BRAVO for QED Queues

(and more stuff about variance of output counts)

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Minimal Background: Queueing Models in this Talk

- Birth-Death queues:
  M/M/1, M/M/1/K, M/M/s/K, M/M/s/K+M ...
- General service times: G/G/1, G/G/1/K ...
- Open stable Jackson networks
- Open stable generalized Jackson networks

A basic packet conservation equation

$$Q(t) = Q(0) + \Big(A(t) - L(t)\Big) - \Big(R(t) + D(t)\Big)$$

Stochastic Counting Processes and their Variance

- Poisson processes:

$$\mathbb{E}[N(t)] = \mathrm{Var}\big(N(t)\big) = \lambda t$$

- Renewal processes:

$$\mathbb{E}[N(t)] \sim \lambda t \qquad \mathrm{Var}\big(N(t)\big) \sim \lambda c^2 t$$

- Counting processes resulting from queues?

$$Q(t) = Q(0) + \Big(A(t) - L(t)\Big) - \Big(R(t) + D(t)\Big)$$

The Output Process $D(\cdot)$

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Why analysis of the output counting process $\{D(t), t \geq 0\}$?

- Orders
- Production
- Arrival process to a downstream queueing system

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Daryl Daley, "*Queueing Output Processes*", Advances in Applied Probability, 1976.

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- Orders
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Daryl Daley, "*Queueing Output Processes*", Advances in Applied Probability, 1976.

Some performance measures of interest

- The law of $\{D(t),\ t \geq 0\}$
- $\mathbb{E}[D(t)]$, $\mathrm{Var}\big(D(t)\big)$
- $\lambda^* := \lim_{t\to\infty} \frac{\mathbb{E}[D(t)]}{t}$, $\quad \overline{V} := \lim_{t\to\infty} \frac{\mathrm{Var}\big(D(t)\big)}{t}$, $\quad \mathcal{D} := \frac{\overline{V}}{\lambda^*}$
- Asymptotic normality: $D(t) \sim \mathcal{N}\Big(\lambda^* t,\ \overline{V} t\Big)$, large t

Talk Outline

- A (new) formula for asymptotic variance of outputs, $\mathcal{D} := \frac{\overline{V}}{\lambda^*}$
- Single servers (older BRAVO results)
- Many server scaling (new BRAVO results)
- Something else: Asymptotic variance in queueing networks

Asymptotic Variance of Outputs

Finite Birth-Death Asymptotic Variance

- Irreducible birth-death process on finite state space
- Birth rates: $\lambda_0, \ldots, \lambda_{J-1}$
- Death rates: $\mu_1, \ldots, \mu_J$
- Stationary distribution: $\pi_0, \ldots, \pi_J$
- $D(t)$ is number of downward transitions (deaths) during $[0, t]$, each "filtered" independently with state-dependent probabilities, $q_1, \ldots, q_J$.
- e.g. The departure process (processed packets) in M/M/s/K+M systems

Of interest:

$$\mathcal{D} = \frac{\overline{V}}{\lambda^*} = \lim_{t\to\infty} \frac{\mathrm{Var}\big(D(t)\big)}{\mathbb{E}[D(t)]}$$

Finite Birth-Death Asymptotic Variance Formula

Theorem: Daryl Daley, Johan van Leeuwaarden, Y.N. 2013

$$\mathcal{D} := \lim_{t\to\infty} \frac{\mathrm{Var}\big(D(t)\big)}{\mathbb{E}[D(t)]} = 1 - 2\sum_{i=0}^{J}(P_i - \Lambda_i^*)\Big(q_{i+1} - \frac{\lambda^*}{\pi_i\lambda_i}(P_i - \Lambda_i^*)\Big),$$

with,

$$P_i := \sum_{j=0}^{i}\pi_j, \qquad \lambda^* := \sum_{j=1}^{J}\mu_j q_j \pi_j, \qquad \Lambda_i^* := \frac{\sum_{j=1}^{i}\mu_j q_j \pi_j}{\lambda^*}.$$

Note: In Weiss, Y.N. 2008, similar expression for case $q_i \equiv 1$

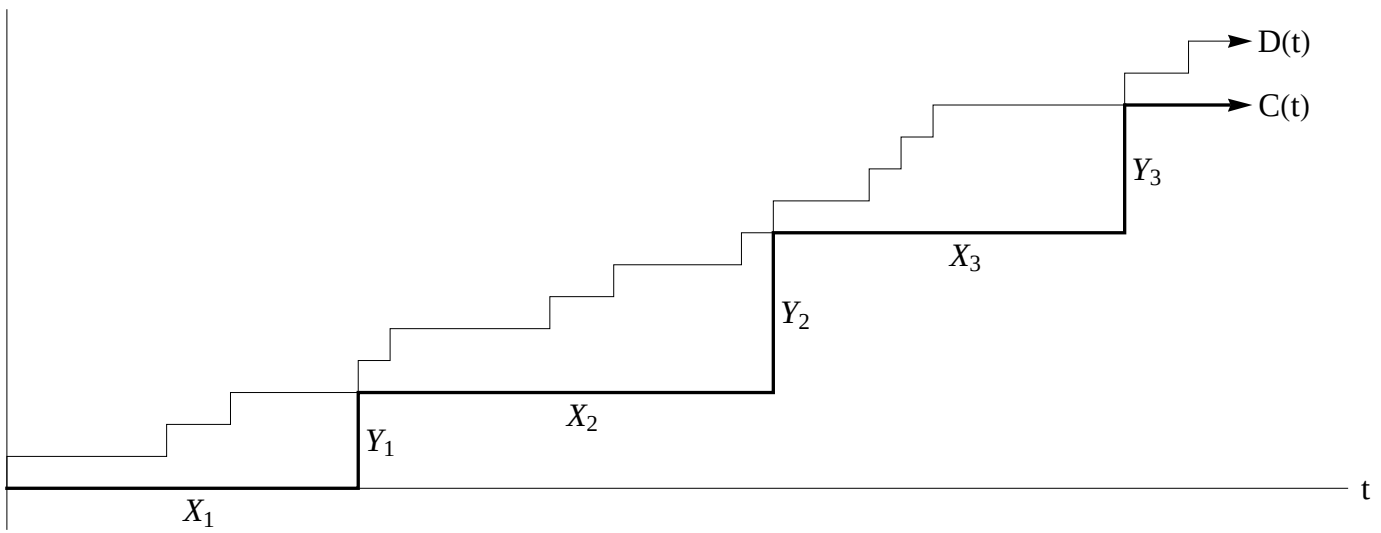
Note: In case $\lambda_i \equiv \lambda$, $q_i \equiv 1$:

$$\mathcal{D} = 1 - 2\frac{\pi_J}{1-\pi_J}\sum_{i=0}^{J} P_i\Big(1 - \pi_J\frac{P_i}{\pi_i}\Big)$$

Idea of Renewal Reward Derivation

"Embed" $D(t)$ in a Renewal-Reward Process, $C(t)$

1. $(X_n, Y_n) \equiv$ (busy cycle, number served) in cycle n
2. $N(t) = \sup\{n : \sum_{i=1}^{n} X_i \leq t\}$, $C(t) = \sum_{i=1}^{N(t)} Y_i$
3. Asymptotic variance rates of $C(t)$ and $D(t)$ are equal
4. Known:
   – Asymptotic variance rate of $C(t)$ is $\frac{1}{\mathbb{E}[X]}\text{Var}\left(Y - \frac{\mathbb{E}[Y]}{\mathbb{E}[X]}X\right)$
   – Systems of equations for 1'st, 2'nd and cross moments of X and Y



