- 3 Brownian motion simulations show variance is affine in  $\frac{c_s^2}{c_a^2}$
- 4 M/D/1 result obtained directly:  $\bar{V} = \lambda \left(1 - \frac{2}{\pi}\right)$
- 5 M/D/1, M/M/1 and affinity in  $\frac{c_s^2}{c_a^2}$  imply the result

- Brownian motion simulations of  $\inf_{0 \leq t \leq 1} \{B_1(t) + \sqrt{\frac{c_s^2}{c_a^2}} B_2(1-t)\}$
- Variance estimates as a function of  $\frac{c_s^2}{c_a^2}$



## Explicit Simulation of $D(t)$ , Taking Sample Variance

- Poisson arrivals ( $c_a^2 = 1$ ), hyper-exp service with  $c_s^2 = 1.2$
- $\bar{V} = (c_a^2 + c_s^2) \left(1 - \frac{2}{\pi}\right) = 0.80$
- $10^5$  repetitions, 50,000 time units, sample variance every 1000

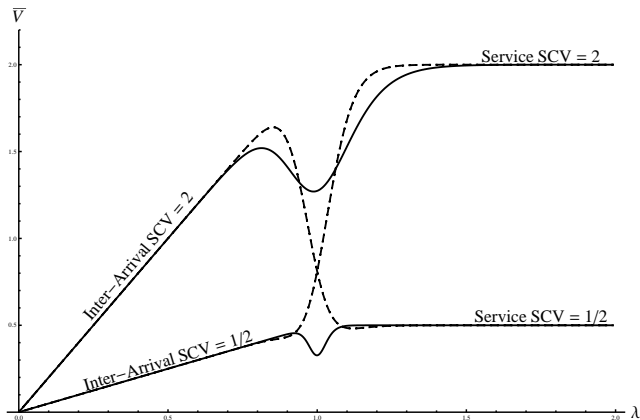


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# GI/G/1/K BRAVO

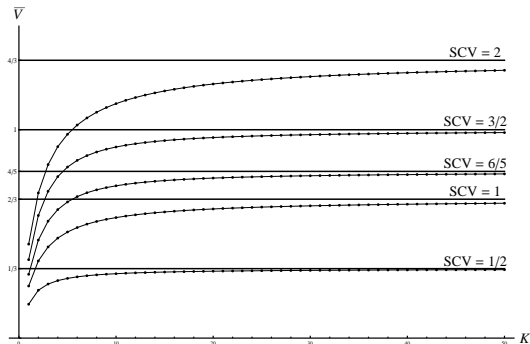
Evaluating  $\bar{V}$  using Markov Arrival Processes (Outputs of PH/PH/1/40):



Clearly there is a BRAVO effect. What is the magnitude?

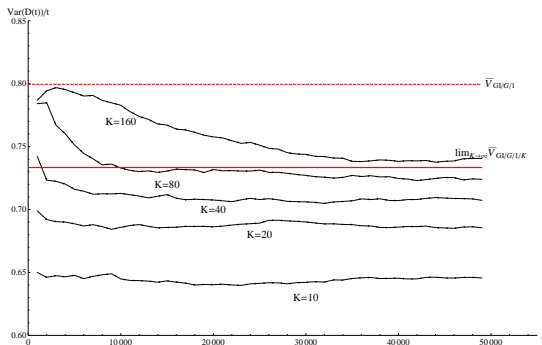
We Think:

$$\bar{V} = \lambda(c_a^2 + c_s^2) \frac{1}{3} + o_K(1)$$

Numerically verifying for  $c_a = c_s$ :Note: we also verified for  $c_a \neq c_s$

Explicit Simulation of  $D(t)$ , Taking Sample Variance

- Poisson arrivals ( $c_a^2 = 1$ ), hyper-exp service with  $c_s^2 = 1.2$
- $\bar{V} \approx (c_a^2 + c_s^2) \frac{1}{3} = 0.73$
- $10^5$  repetitions, 50,000 time units, sample variance every 1000





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# Discussion

## Question

What "causes" BRAVO?

## Answer

...

## Question

Applications to operations management?

## Answer

...

Thank you  
for your attention