Single Server BRAVO
(older results)

M/M/1/K Queue

Here π_i is truncated geometric distribution when $\lambda \neq \mu$ and a uniform distribution when $\lambda = \mu$

Using $\mathcal{D} = 1 - 2\frac{\pi_J}{1-\pi_J}\sum_{i=0}^{J} P_i\left(1 - \pi_J\frac{P_i}{\pi_i}\right)$:

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Using $\mathcal{D} = 1 - 2\frac{\pi_J}{1-\pi_J}\sum_{i=0}^{J} P_i\left(1 - \pi_J\frac{P_i}{\pi_i}\right)$:

$$\mathcal{D} = \begin{cases} 1 + o_K(1), & \lambda \neq \mu, \\ \frac{2}{3} + o_K(1), & \lambda = \mu. \end{cases}$$

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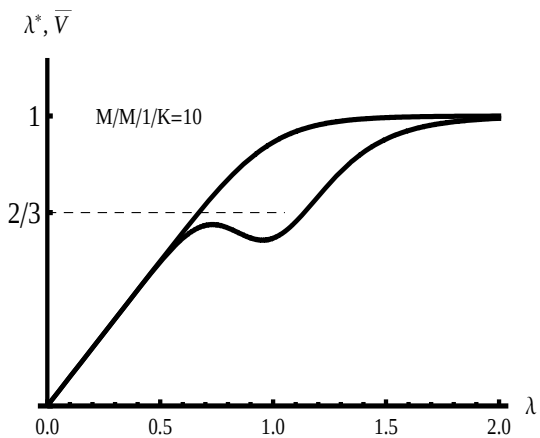

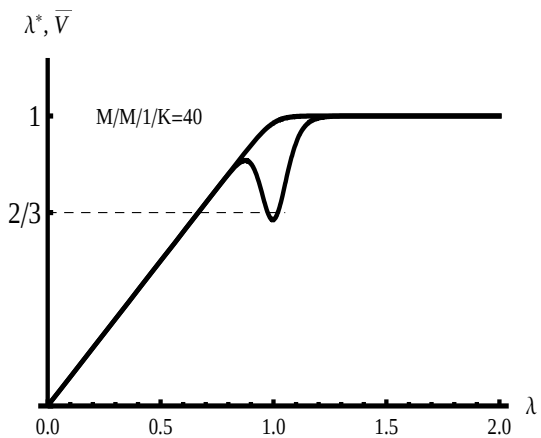

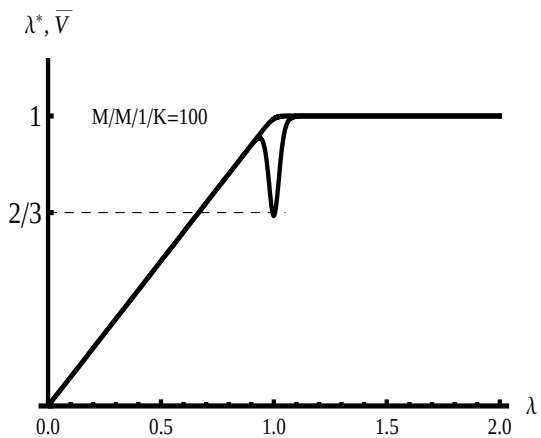

M/M/1/K Queue

Here π_i is truncated geometric distribution when $\lambda \neq \mu$ and a uniform distribution when $\lambda = \mu$

Using $\mathcal{D} = 1 - 2\frac{\pi_J}{1-\pi_J}\sum_{i=0}^{J} P_i\left(1 - \pi_J \frac{P_i}{\pi_i}\right)$:

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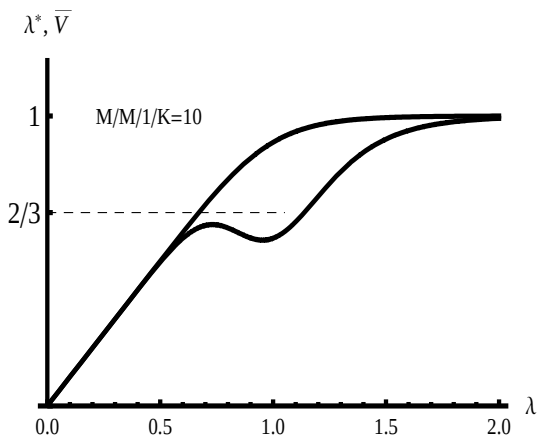

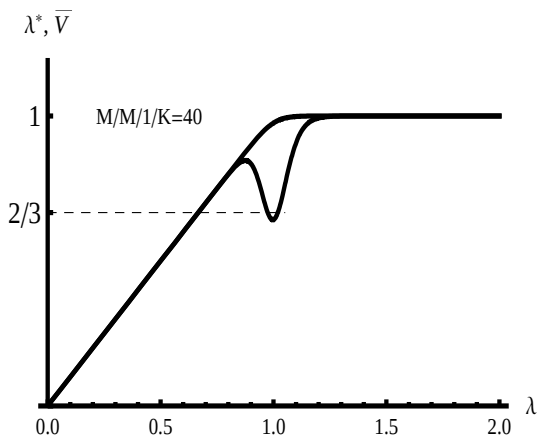

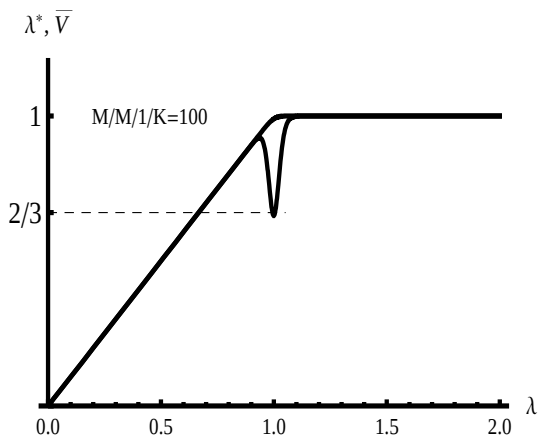

We call this **BRAVO**:

Balancing **R**educes **A**symptotic **V**ariance of **O**utputs

M/M/1 Queue

When $K = \infty$, the formula for $\mathcal{D}$ does not hold. In this case,

$$\mathcal{D} = \begin{cases} 1, & \lambda \neq \mu, \\ ?, & \lambda = \mu. \end{cases}$$

A guess is $\frac{2}{3}$, since for $K < \infty$, $\mathcal{D} = \frac{2}{3} + o_K(1)$...

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Theorem:
Ahmad Al-Hanbali, Michel Mandjes, Y. N., Ward Whitt, 2011

For the M/M/1 queue with $\lambda = \mu$ and arbitrary initial conditions of $Q(0)$ (with finite second moments),

$$\mathcal{D} = 2\left(1 - \frac{2}{\pi}\right) \approx 0.727.$$

Proof based on analysis of classic Laplace transform of generating function of $D(\chi)$ where χ is an exponential random variable.

G/G/1 Queue

Moving away from the memory-less assumptions,

$$\mathcal{D} = \begin{cases} c_a^2, & \lambda < \mu, \\ ?, & \lambda = \mu, \\ c_s^2, & \lambda > \mu. \end{cases}$$

For M/M/1 it was $2(1 - \frac{2}{\pi})$...

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$$\mathcal{D} = \begin{cases} c_a^2, & \lambda < \mu, \\ ?, & \lambda = \mu, \\ c_s^2, & \lambda > \mu. \end{cases}$$

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Theorem:
Ahmad Al-Hanbali, Michel Mandjes, Y.N., Ward Whitt, 2011

For the G/G/1 queue with $\lambda = \mu$, arbitrary finite second moment initial conditions $\big(Q(0), V(0), U(0)\big)$, and finite fourth moments of the inter-arrival and service times,

$$\mathcal{D} = (c_a^2 + c_s^2)\Big(1 - \frac{2}{\pi}\Big).$$

Proof based on diffusion limit of $(D(n\cdot) - \lambda n\cdot)/\sqrt{\lambda n\cdot}$ as $n \to \infty$ (Iglehart and Whitt 1971). Fourth moments are a technical condition used in establishing uniform integrability.

G/G/1/K Queue

$$\mathcal{D} = \begin{cases} c_a^2 + o_K(1), & \lambda < \mu, \\ ?, & \lambda = \mu, \\ c_s^2 + o_K(1), & \lambda < \mu. \end{cases}$$

For $M/M/1/K$ it was $\frac{2}{3} + o_K(1)$, for $G/G/1$ it was $(c_a^2 + c_s^2)(1 - \frac{2}{\pi})$...

G/G/1/K Queue

$$\mathcal{D} = \begin{cases} c_a^2 + o_K(1), & \lambda < \mu, \\ ?, & \lambda = \mu, \\ c_s^2 + o_K(1), & \lambda < \mu. \end{cases}$$

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Conjecture (numerically tested): Y.N., 2011

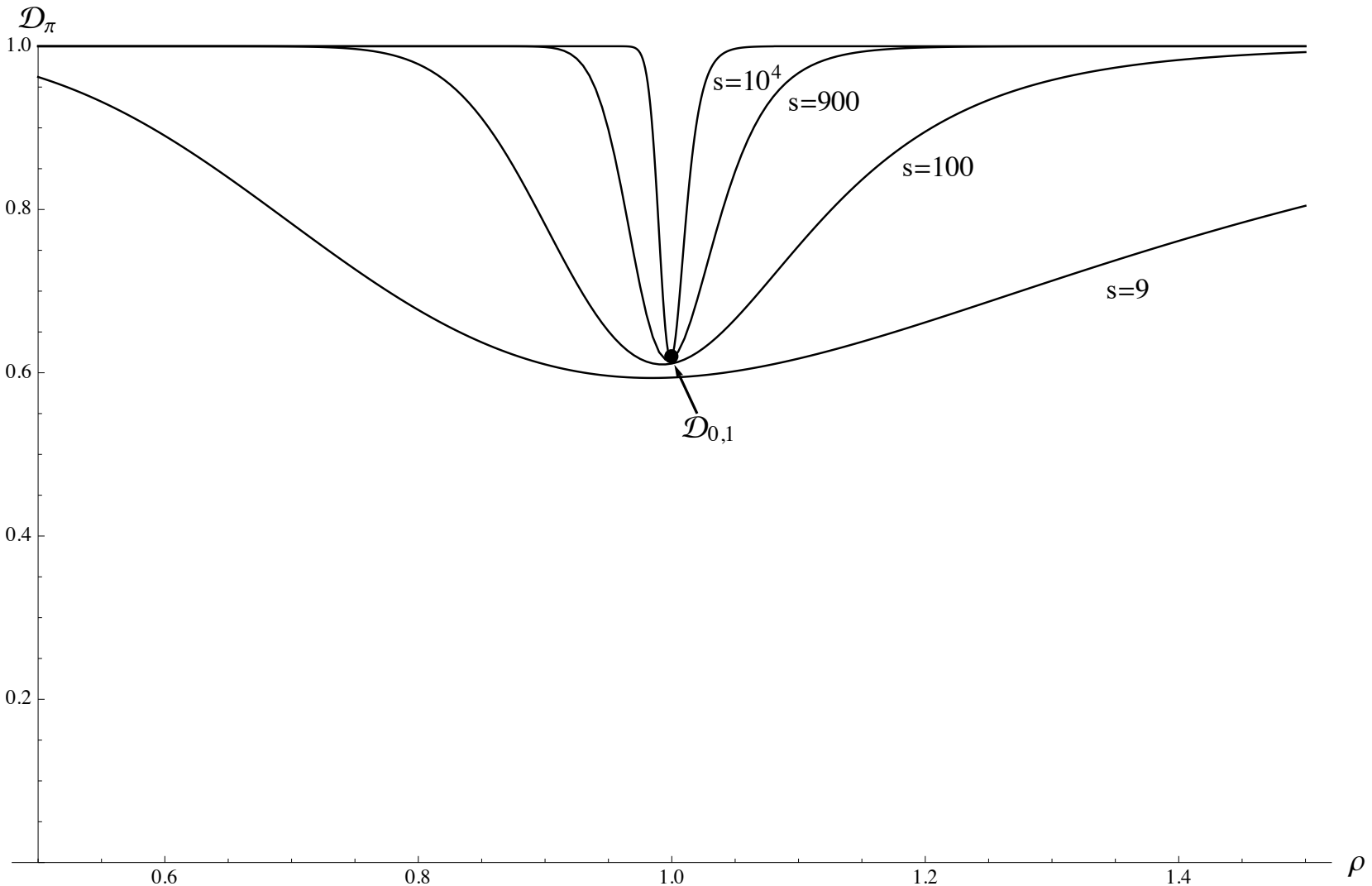
For the G/G/1/K queue with $\lambda = \mu$ and arbitrary initial conditions and light-tailed service and inter-arrival times,

$$\mathcal{D} = (c_a^2 + c_s^2)\frac{1}{3} + O(\frac{1}{K}).$$

Numerical verification done by representing the system as PH/PH/1/K MAPs

Many Servers

M/M/s/$\lfloor\sqrt{s}\rfloor$



Quality and Efficiency Driven (QED) Scaling Regime

A sequence of systems

Consider a sequence of $M/M/s/K$ queues with increasing $s = 1, 2, ...$ and with $\rho_s := \frac{\lambda}{s\mu}$ and K_s such that,

$$(1 - \rho_s)\sqrt{s} \to \beta \in (-\infty, \infty)$$
$$\frac{K_s}{\sqrt{s}} \to \eta \in (0, \infty)$$

So for large s:

$$\rho_s \approx 1 - \beta/\sqrt{s}$$
$$K_s \approx \eta\sqrt{s}$$

Halfin, Whitt, 1981, Garnett, Mandelbaum, Reiman 2002, Borst, Mandelbaum, Reiman, 2004, Whitt, 2004, Pang, Talreja, Whitt, 2007, Janssen, van Leeuwaarden, Zwart, 2011, Kaspi, Ramanan 2011...

Favorable QED Properties

- Probability of delay converges to a value $\in (0, 1)$
- Mean waiting times are typically $O(s^{-1/2})$
- Large queue lengths almost never occur
- Quick mixing times
- In applications: Call-centers (etc...) describes behavior well and allows for asymptotic approximate optimization of staffing etc...
- How about BRAVO?

BRAVO for QED Queues

M/M/s/K QED BRAVO

Theorem: Daryl Daley, Johan van Leeuwaarden, Y.N. 2013

Consider QED scaling with $\beta \neq 0$:

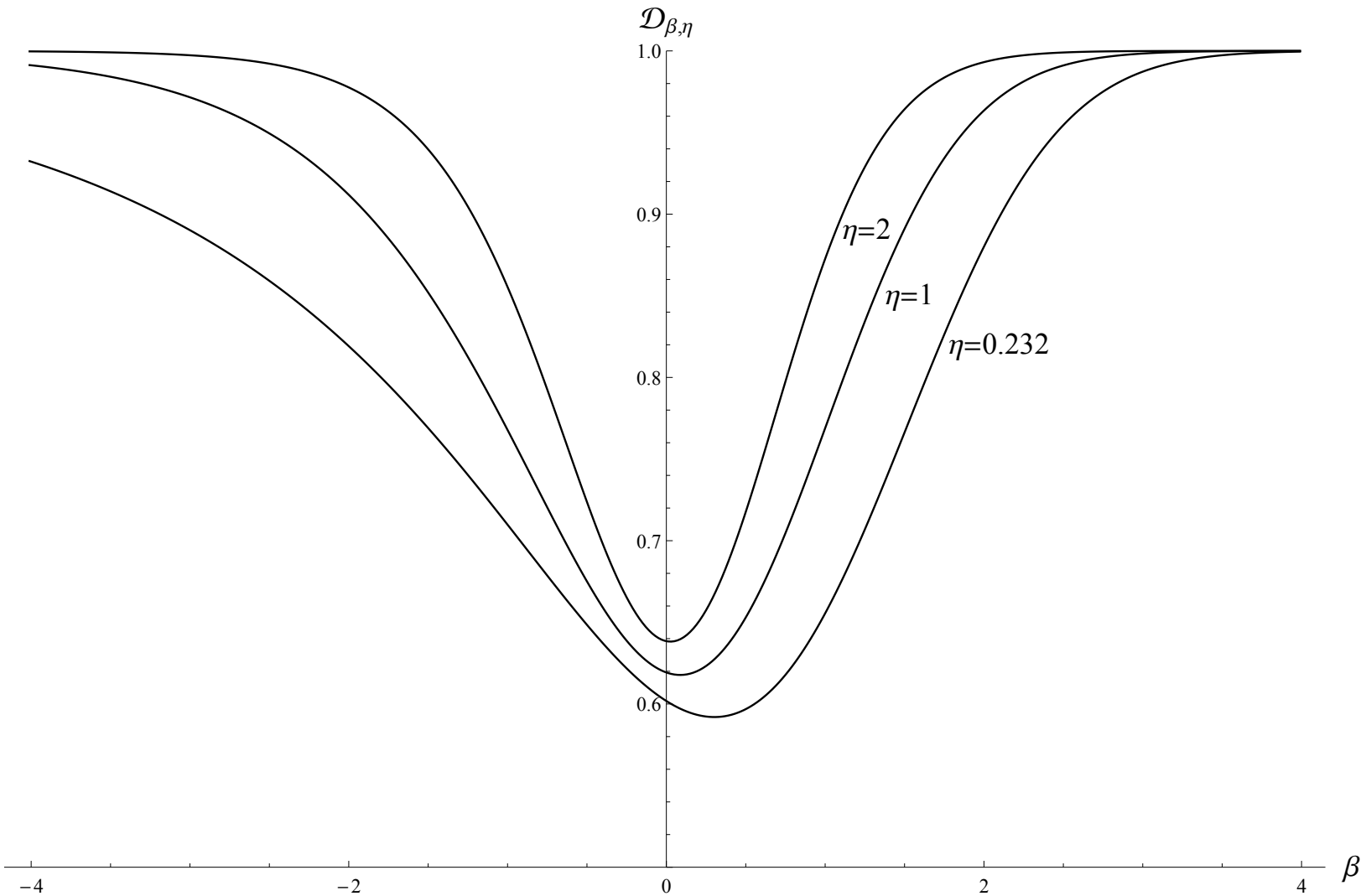
$$\mathcal{D}_{\beta,\eta} := \lim_{s,K\to\infty} \lim_{t\to\infty} \frac{Var\big(D(t)\big)}{\mathbb{E}\big(D(t)\big)},$$

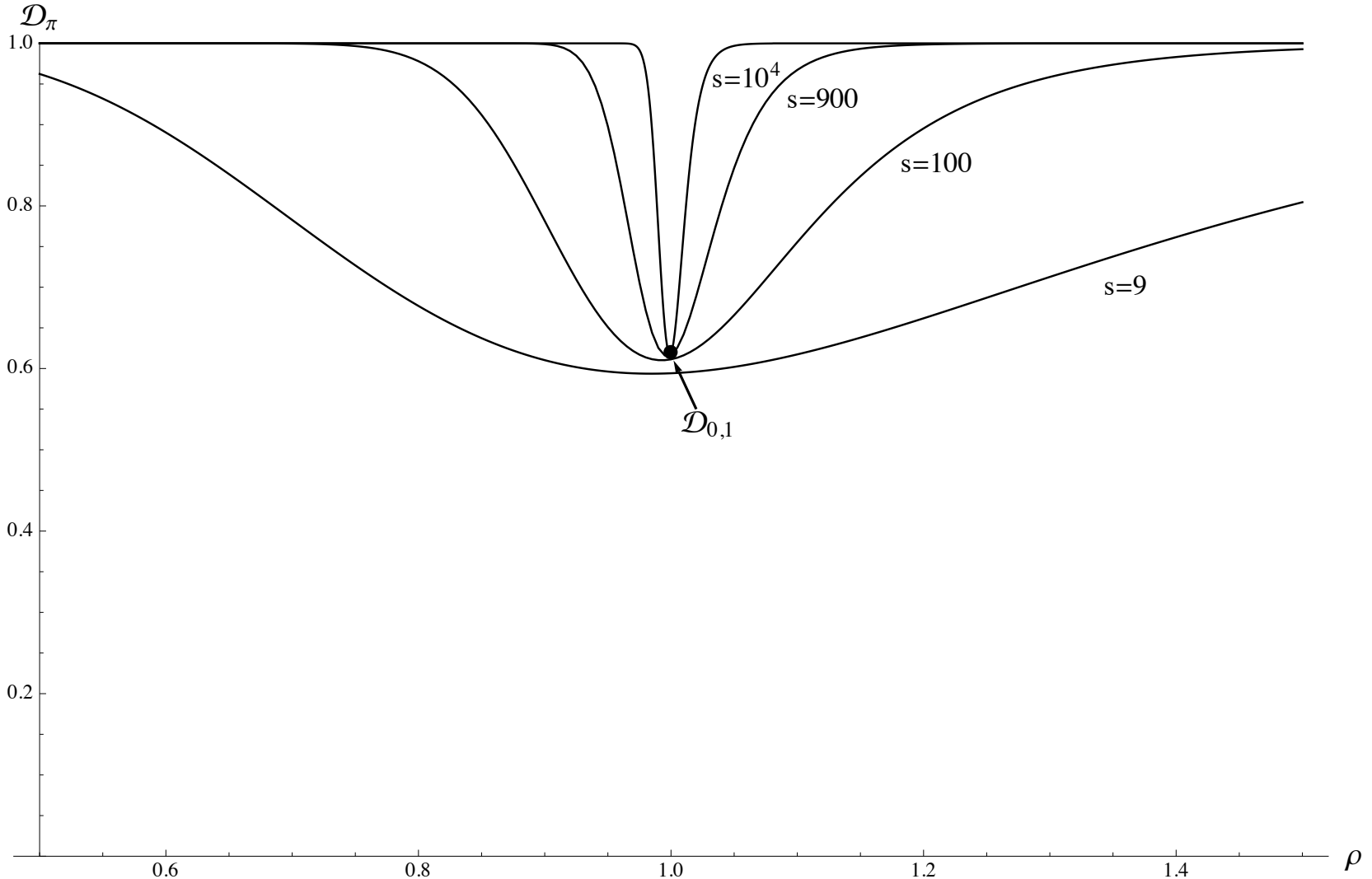
$$\mathcal{D}_{\beta,\eta} = 1 - \frac{2\beta^2 e^{-\beta\eta} h^2}{\phi(\beta)} \int_{-\beta}^{\infty} \Big(1 - \beta e^{-\beta\eta} h \frac{\Phi(-u)}{\phi(u)}\Big)\, \Phi(-u)\, du$$
$$+ 2e^{-\beta\eta} h(1 + e^{-\beta\eta} h)\Big(1 - \beta\eta - e^{-\beta\eta} + (1 - 2\beta\eta e^{-\beta\eta} - e^{-2\beta\eta})h\Big)$$

where

$$h = \lim_{s\to\infty} \frac{\mathbb{P}\big(Q_s \geq s\big)}{1 - e^{-\beta\eta}} = \frac{1}{1 - e^{-\beta\eta} + \frac{\beta\Phi(\beta)}{\phi(\beta)}}$$

BRAVO viewed through the QED lens




$\mathcal{D}_\pi$
1.0
0.8
0.6
0.4
0.2
0.6
0.8
1.0
1.2
1.4
ρ
s=10^4
s=900
s=100
s=9
$\mathcal{D}_{0,1}$

M/M/s/K QED BRAVO with $\rho \equiv 1$ ($\beta = 0$)

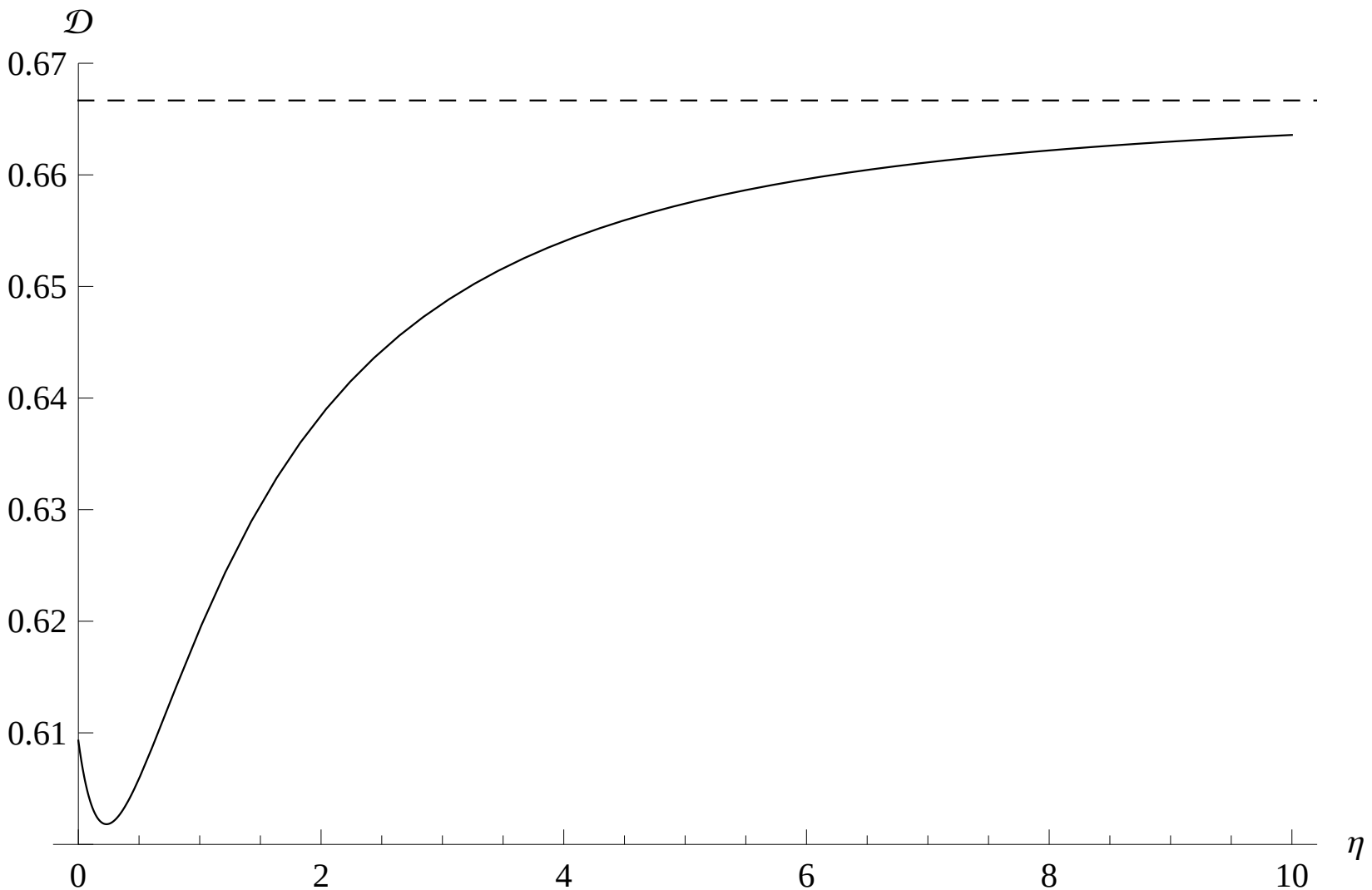
Theorem: Daryl Daley, Johan van Leeuwaarden, Y.N. 2013

Assume $\rho \equiv 1$ and $\frac{K_s}{\sqrt{s}} \to \eta \in (0, \infty)$. Then

$$\mathcal{D}_{0,\eta} := \lim_{s,K\to\infty} \lim_{t\to\infty} \frac{Var\big(D(t)\big)}{\mathbb{E}\big(D(t)\big)},$$

$$\mathcal{D}_{0,\eta} = \frac{2}{3} - \frac{\big(6 - \frac{3\pi}{2}\big)\,\eta - \frac{1}{2}\pi\sqrt{\frac{\pi}{2}} + 3\sqrt{2\pi}(1 - \log 2)}{3\left(\eta + \sqrt{\frac{\pi}{2}}\right)^3}.$$

$\mathrm{M}/\mathrm{M}/s/\lfloor \eta\sqrt{s} \rfloor \quad s \to \infty \quad$ at $\rho \equiv 1$ $(\beta = 0)$



Idea of BRAVO QED Proofs

Use

$$\mathcal{D} = 1 - 2\frac{\pi_J}{1-\pi_J}\sum_{i=0}^{J} P_i\left(1 - \pi_J\frac{P_i}{\pi_i}\right).$$

Using QED scaling:

$$(1-\rho_s)\sqrt{s} \to \beta, \qquad \frac{K_s}{\sqrt{s}} \to \eta,$$

"simply evaluate" the limit,

$$\lim_{s,K\to\infty} \frac{\pi_J}{1-\pi_J}\sum_{i=0}^{J} P_i\left(1 - \pi_J\frac{P_i}{\pi_i}\right).$$

Intermediate Summary: BRAVO

Known BRAVO constants:

- Single server finite buffer: $2/3$
  (for G/G replace 2 by $c_a^2 + c_s^2$)
- Single server infinite buffer $2(1 - 2/\pi)$:
  (for G/G replace 2 by $c_a^2 + c_s^2$)
- Memoryless many servers finite buffer: $\mathcal{D}_{0,\eta} \in [0.6, 2/3]$

Not yet known:

- Memoryless many servers infinite buffer
- Many servers without memoryless assumptions
- Systems with reneging or other packet loss mechanisms

Other questions: How can BRAVO be harnessed in practice?
Why does BRAVO occur?

Further properties of $\mathrm{Var}\big(D(t)\big)$

The Stable M/G/1 Queue

Theorem: Sophie Hautphenne, Yoav Kerner, Y. N., Peter Taylor, 2013

Consider the stable M/G/1 queue with finite third service moment, parameterized by (*arrival rate, load, scv, skewness*)$= (\lambda, \rho, c^2, \gamma)$.

Stationary version:

$$\mathrm{Var}\big(D(t)\big) = \lambda t + L_e \frac{\rho}{(1-\rho)^2} + o(1),$$

$$L_e = \frac{(3c^4 - 4\gamma c^3 + 6c^2 - 1)\rho^3 + (4\gamma c^3 - 12c^2 + 4)\rho^2 + (6c^2 - 6)\rho}{6}.$$

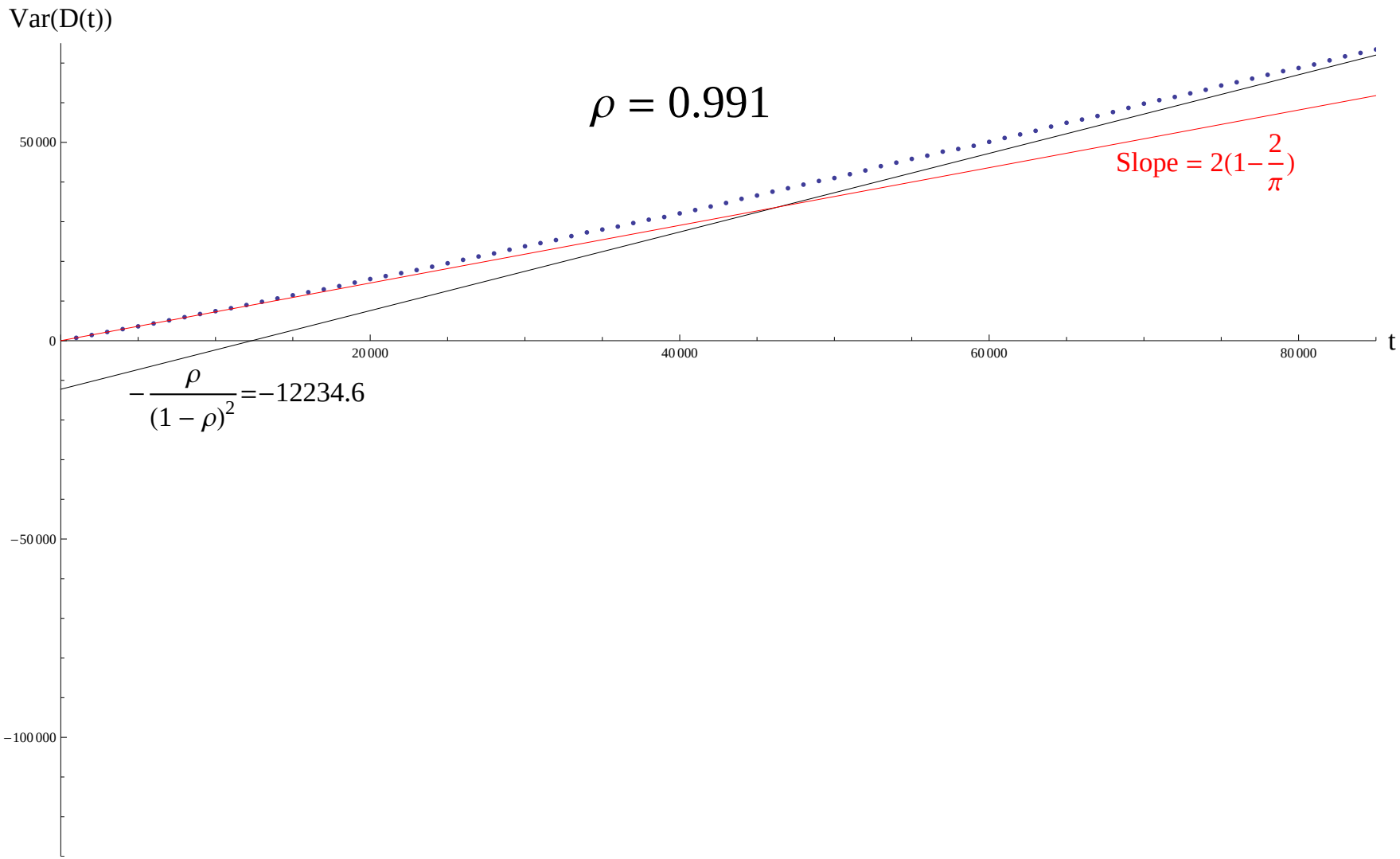
Starting empty version:

$$\mathrm{Var}\big(D(t)\big) = \lambda t - (1 - L_0) \frac{\rho}{(1-\rho)^2} + o(1),$$

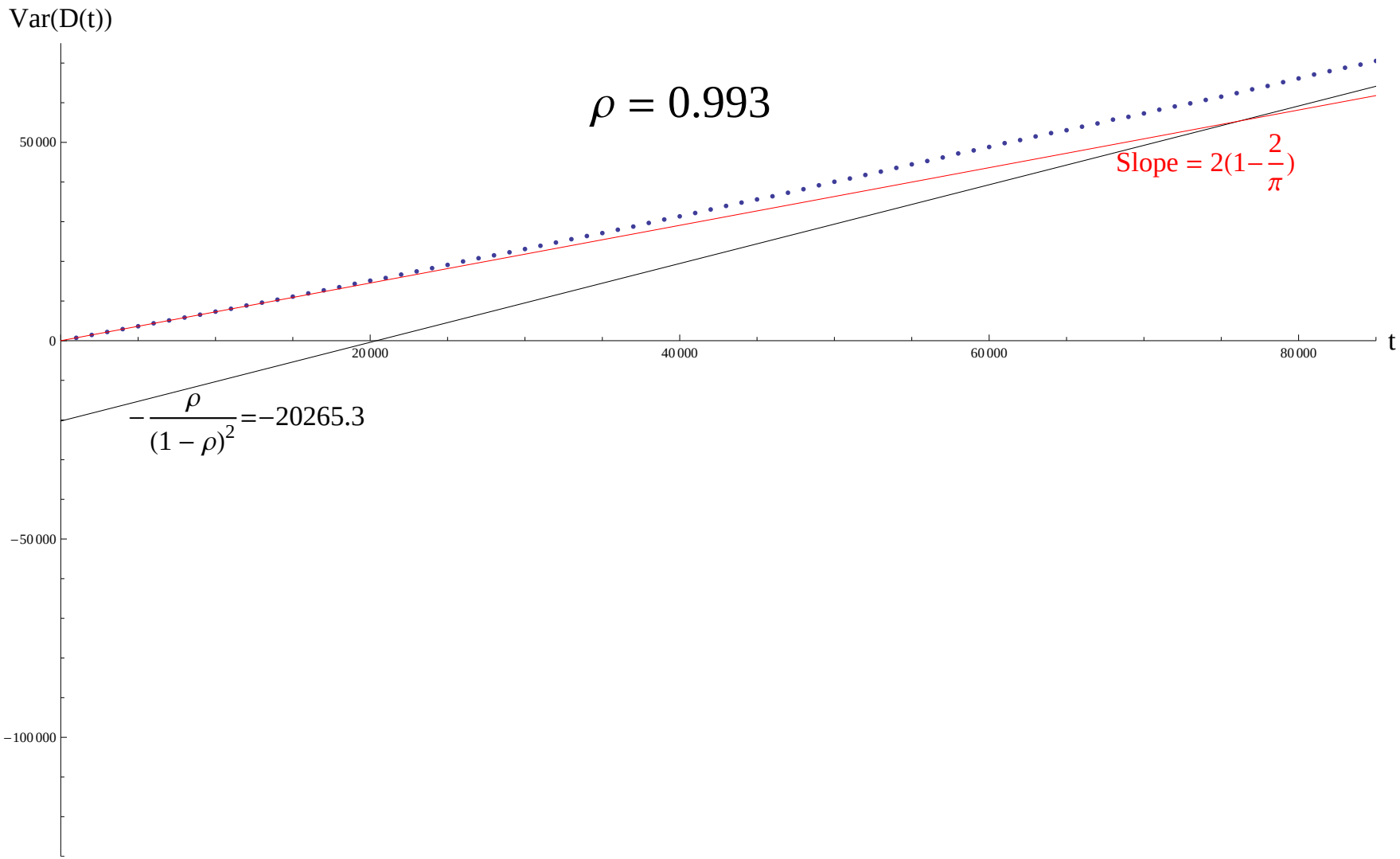
$$L_0 = \frac{(3c^4 - 4\gamma c^3 + 6c^2 - 1)\rho^3 + (4\gamma c^3 - 6c^2 - 2)\rho^2 - (6c^2 - 6)\rho}{12}.$$

M/M/1: $c^2 = 1, \gamma = 2$. $L_e = 0$, $L_0 = 0$.

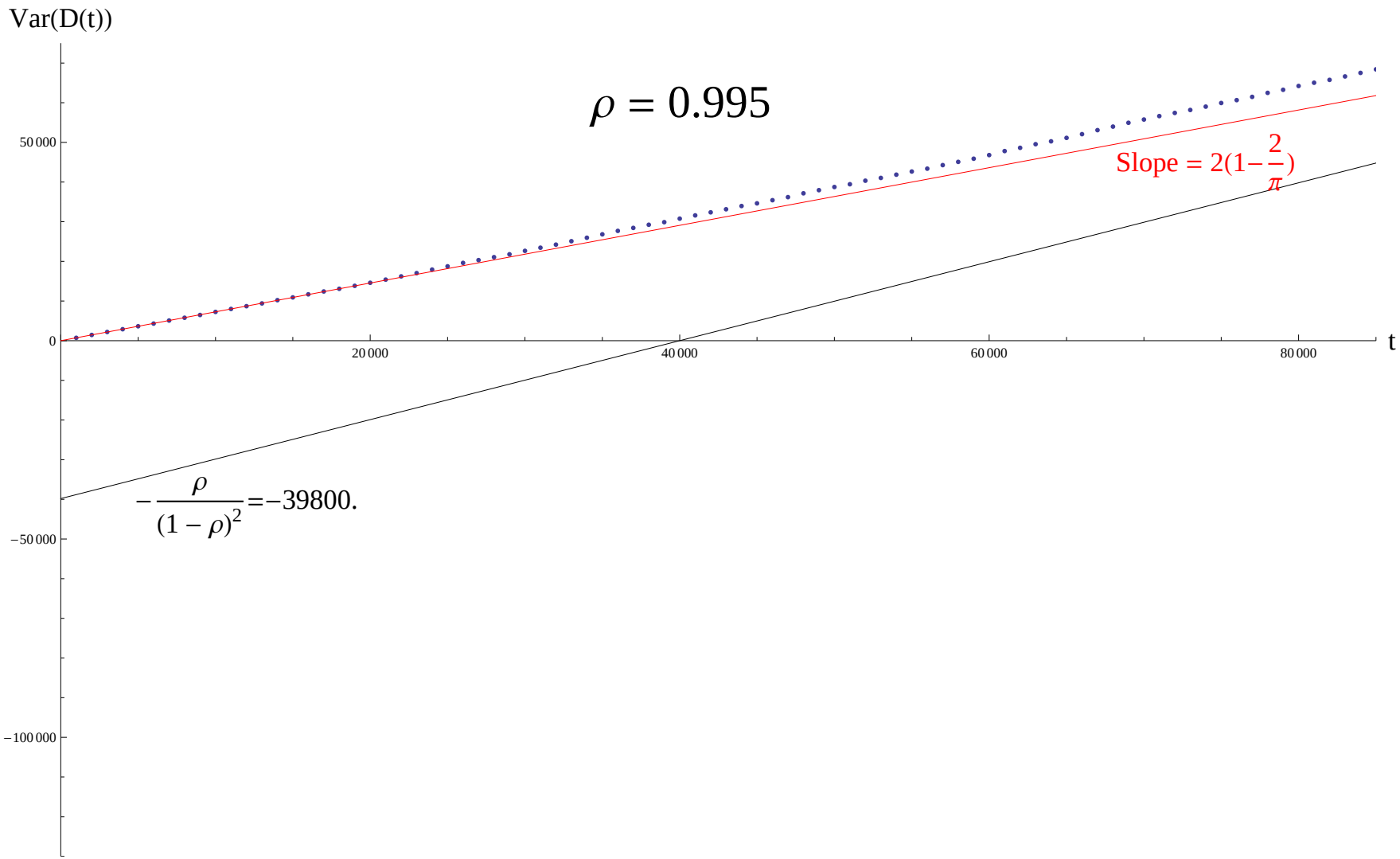
M/M/1 Queue


Var(D(t))
$\rho = 0.991$
Slope = $2(1-\frac{2}{\pi})$
50 000
0
−50 000
−100 000
20 000
40 000
60 000
80 000
t
$-\frac{\rho}{(1-\rho)^2} = -12234.6$

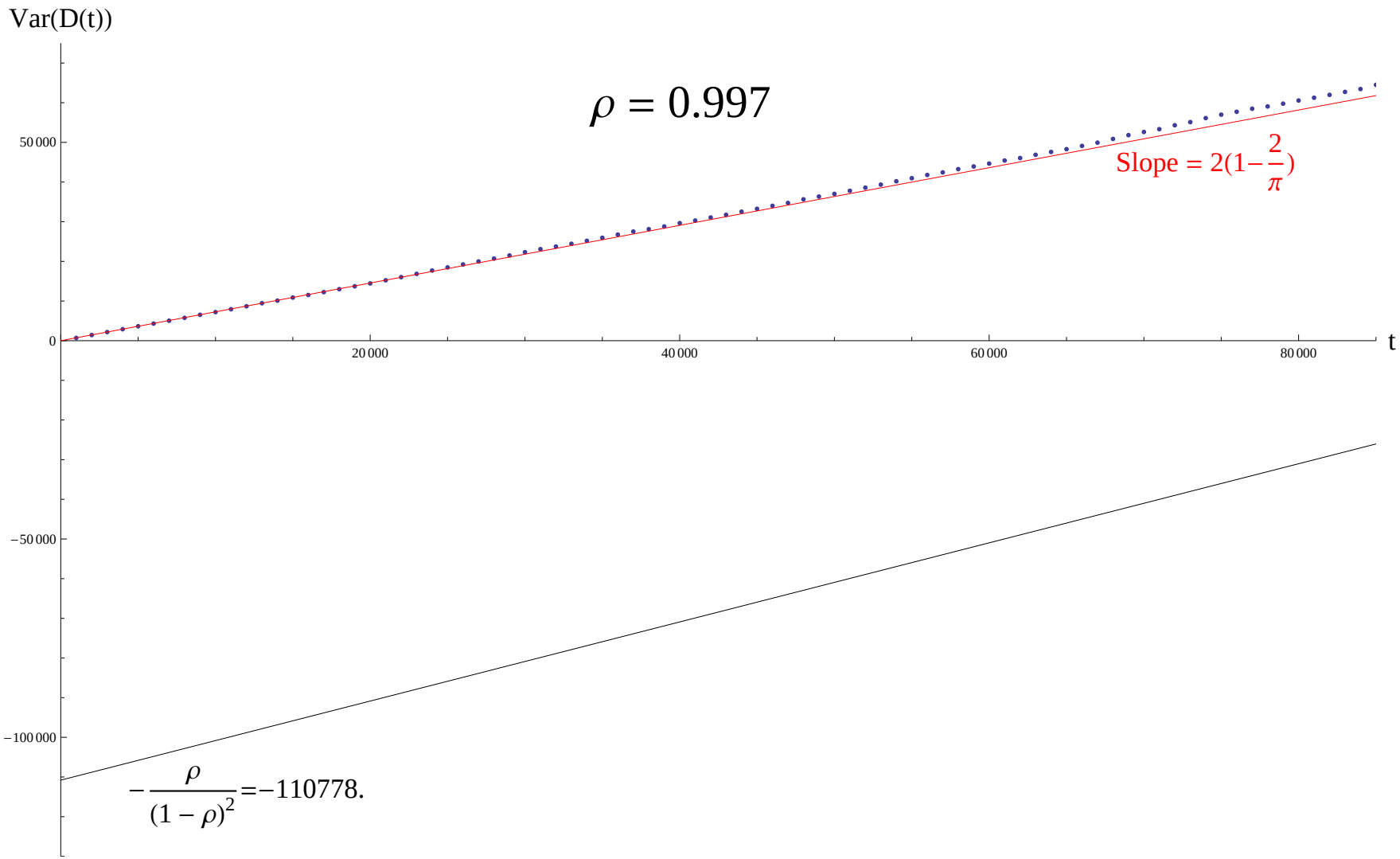
M/M/1 Queue


Var(D(t))
$\rho = 0.993$
Slope = $2(1-\frac{2}{\pi})$
50 000
0
−50 000
−100 000
20 000
40 000
60 000
80 000
t
$-\frac{\rho}{(1-\rho)^2} = -20265.3$

M/M/1 Queue


Var(D(t))
$\rho = 0.995$
Slope = $2(1-\frac{2}{\pi})$
50 000
0
−50 000
−100 000
20 000
40 000
60 000
80 000
t
$-\frac{\rho}{(1-\rho)^2} = -39800.$

M/M/1 Queue


Var(D(t))
$\rho = 0.997$
Slope = $2(1-\frac{2}{\pi})$
50 000
0
−50 000
−100 000
20 000
40 000
60 000
80 000
t
$-\frac{\rho}{(1-\rho)^2} = -110778.$

Asymptotic Variance of Flows in General Open Stable Queueing Networks

Stable Open Queueing Networks

Generalized Jackson network with external arrival vector α, routing matrix P and service capacity vector μ.

Stable if $\nu := (I - P')^{-1}\alpha < \mu$.

Counting processes: $E(\cdot)$ are entrances to nodes (exogenous and endogenous). $D_{i,j}(\cdot)$, are packets passed from node i to j.

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Simple:

$$\nu_i := \lim_{t\to\infty} \frac{\mathbb{E}[E_i(t)]}{t}.$$

Stable Open Queueing Networks

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Simple:

$$\nu_i := \lim_{t\to\infty} \frac{\mathbb{E}[E_i(t)]}{t}.$$

Our contribution: Computable exact formulas for:

$$\sigma_{i,j} := \lim_{t\to\infty} \frac{Cov\Big(E_i(t), E_j(t)\Big)}{t}, \qquad \sigma_{i_1\to j_1, i_2\to j_2} := \lim_{t\to\infty} \frac{Cov\Big(D_{i_1,j_1}(t), D_{i_2,j_2}(t)\Big)}{t}.$$

Main Queueing Network Result

Theorem: Werner Scheinhardt, Y. N. 2013

Define,

$$\Sigma^{(D)} := H\,\Sigma^{(F)}\,H', \qquad \Sigma^{(E)} := \Big(BH + [I \;\; 0]\Big)\Sigma^{(F)}\Big(BH + [I \;\; 0]\Big)'.$$

Here $\Sigma^{(F)}$ is the covariance matrix of the primitive input sequences (arrivals and routing) and the matrices B and H are easily constructed with only the inversion $(I - P')^{-1}$.

Under general assumptions of stable (single or multi-class) queueing networks, the processes $D(\cdot)$ and $E(\cdot)$ converge weakly to Brownian motions with the above covariance matrices.

Further,

$$\sigma_{i_1 \to j_1, i_2 \to j_2} = \Sigma^{(D)}_{(i_1-1)K+j_1,\ (i_2-1)K+j_2}, \qquad \sigma_{i,j} = \Sigma^{(E)}_{i,j}.$$

Note: σ's do not depend on the service variance. Compare with QNA (Queueing Network Analyzer, Whitt 80's).

Wrap Up

Is all this useful?

- BRAVO: Can it be incorporated in system planning, estimation, or control?
- Network Flows: Can heuristic queueing network decomposition schemes be improved?
- Other uses?

References

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